

Discrete HMMs, part 2: Viterbi!

Steven Bedrick CS/EE 5/655, 10/27/14

Quick Review:

Stochastic PoS techniques rely entirely on probability.

The goal of a stochastic PoS tagger is to find:

$$\hat{t}_1^n = \arg\max_{t_1^n} P(t_1^n | w_1^n)$$

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Probability of word given tag

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 Probability of Probability of tag word given tag given previous tag

скрытая марковская модель



Andrei Andreievich Markov 1856–1922

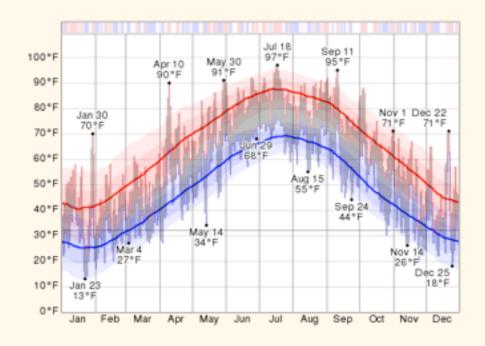
HMMs are a type of stochastic model used to examine sequential data.

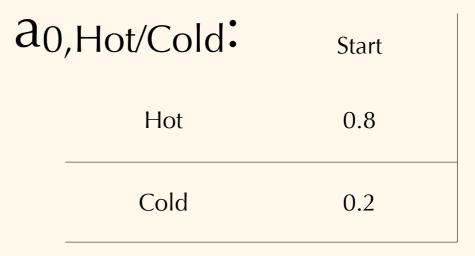
The basic idea: there are two parameters changing over time, but we can only directly observe one of them. We want to know about the other.

Q: { Hot, Cold }

		Hot	Cold
A: -	Hot	0.7	0.3
	Cold	0.6	0.4

		1	2	3
B: -	Hot	0.2	0.4	0.4
	Cold	0.5	0.4	0.1
		0.5	0.4	0.1





Note: for this demonstration, we are ignoring a_F .

There are three fundamental kinds of questions that we can ask with an HMM:

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Let's say we have a sequence of diary entries:

$$O = 3, 1, 3$$

How likely is this sequence given the model described earlier? $P(O|\lambda)$

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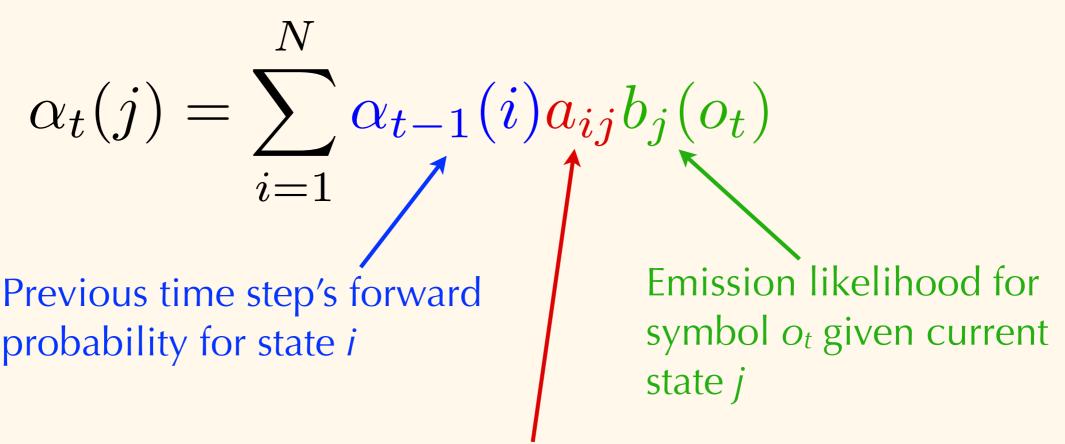
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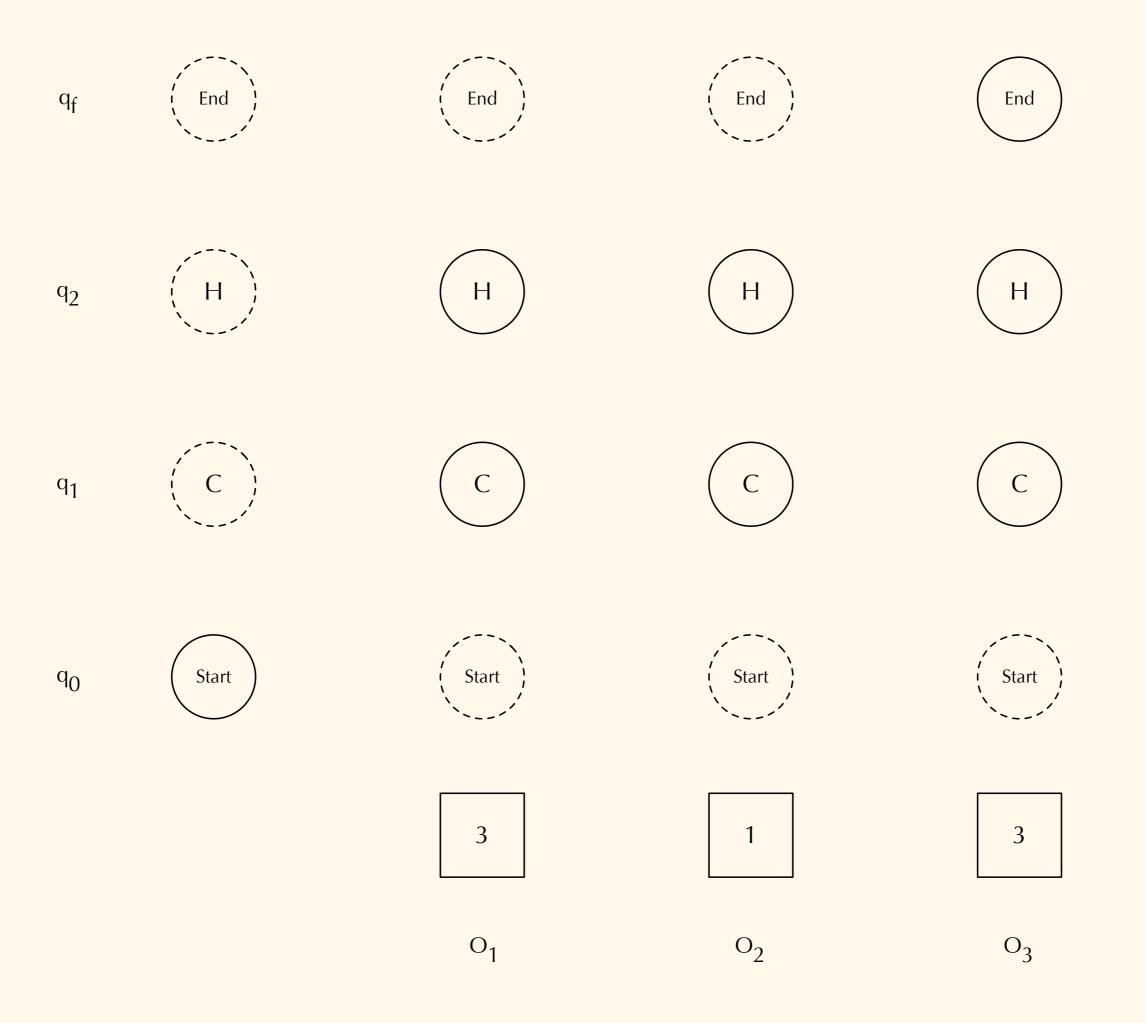
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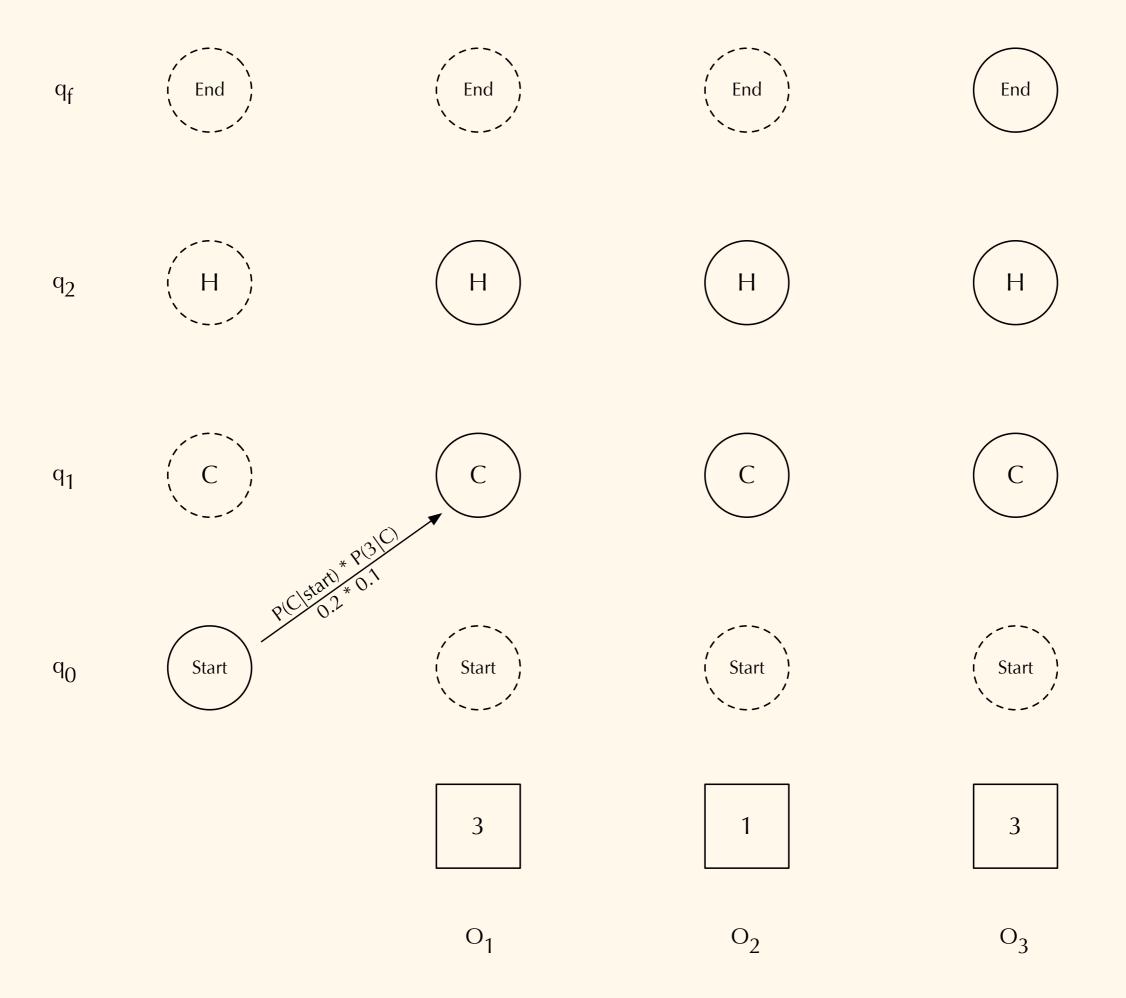
Formally: $\alpha_t(j) = P(o_1, o_2...o_t, q_t = j | \lambda)$

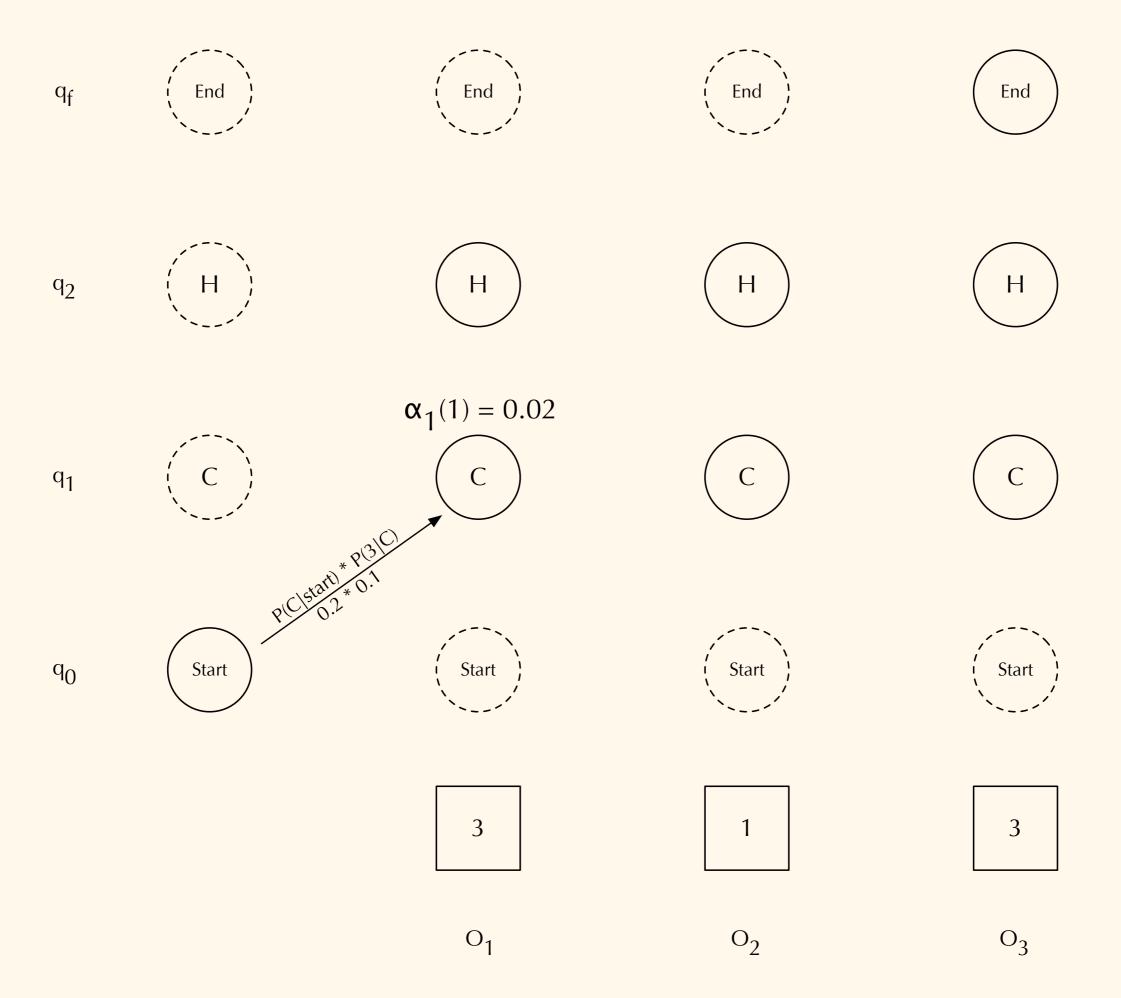
Calculating $\alpha_t(j) = P(o_1, o_2...o_t, q_t = j | \lambda)$ is fairly straightforward:

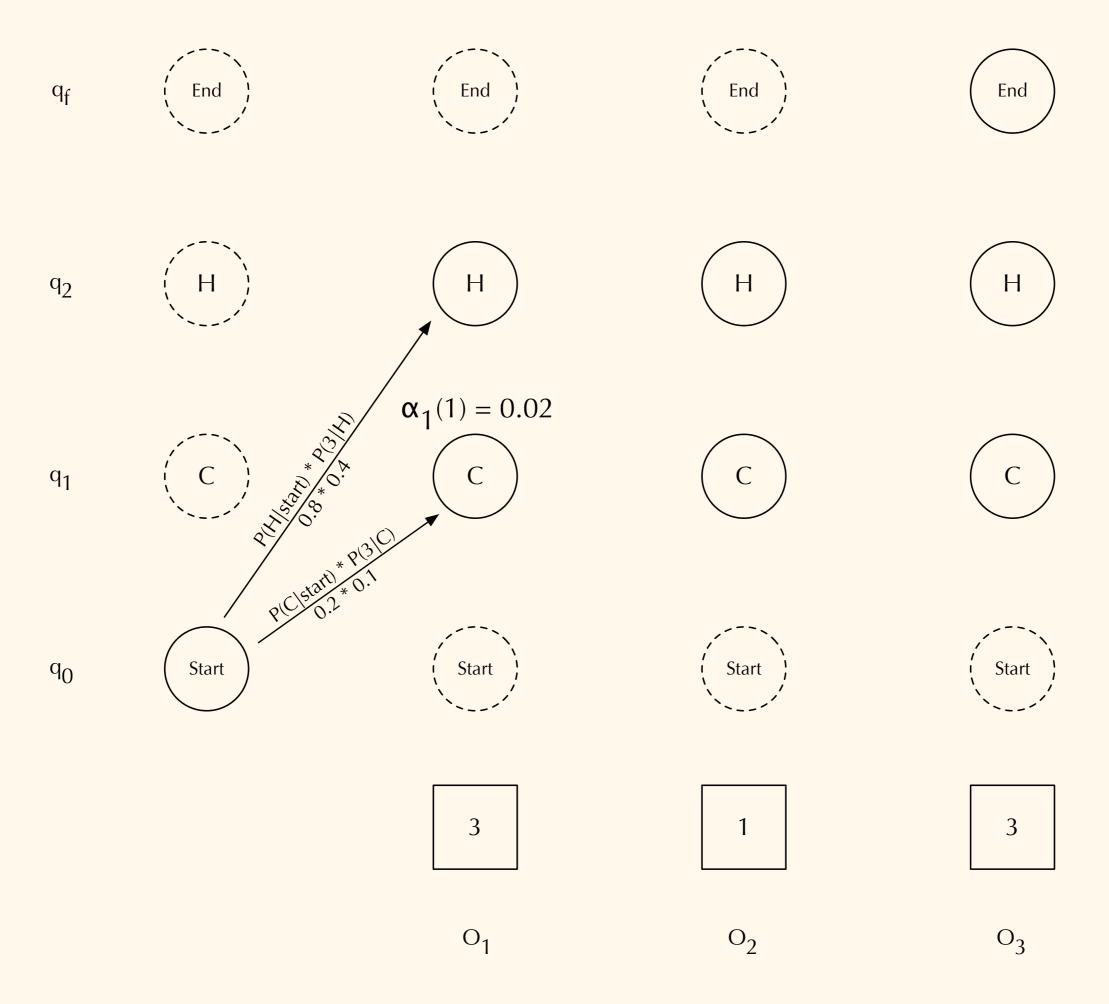


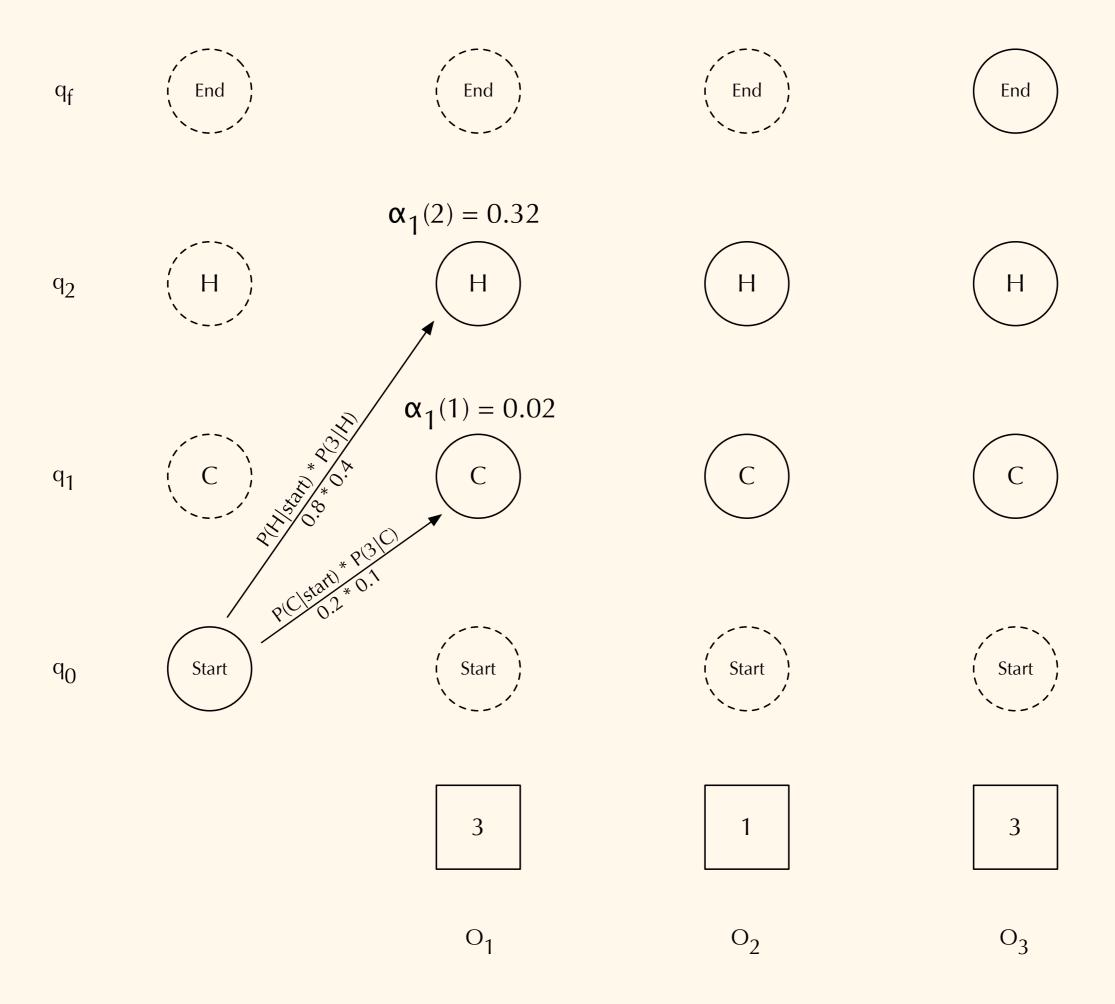
Transition prob. from previous state *i* to current state *j*

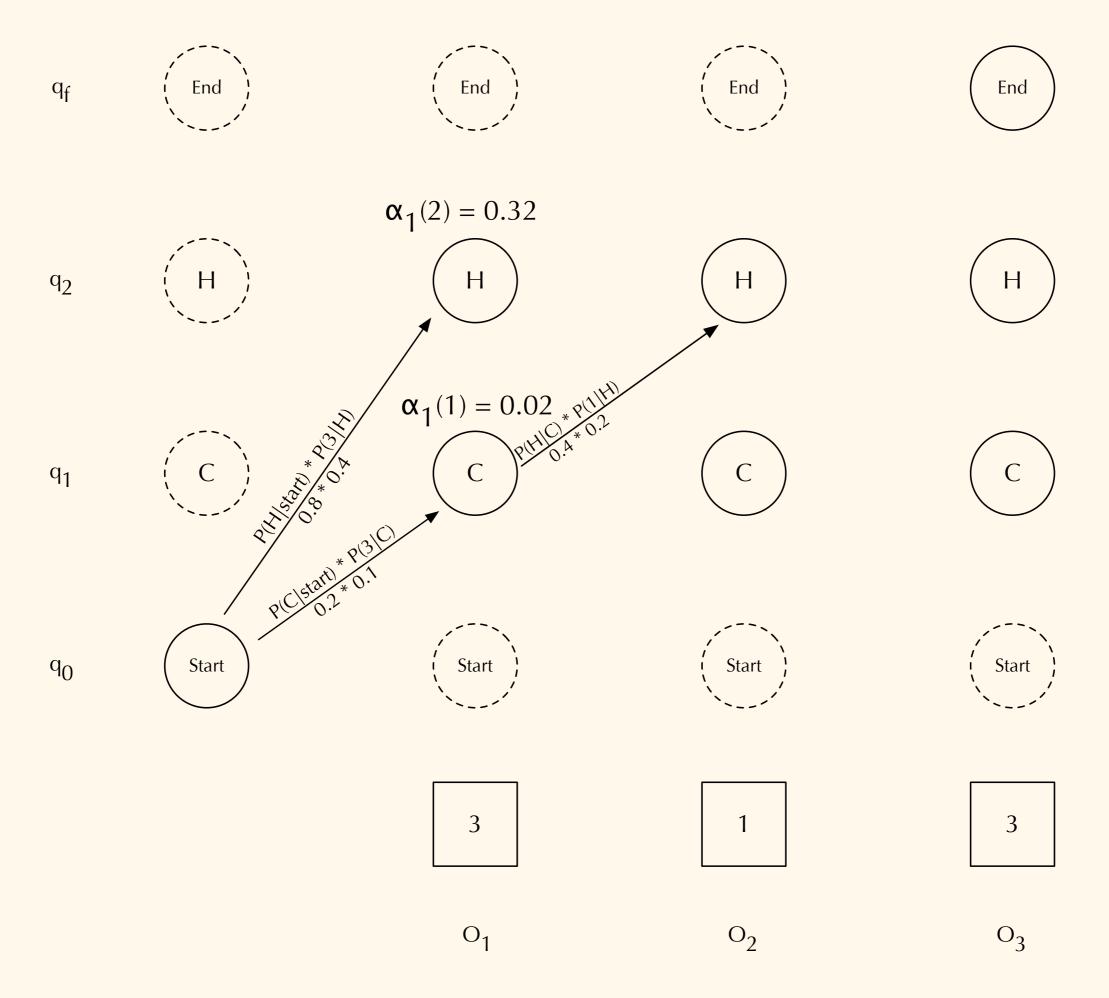


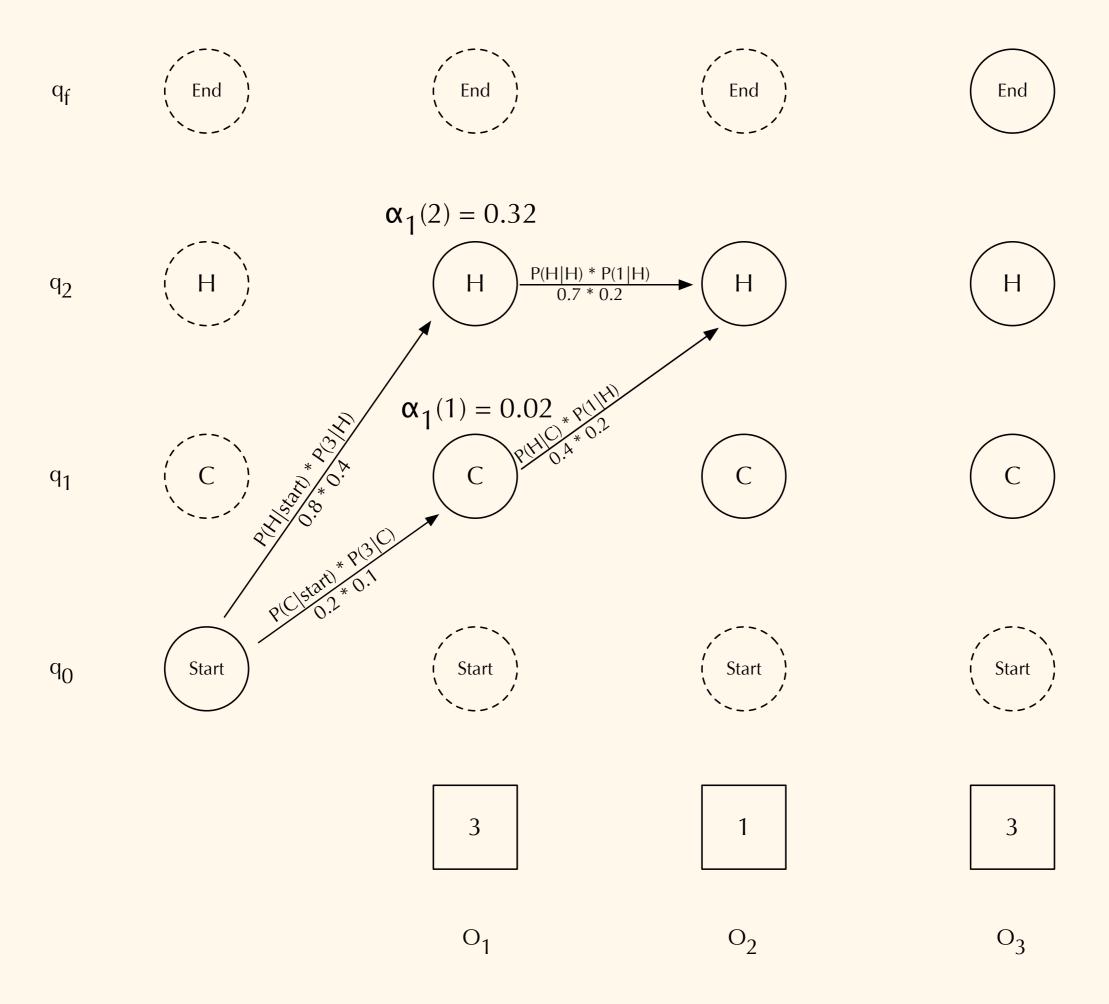


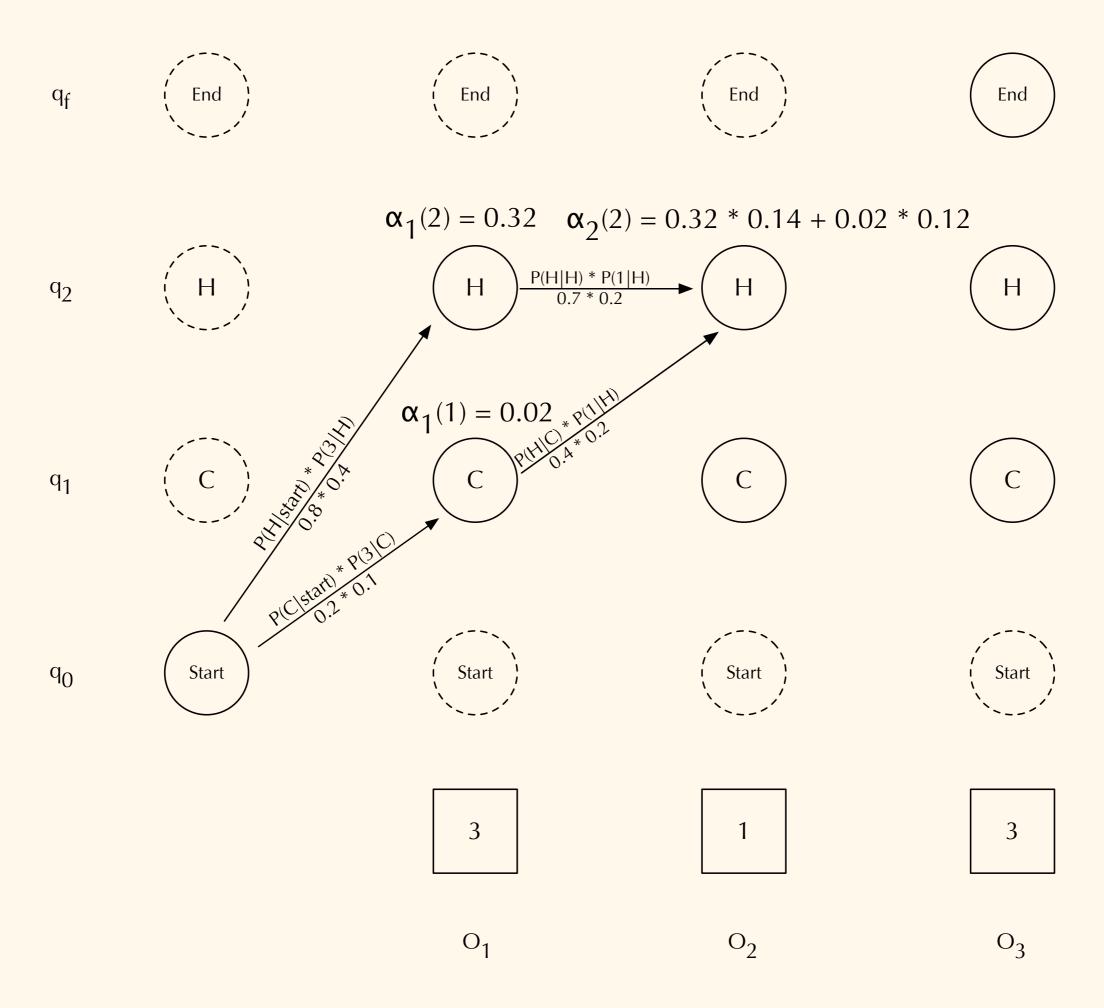


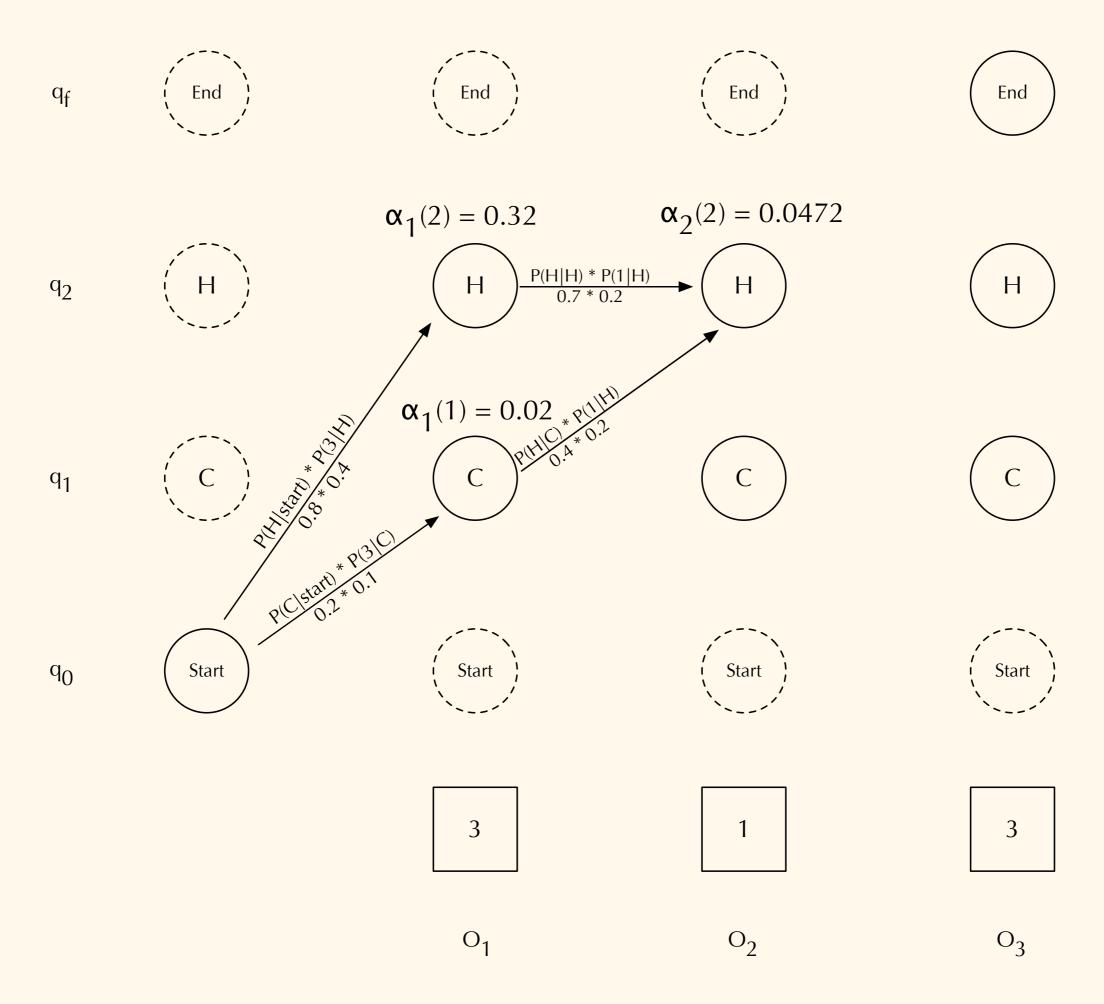


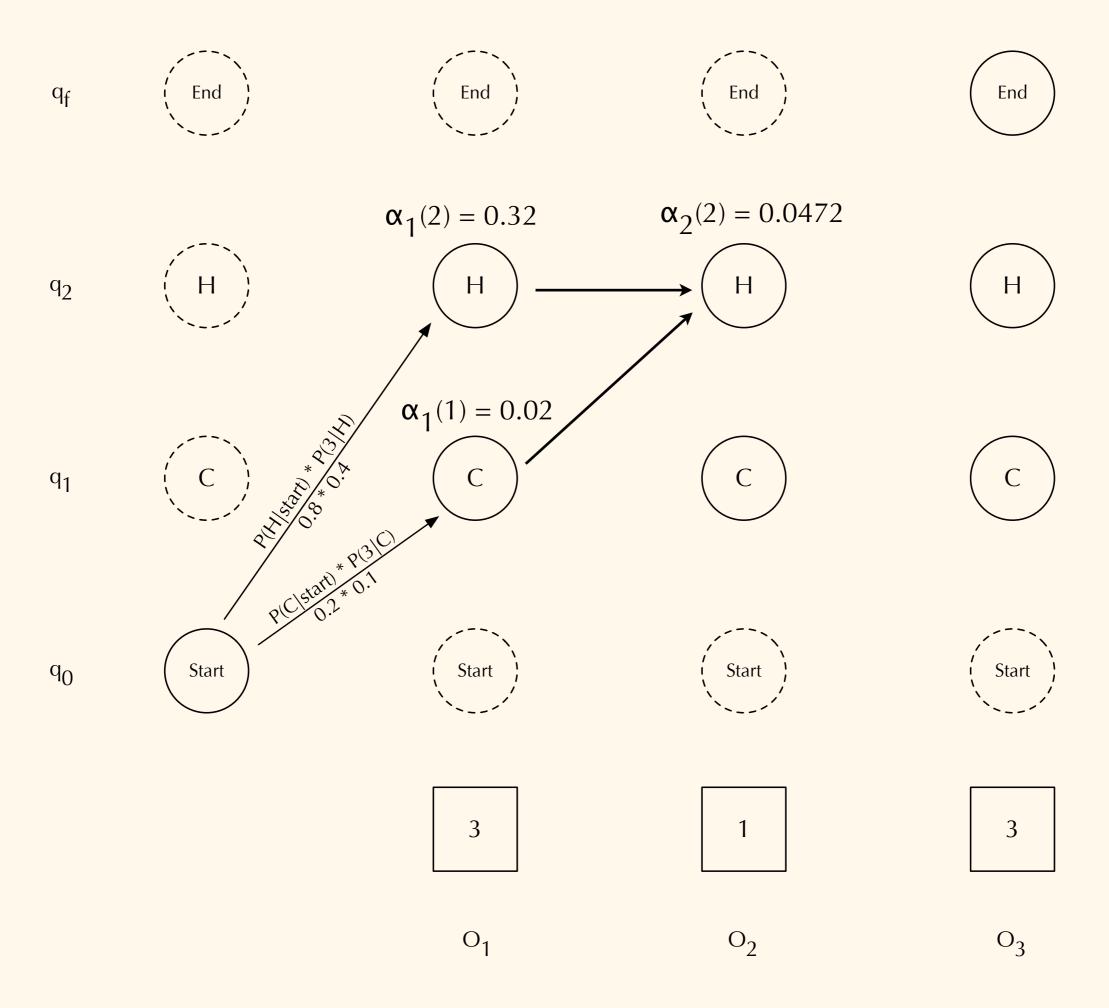


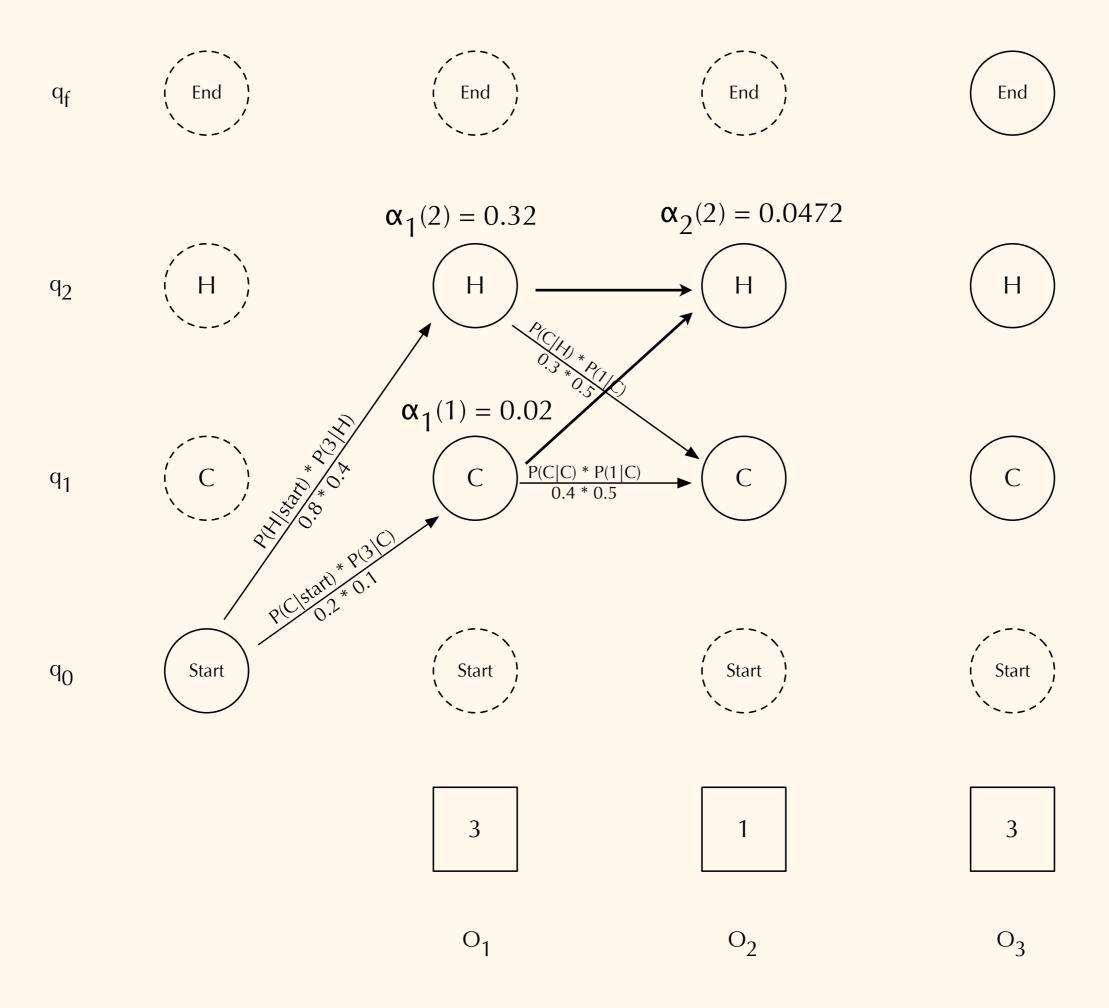


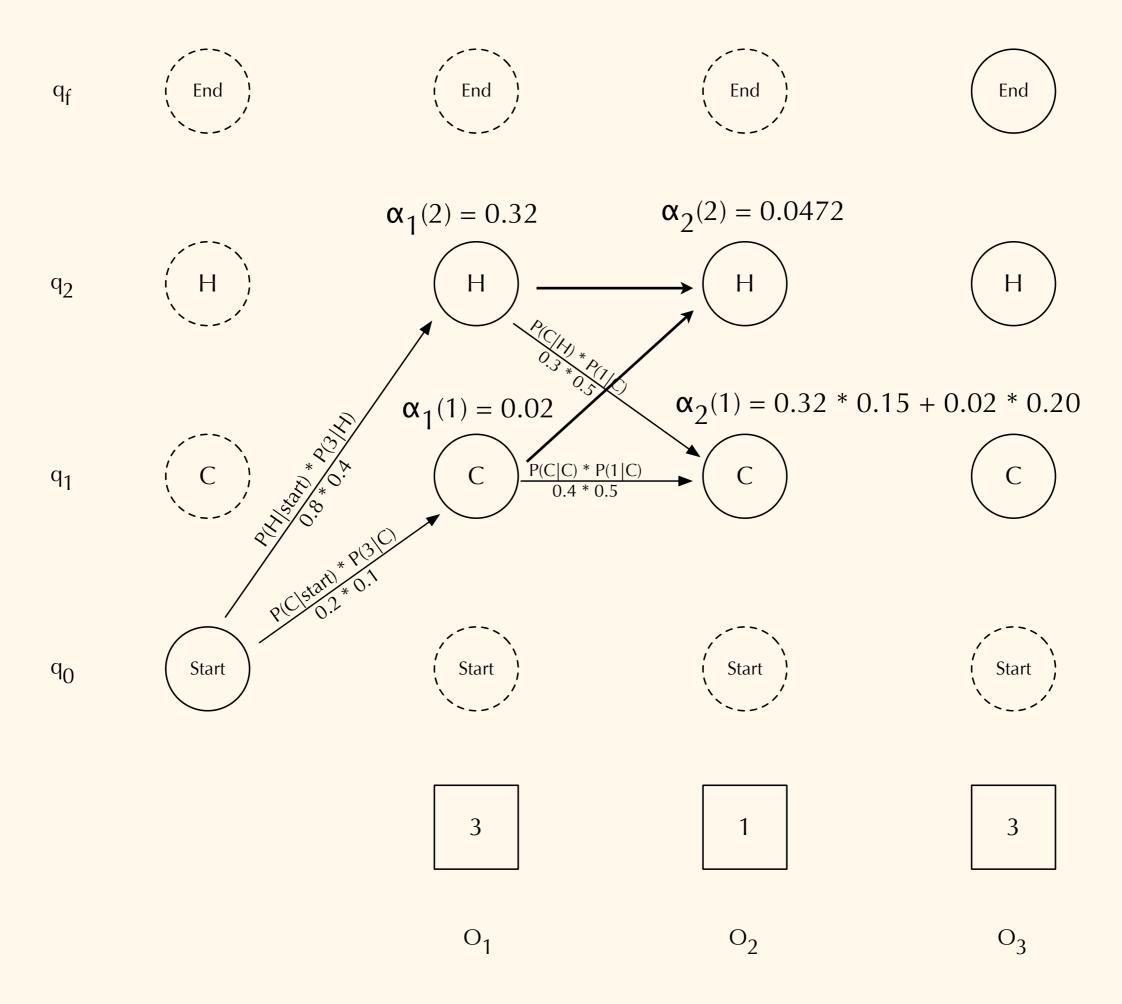


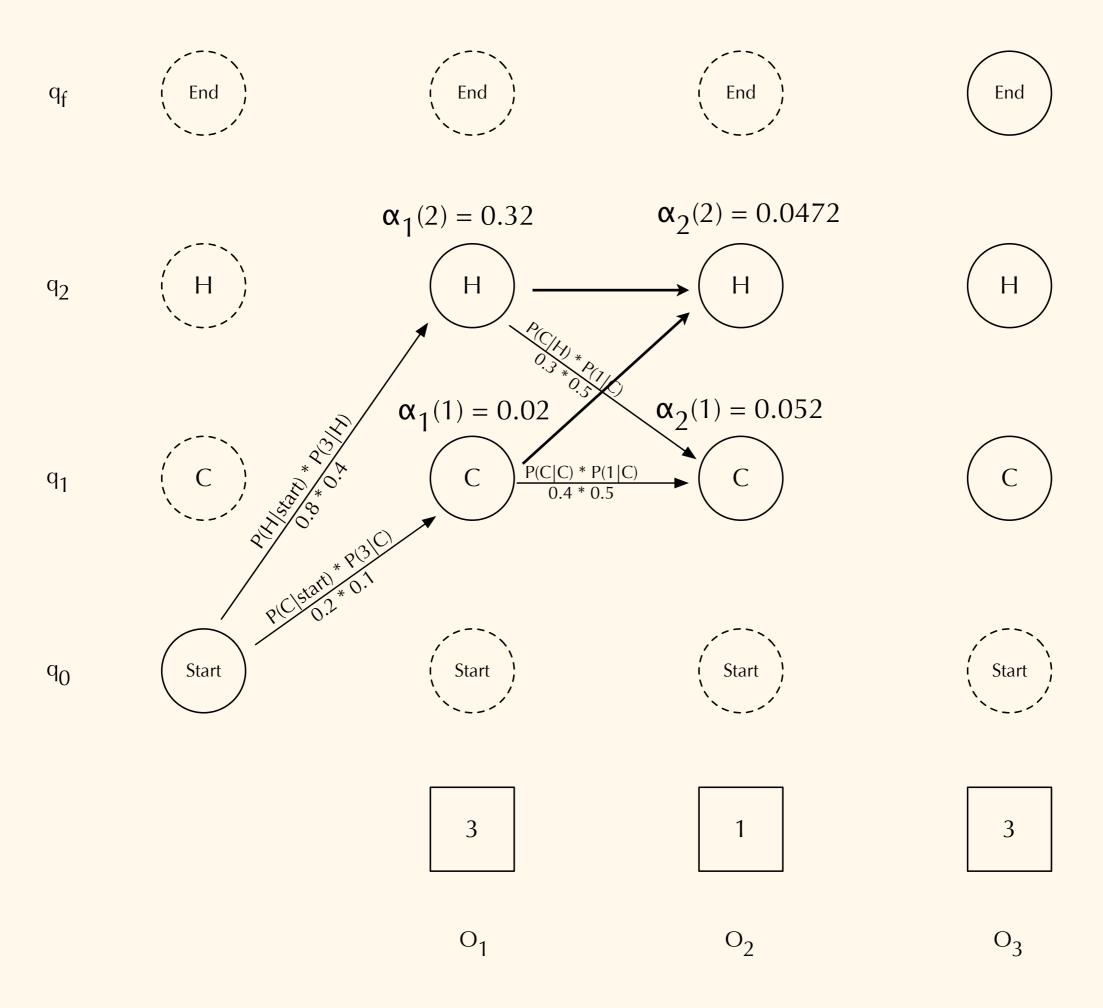


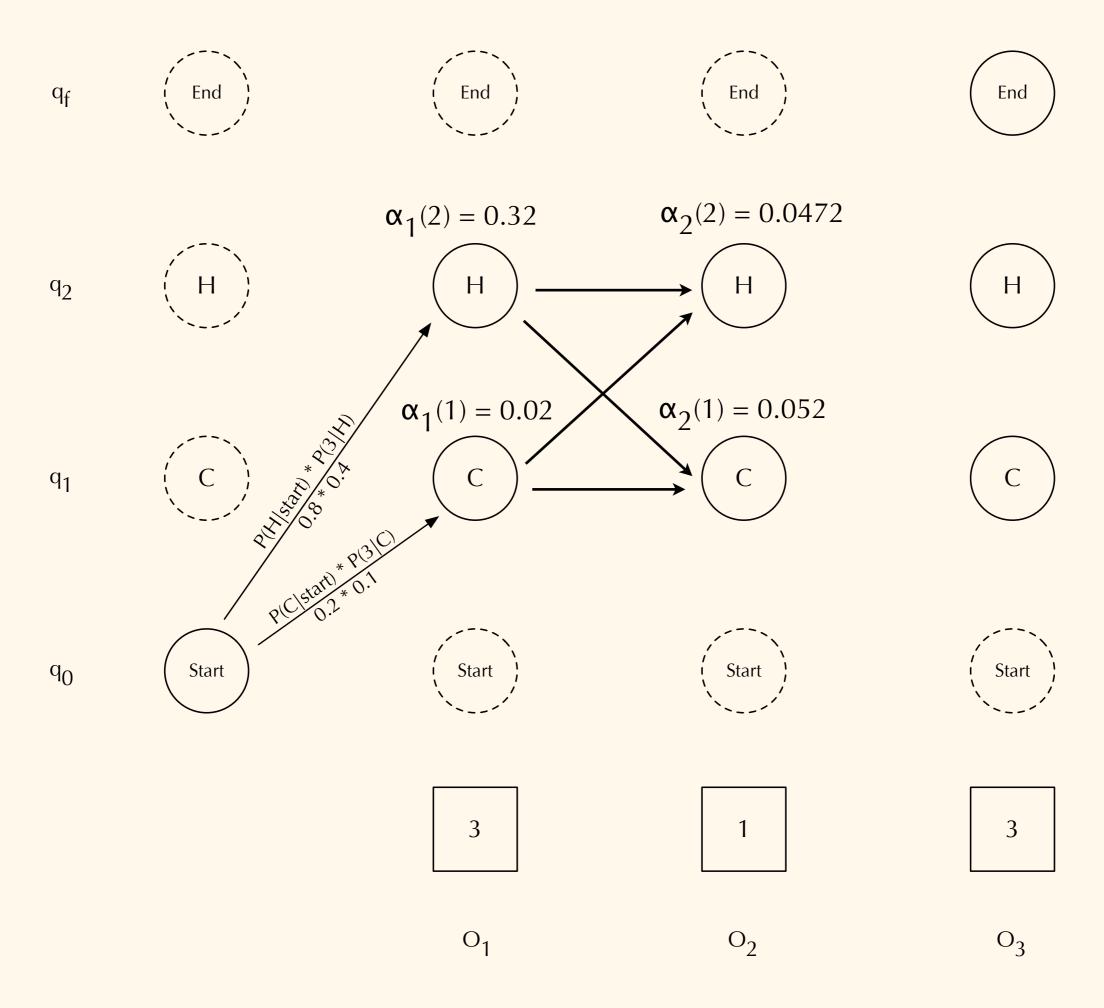


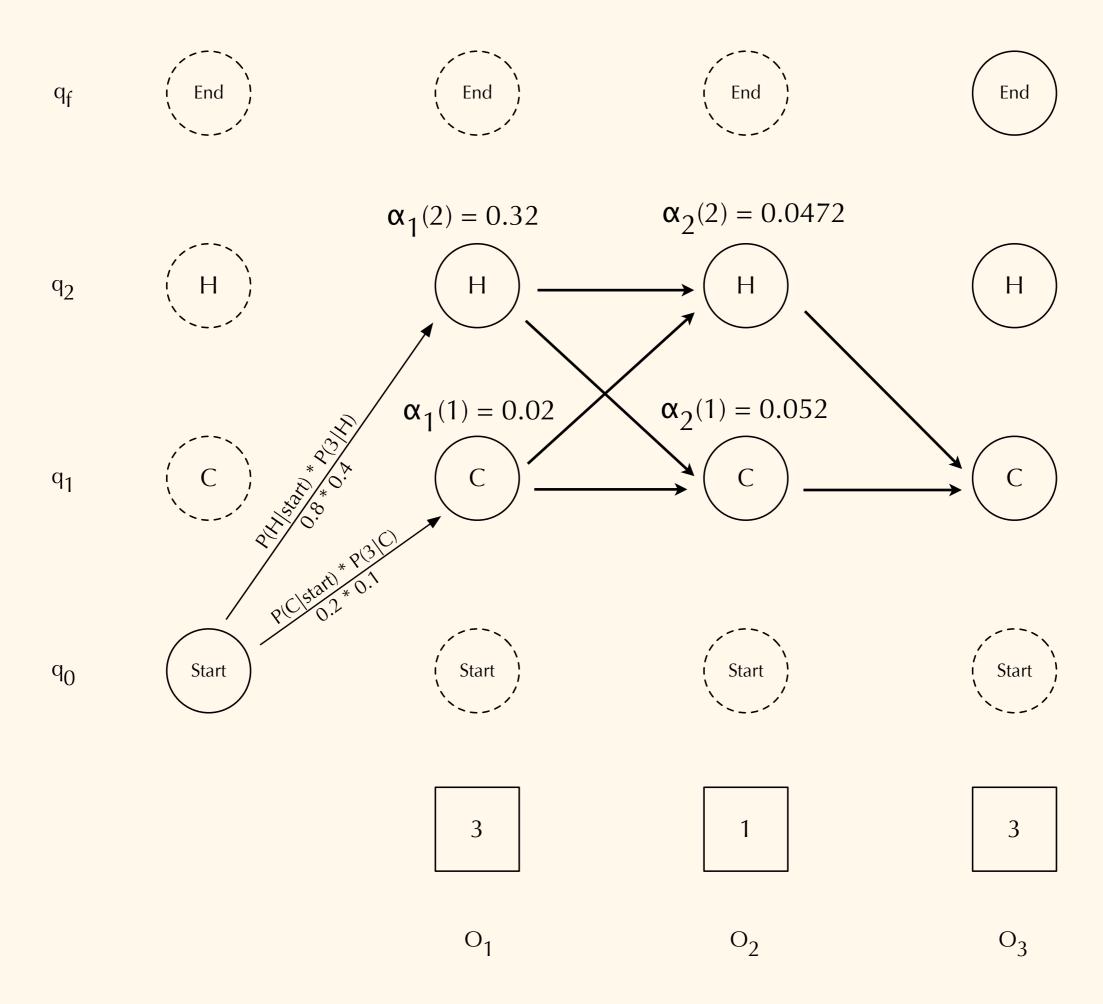


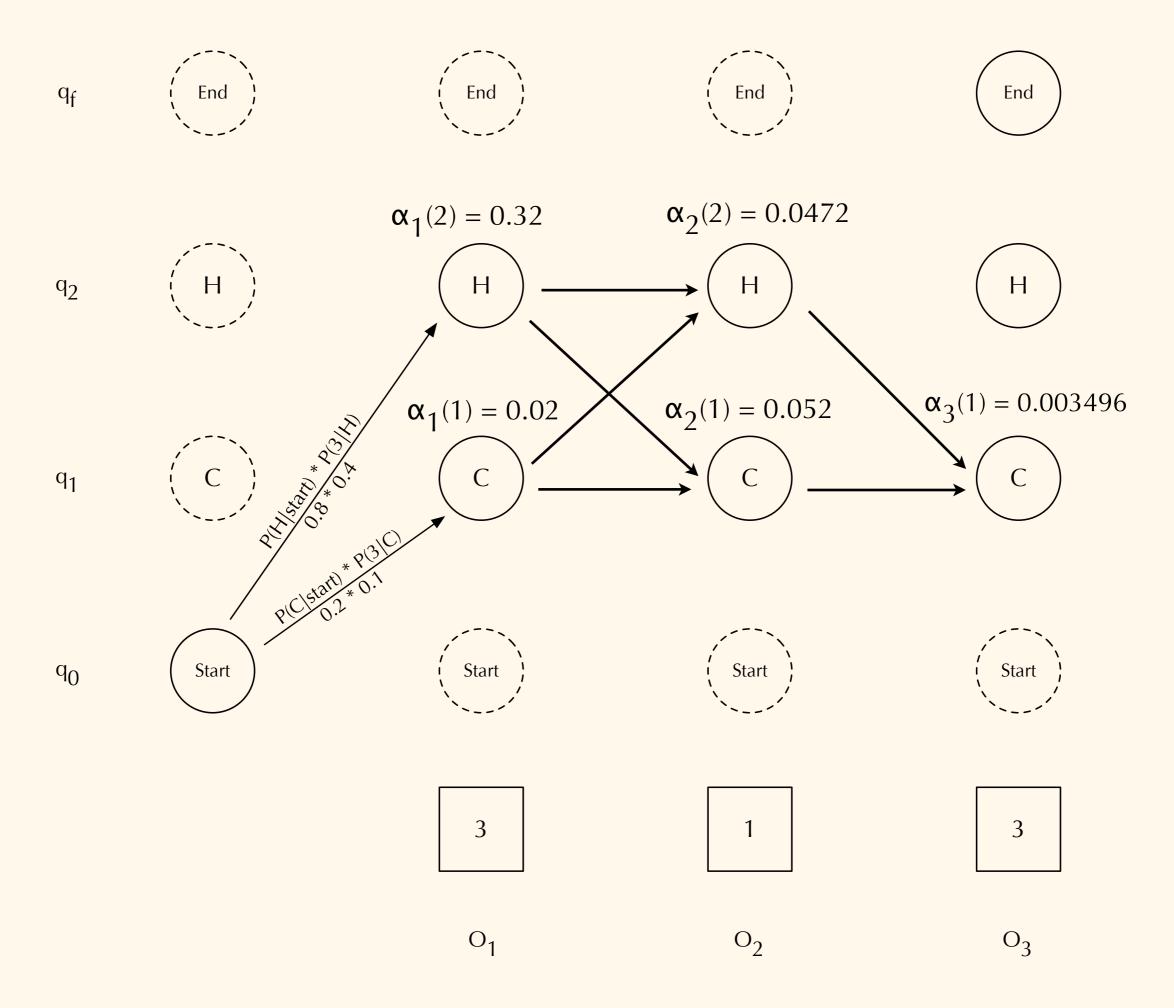


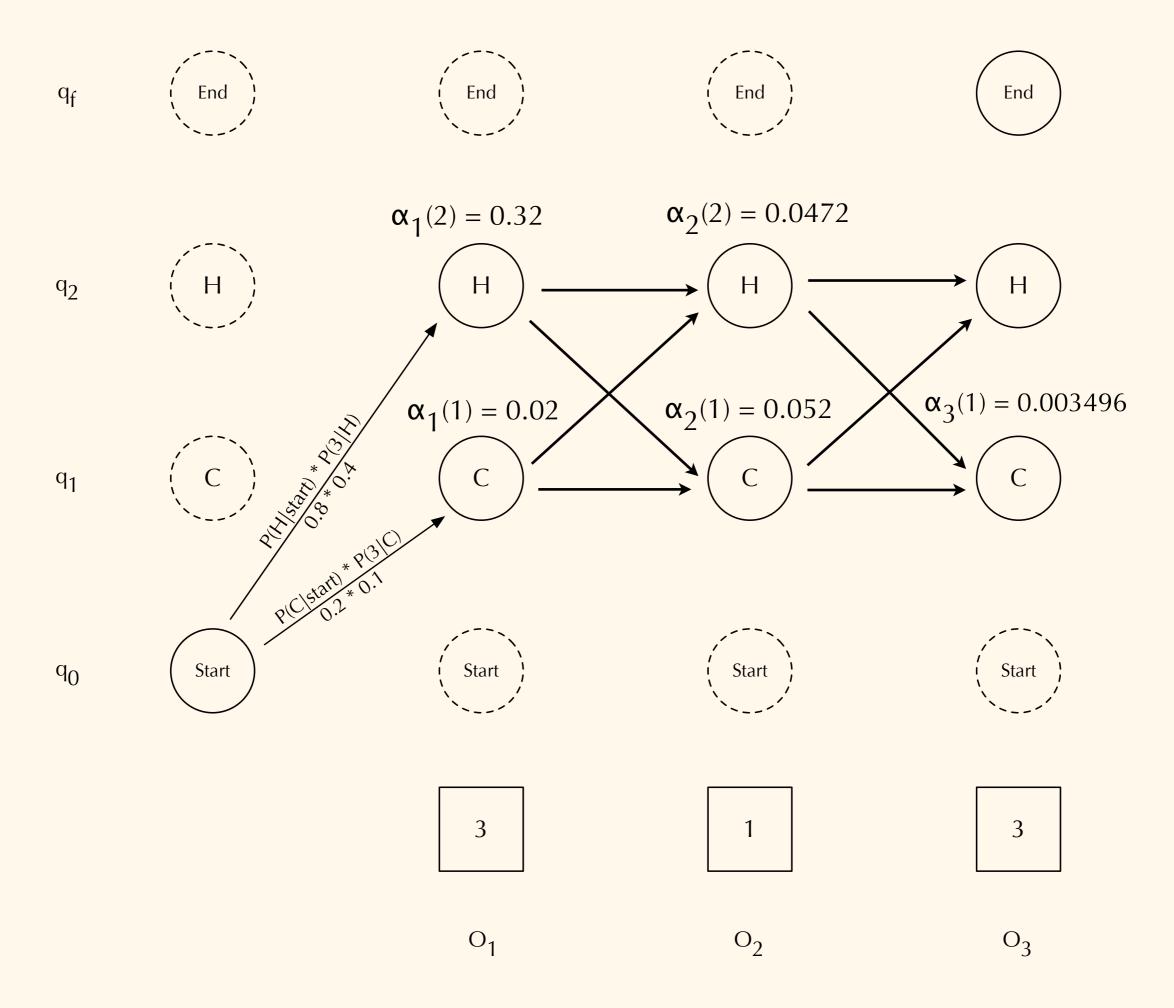


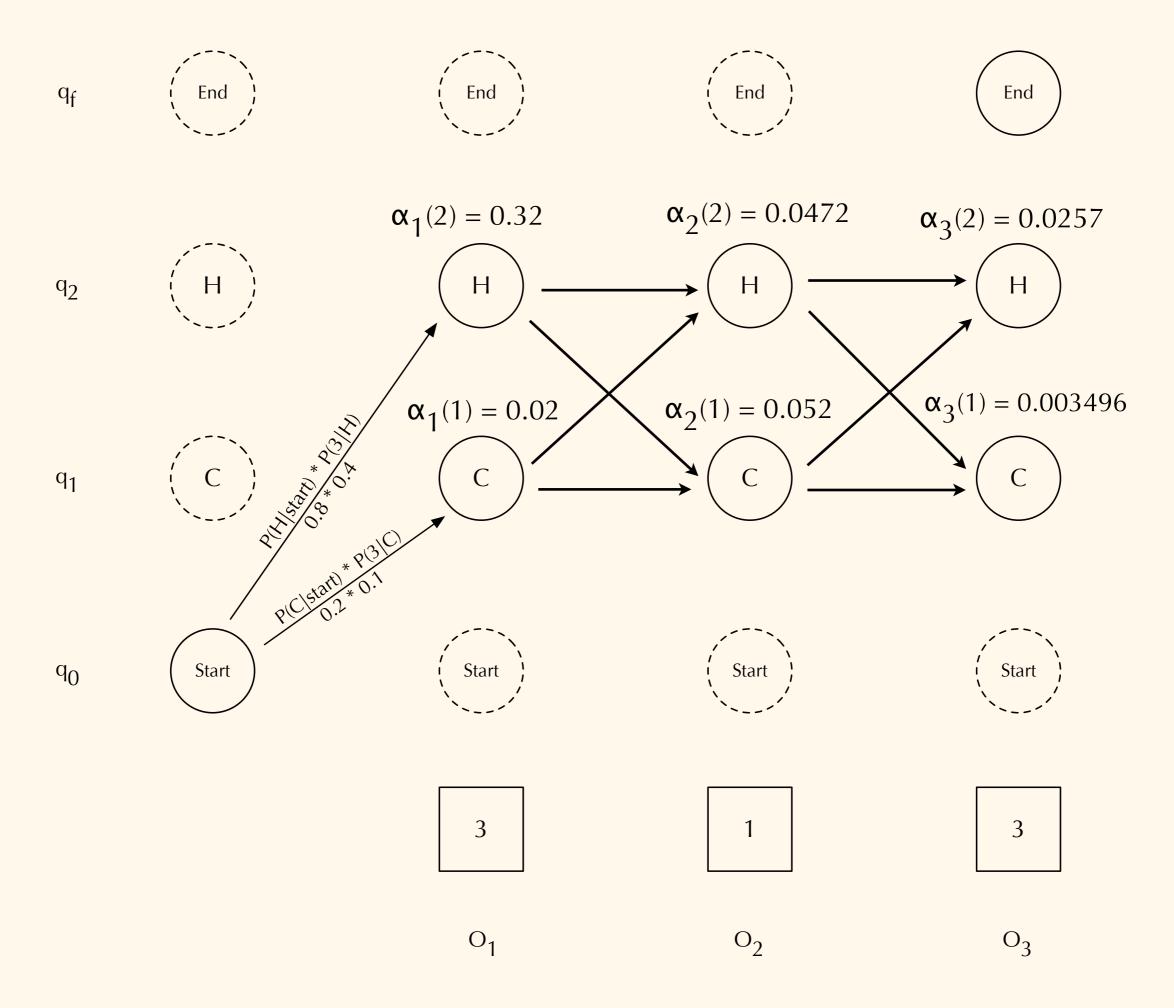


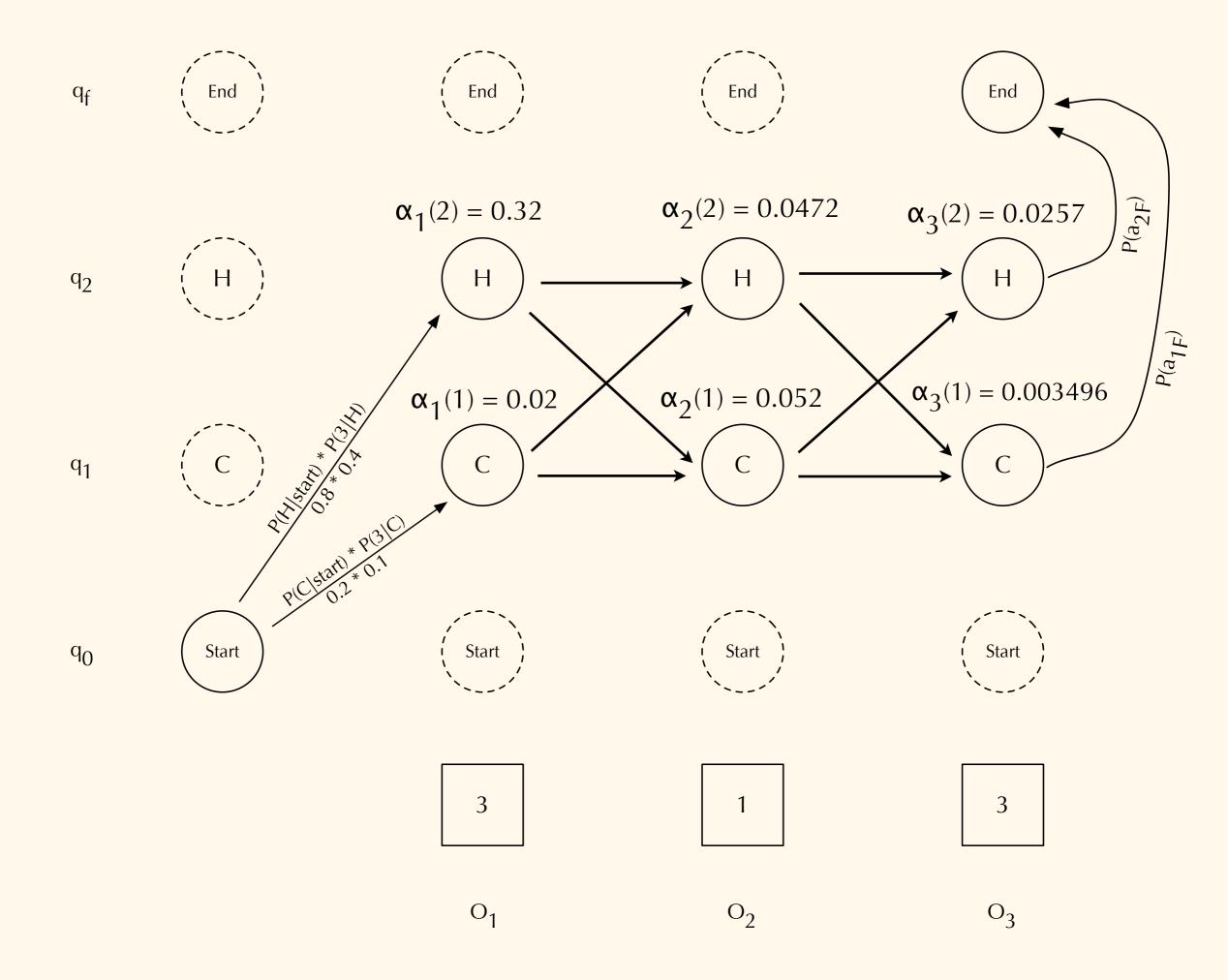


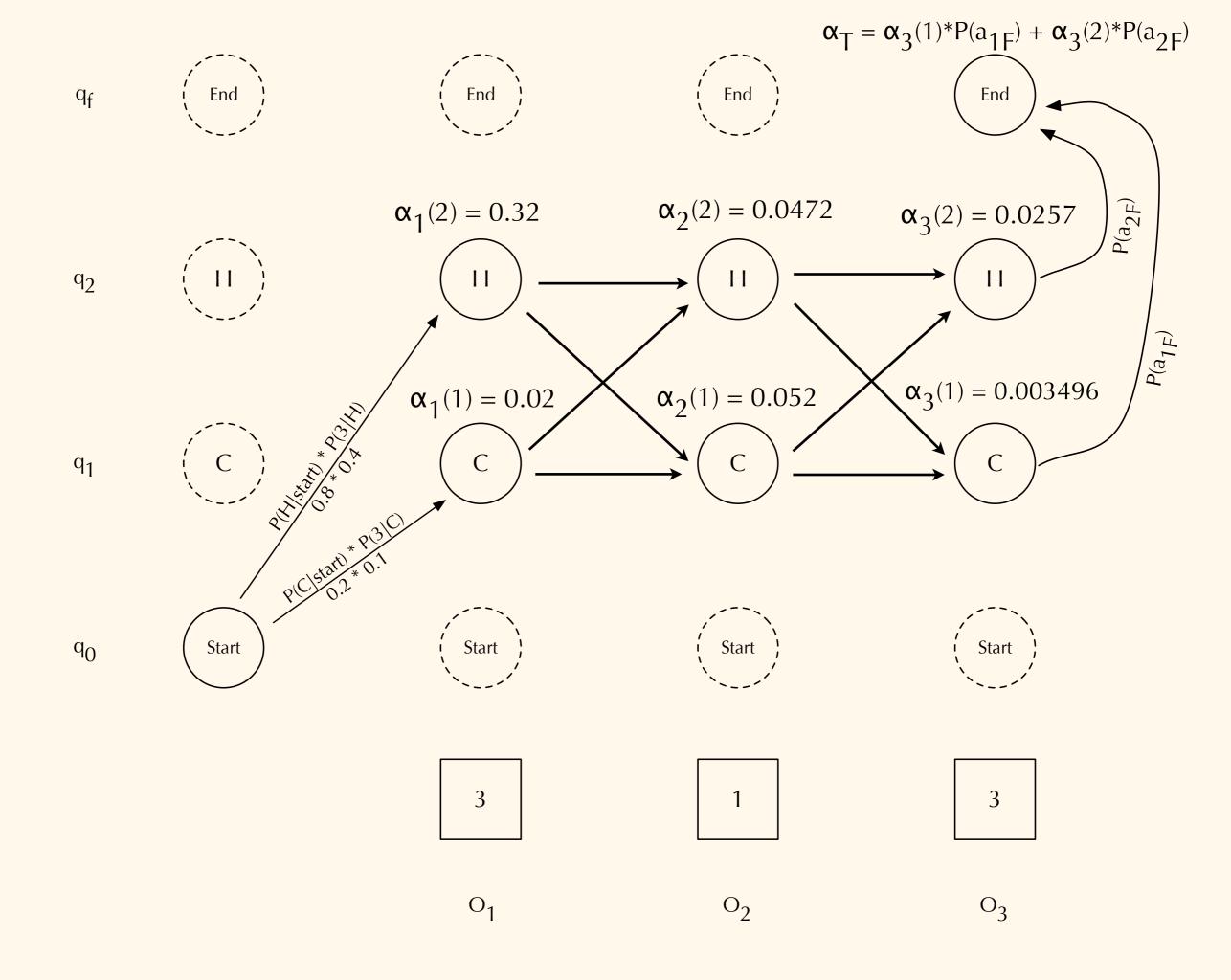


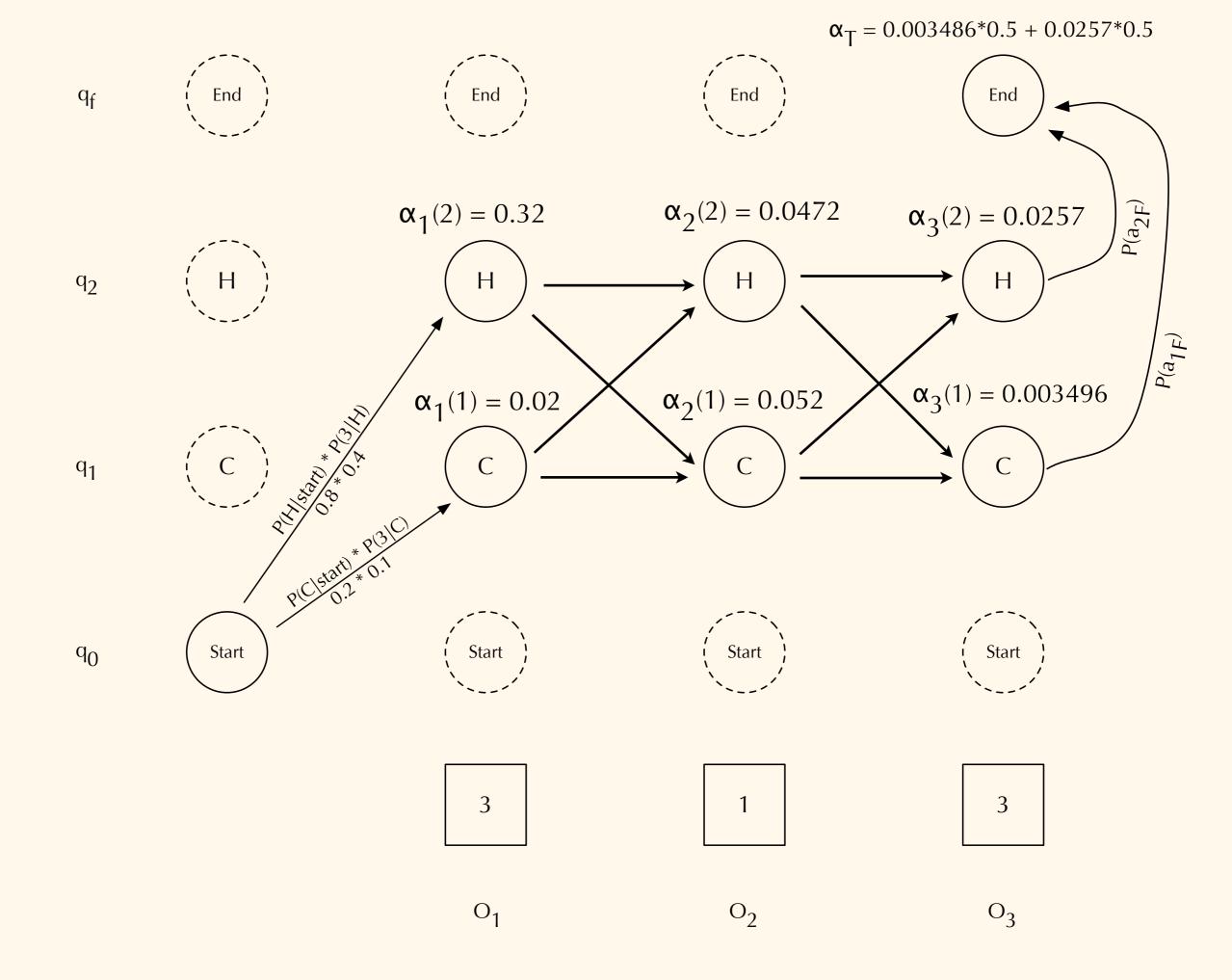


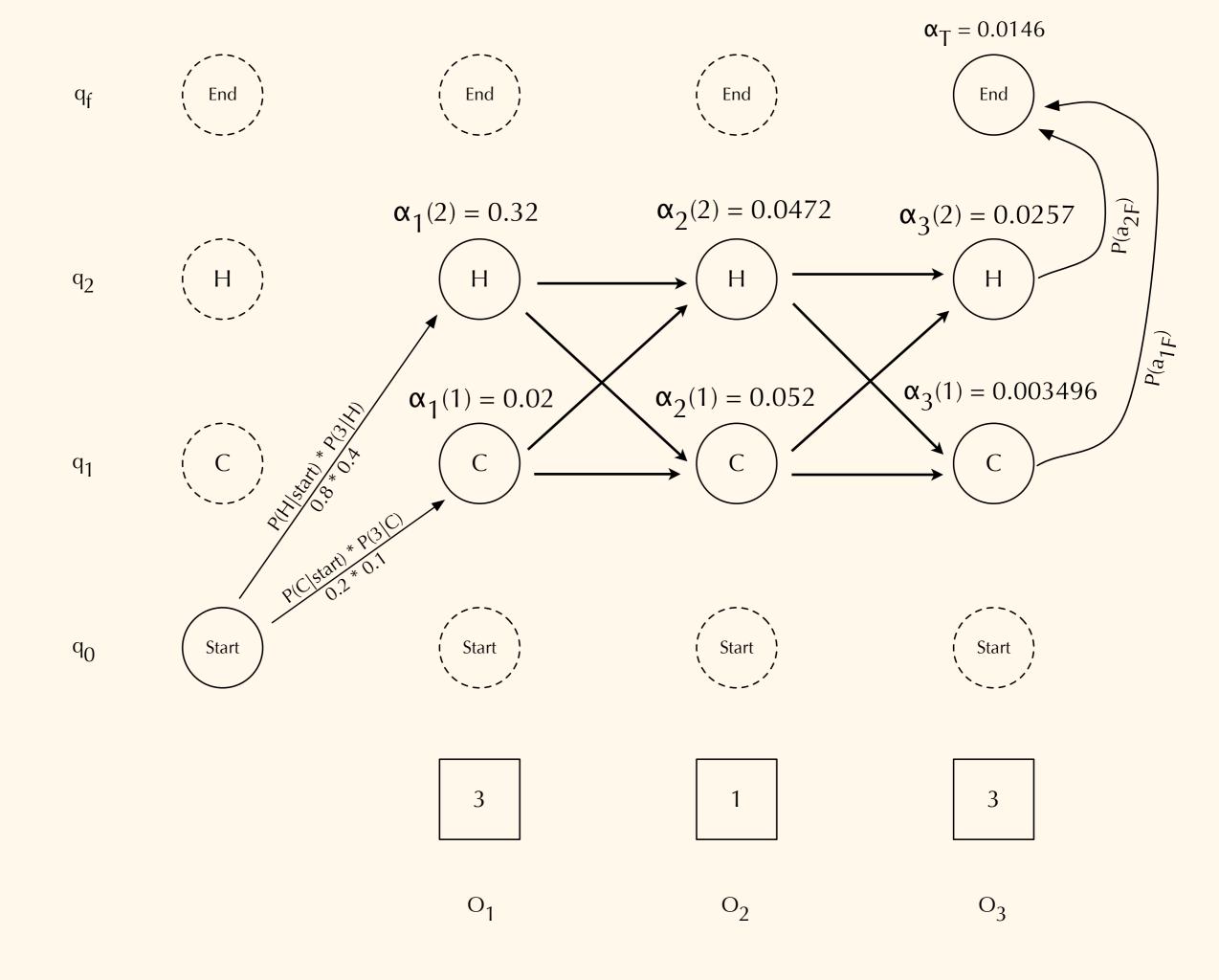












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O(N^T) possible solutions...

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Modifying the forward algorithm slightly gives us the *Viterbi algorithm* for decoding.

The main difference: instead of *summing* possible paths to each state, we take the *max*...

... and keep track of which one it was!

$$\alpha_t(j) = \sum_{i=1}^{N} \alpha_{t-1}(i) a_{ij} b_j(o_t)$$

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Viterbi algorithm trellis locations:

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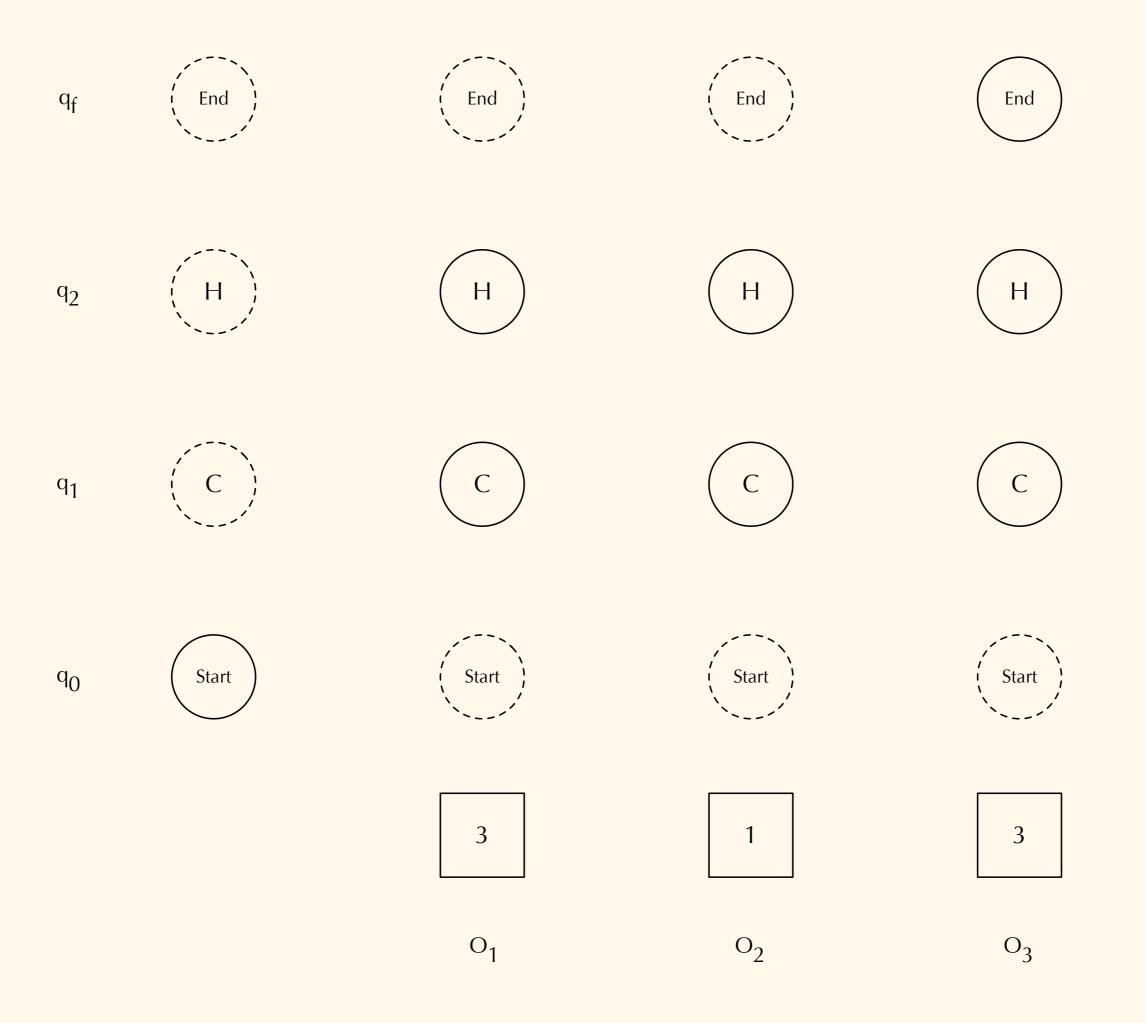
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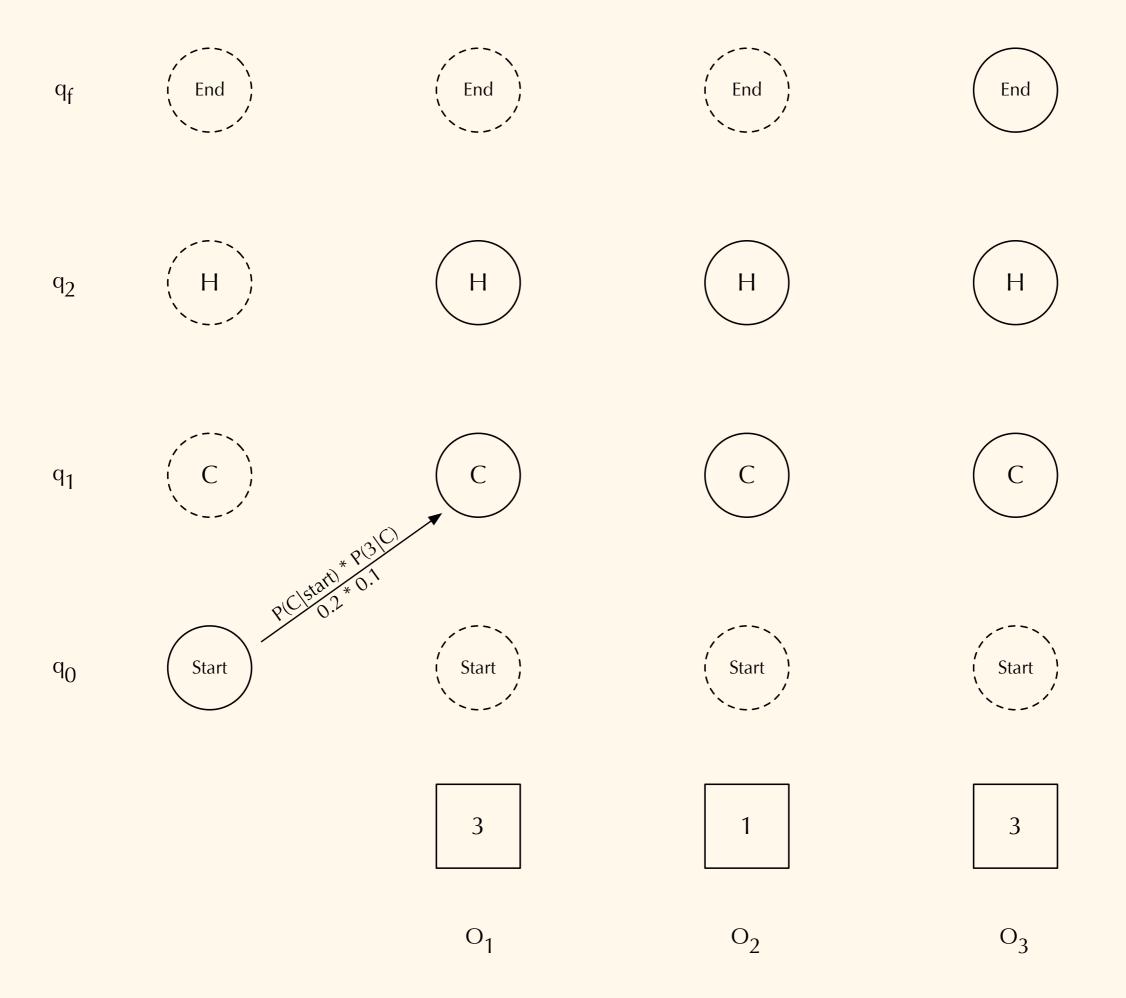
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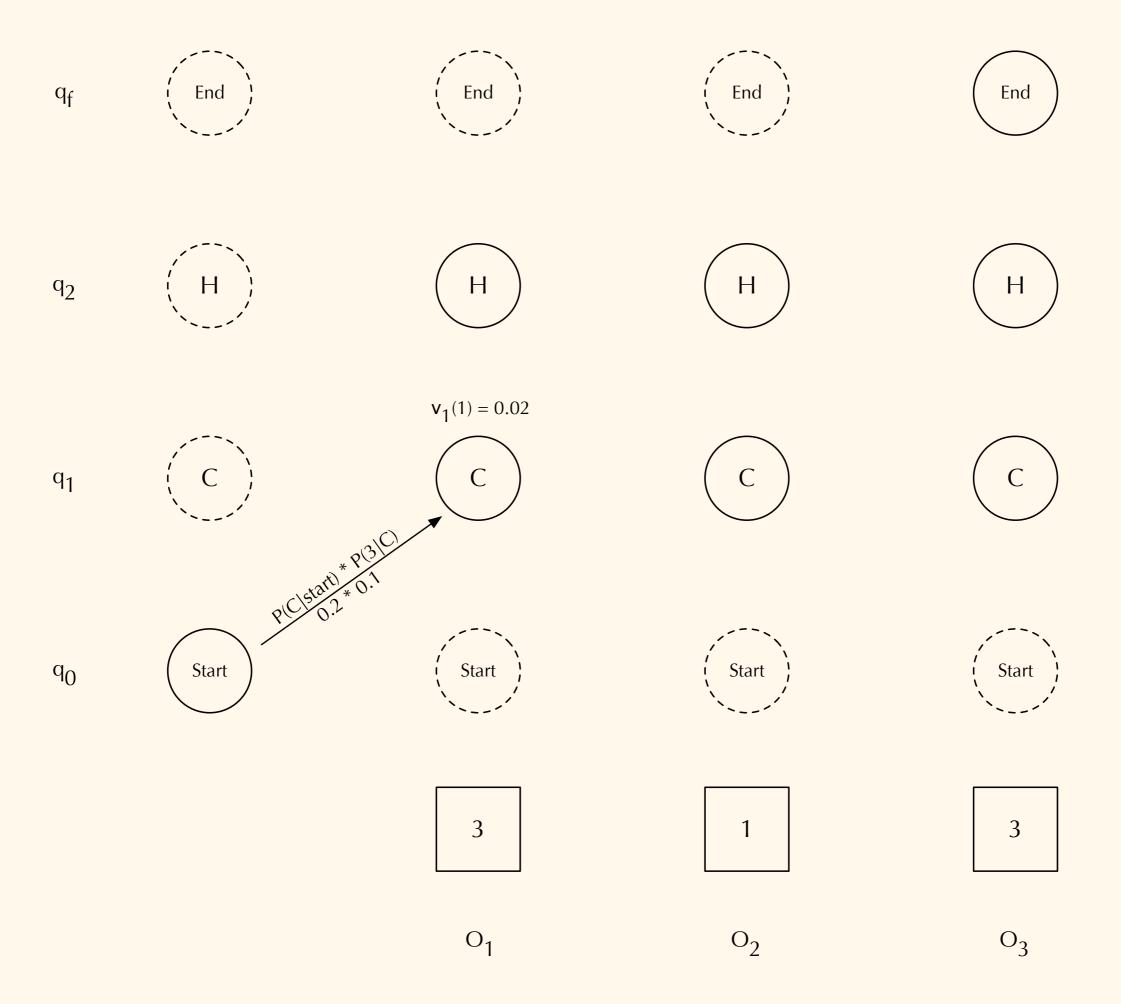
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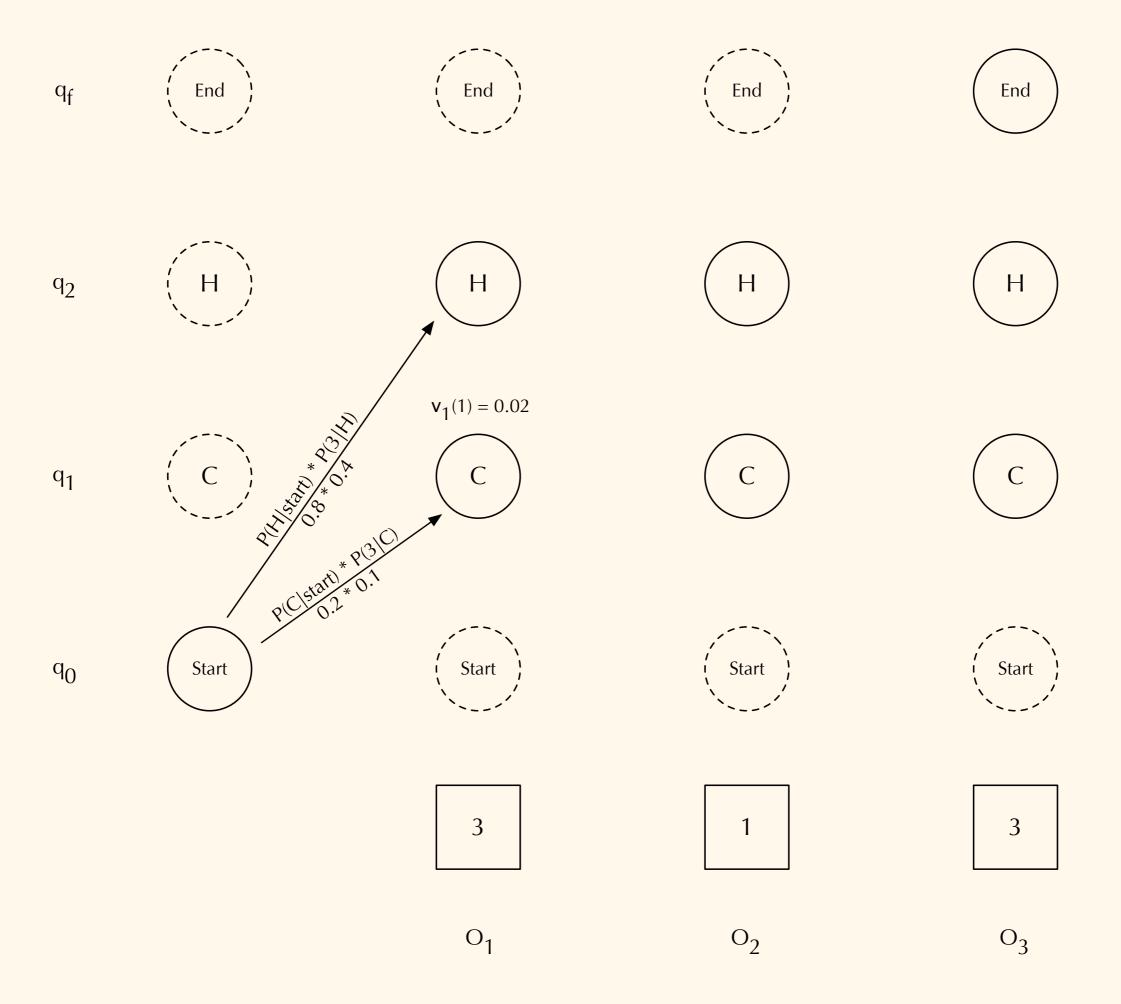
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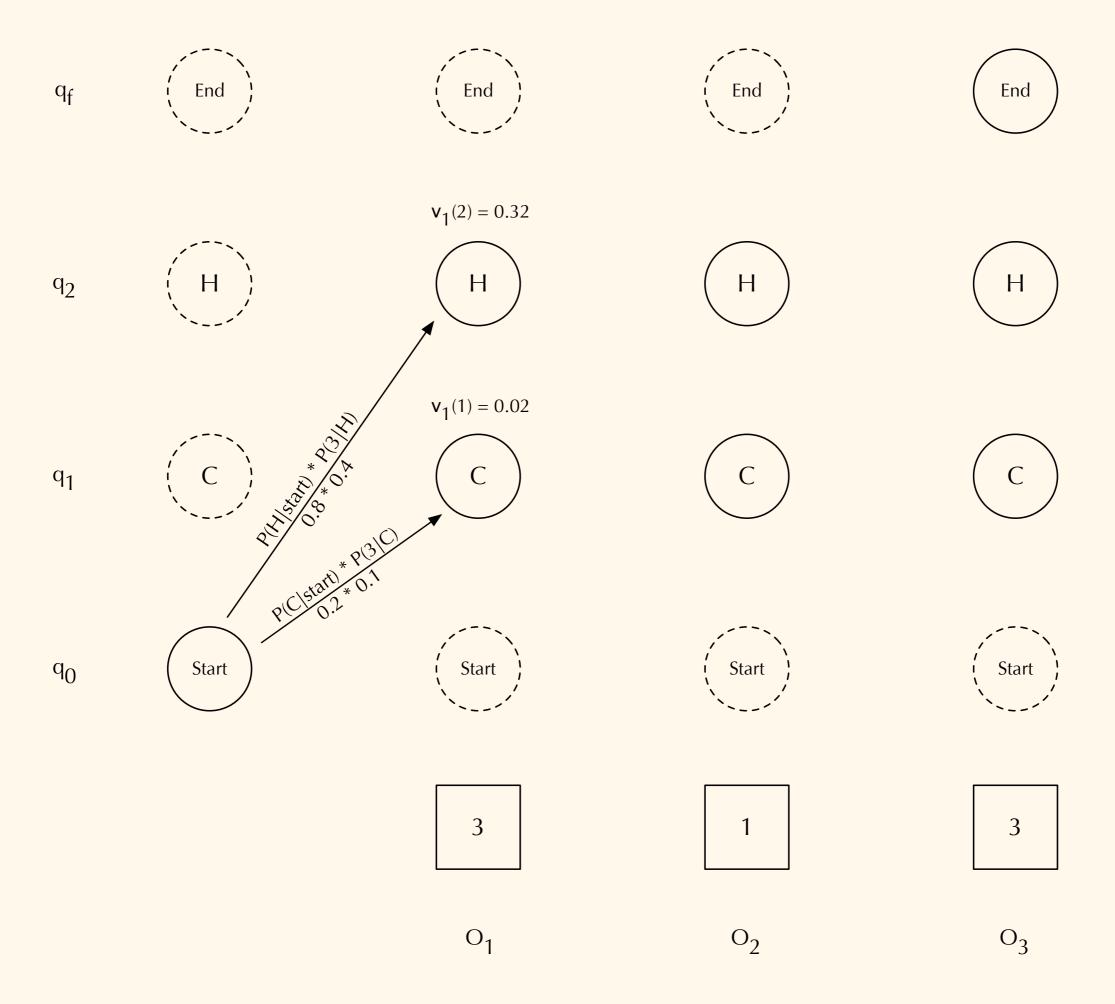
$$bt_t(j) = \underset{i=1}{\operatorname{argmax}} v_{t-1} a_{ij} b_j(o_t)$$

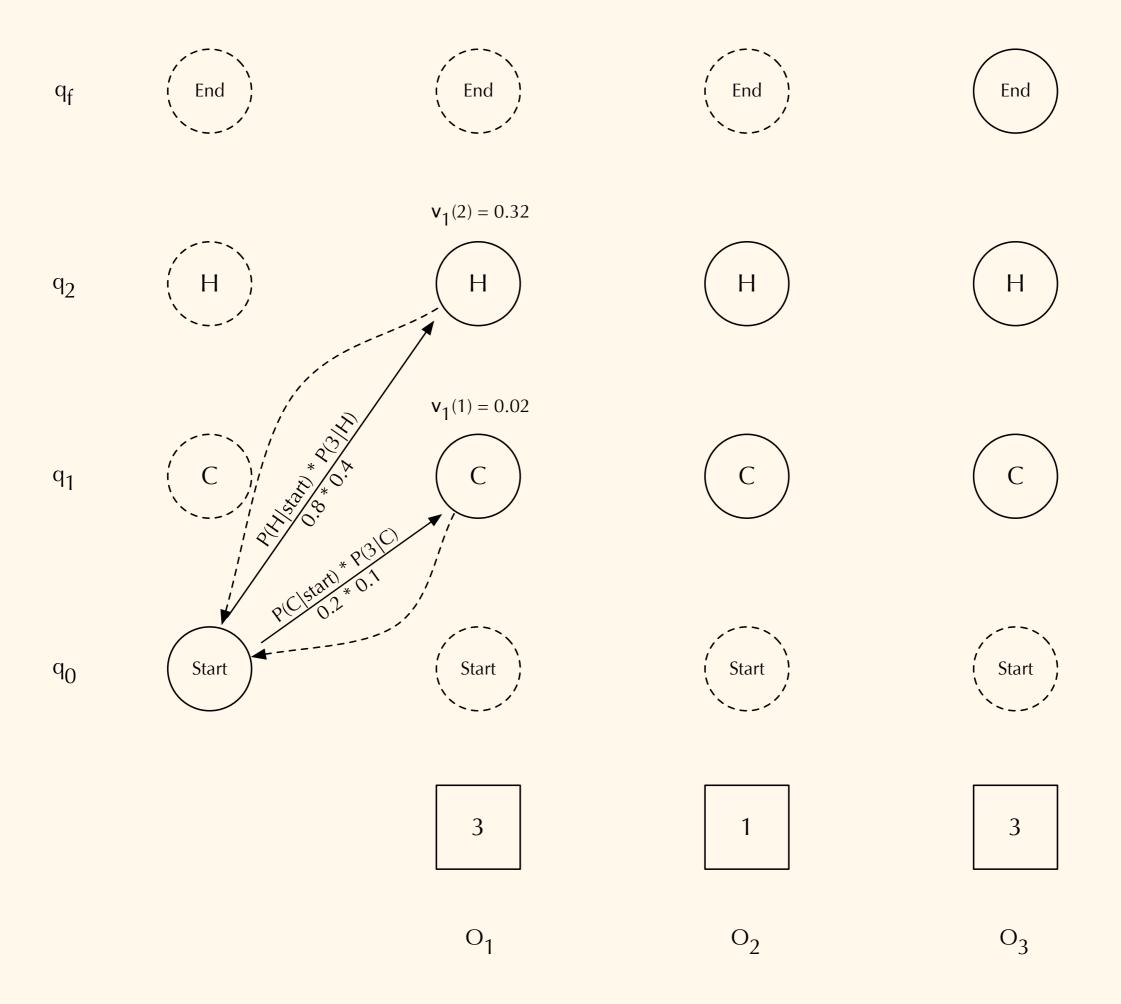


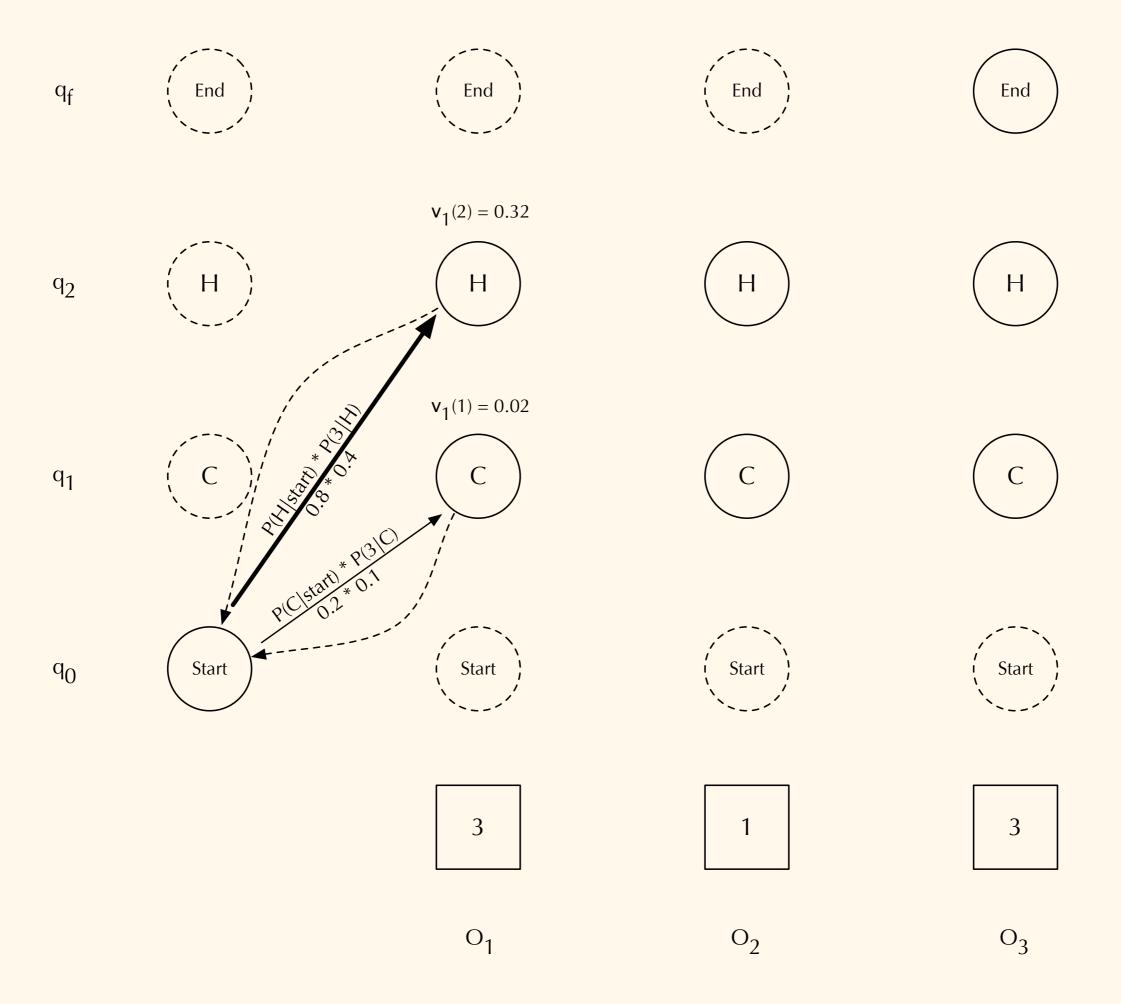


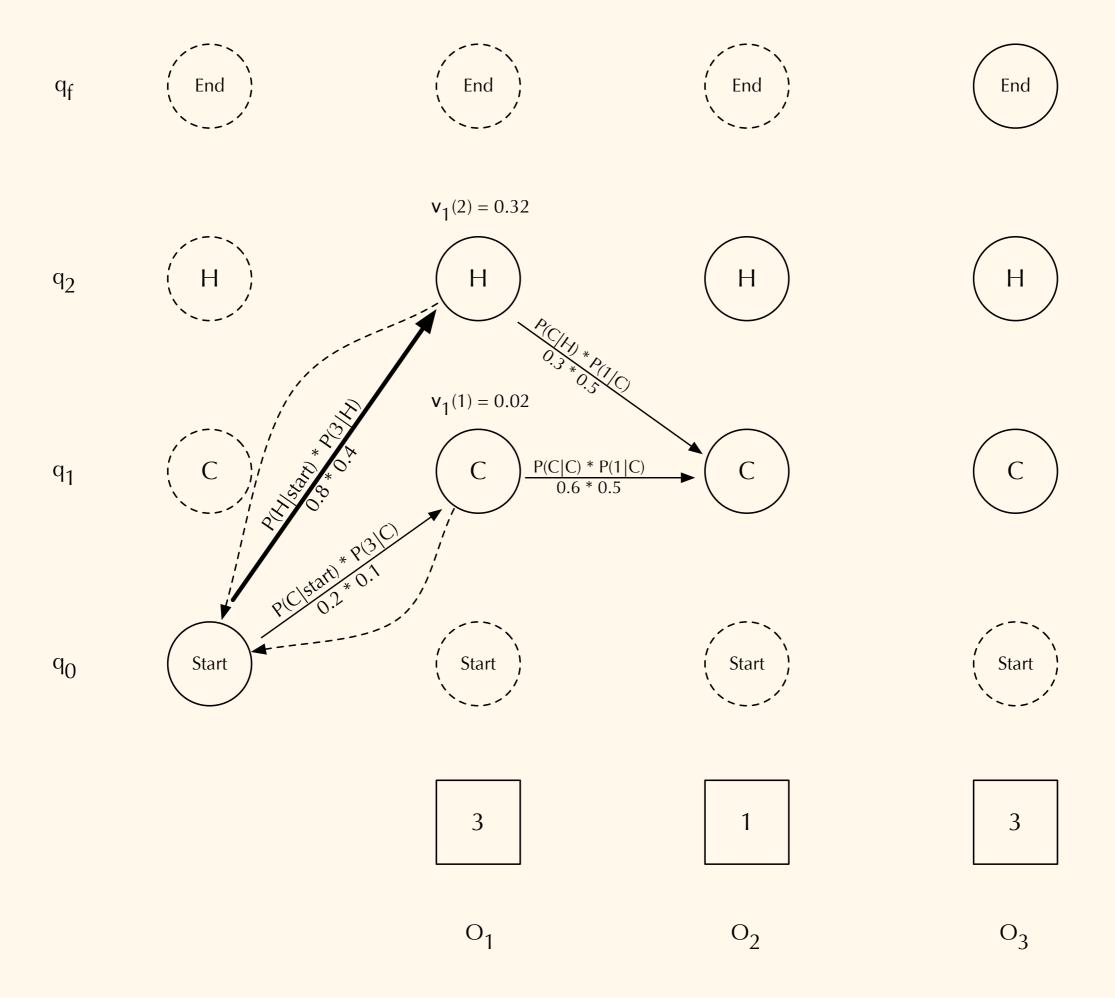


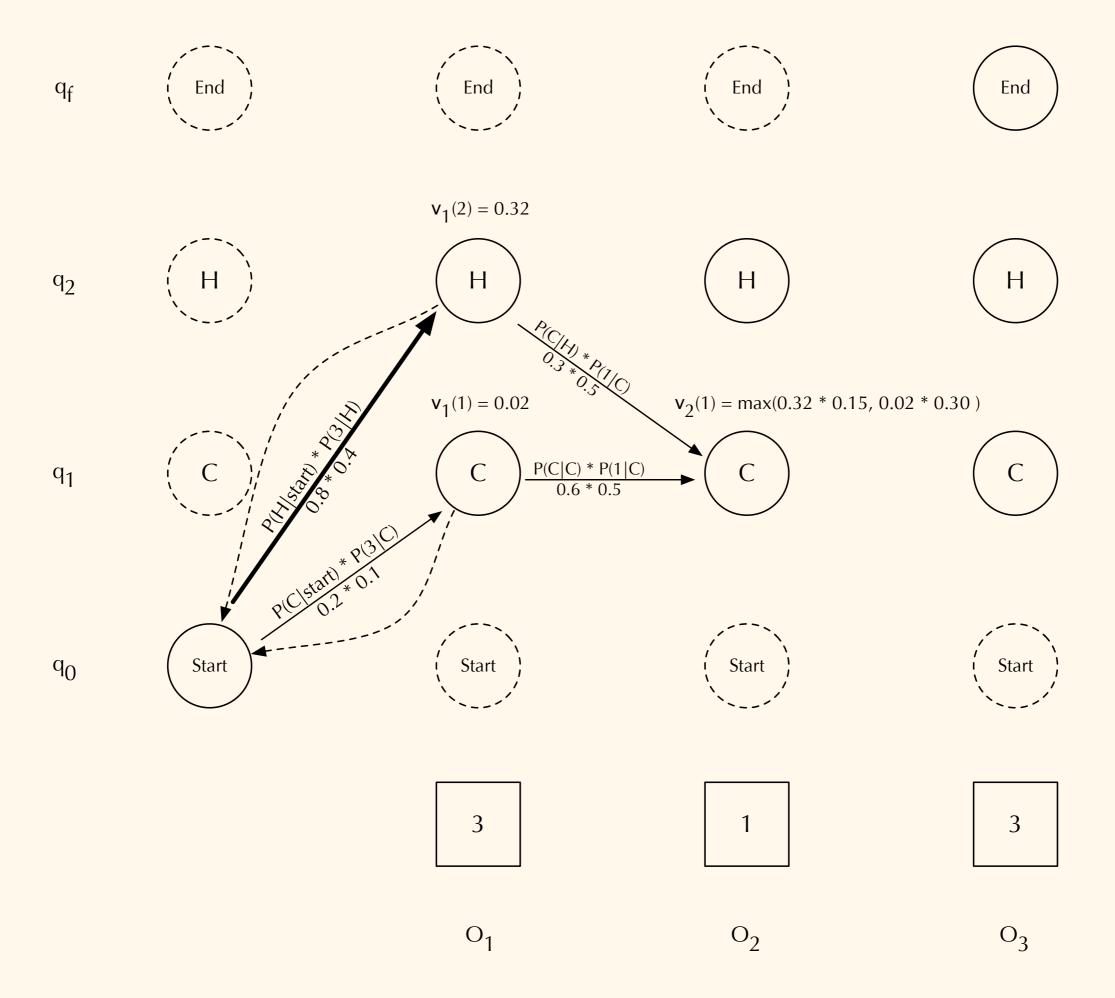


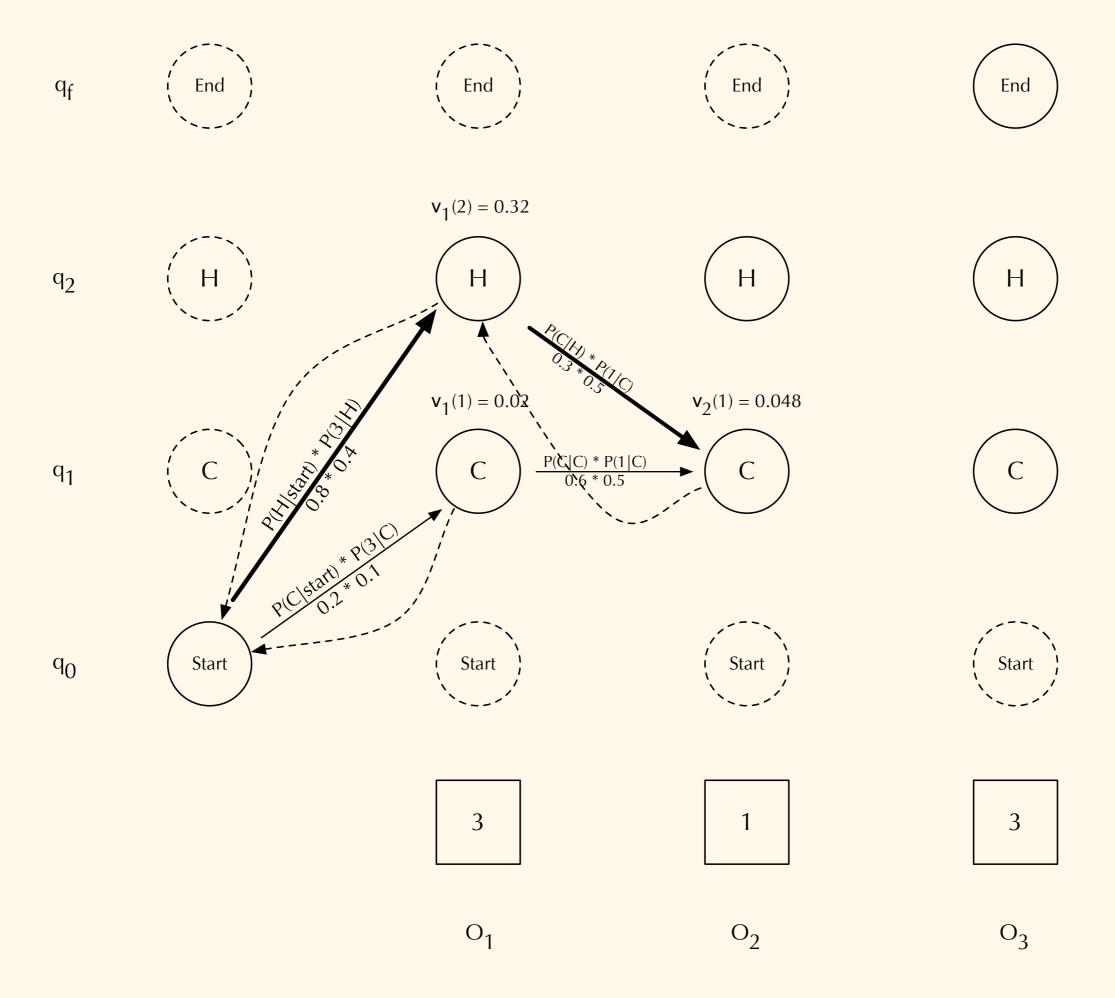


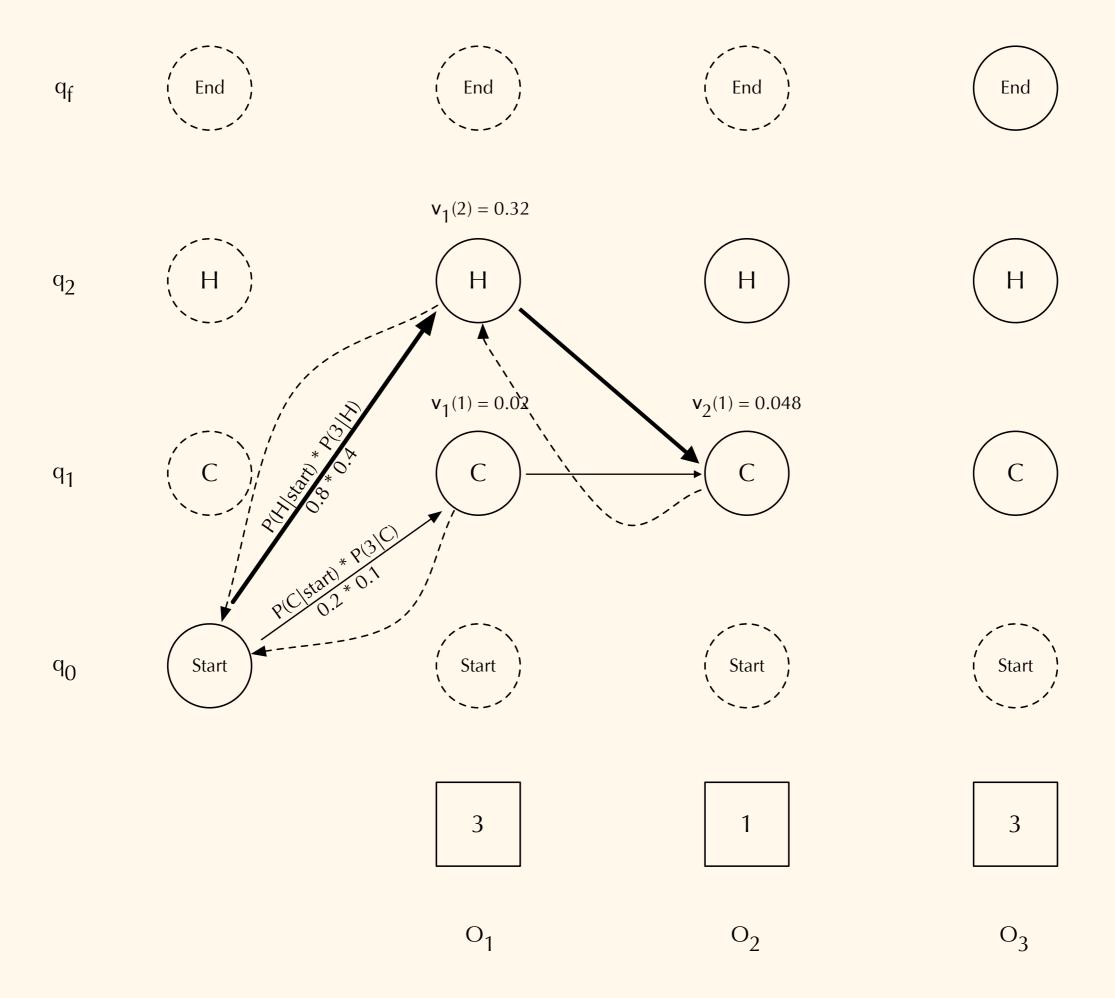


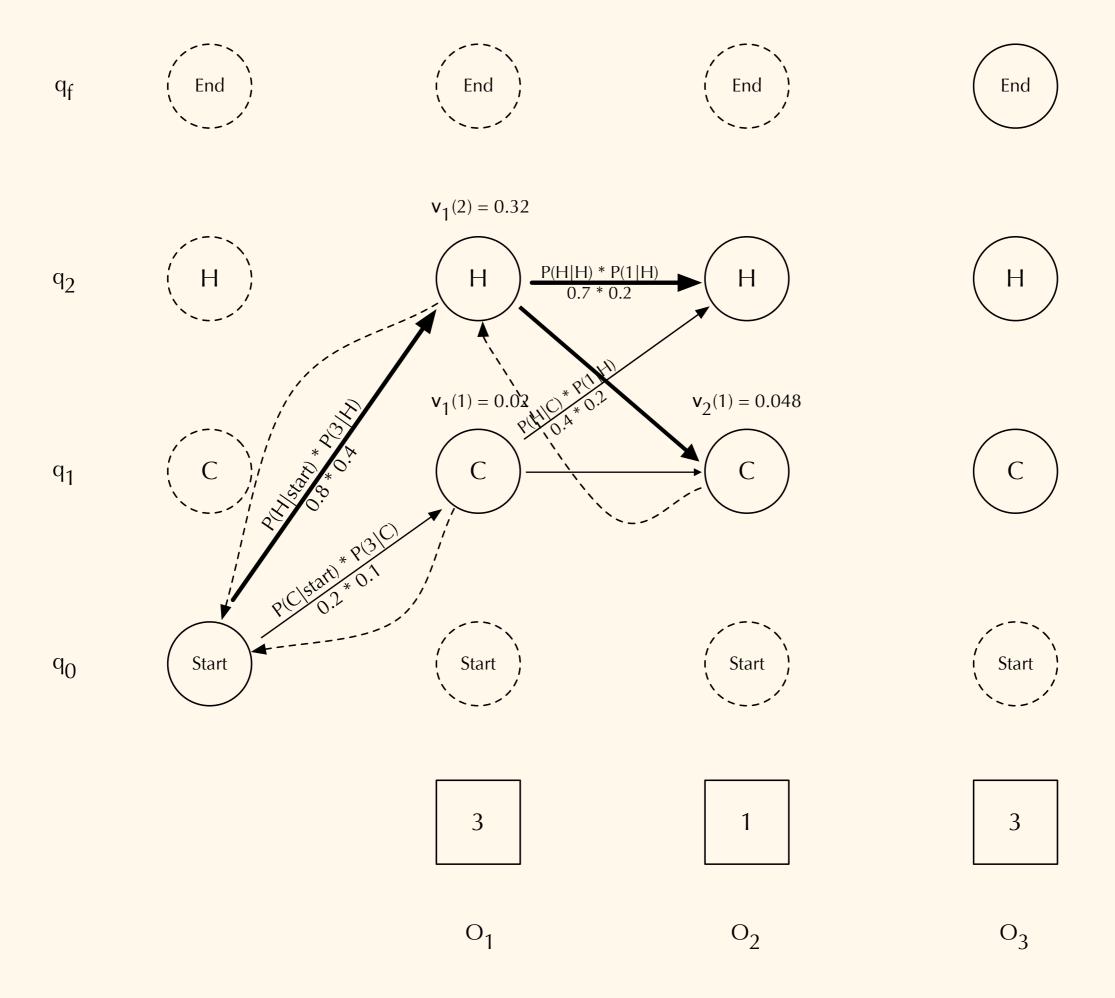


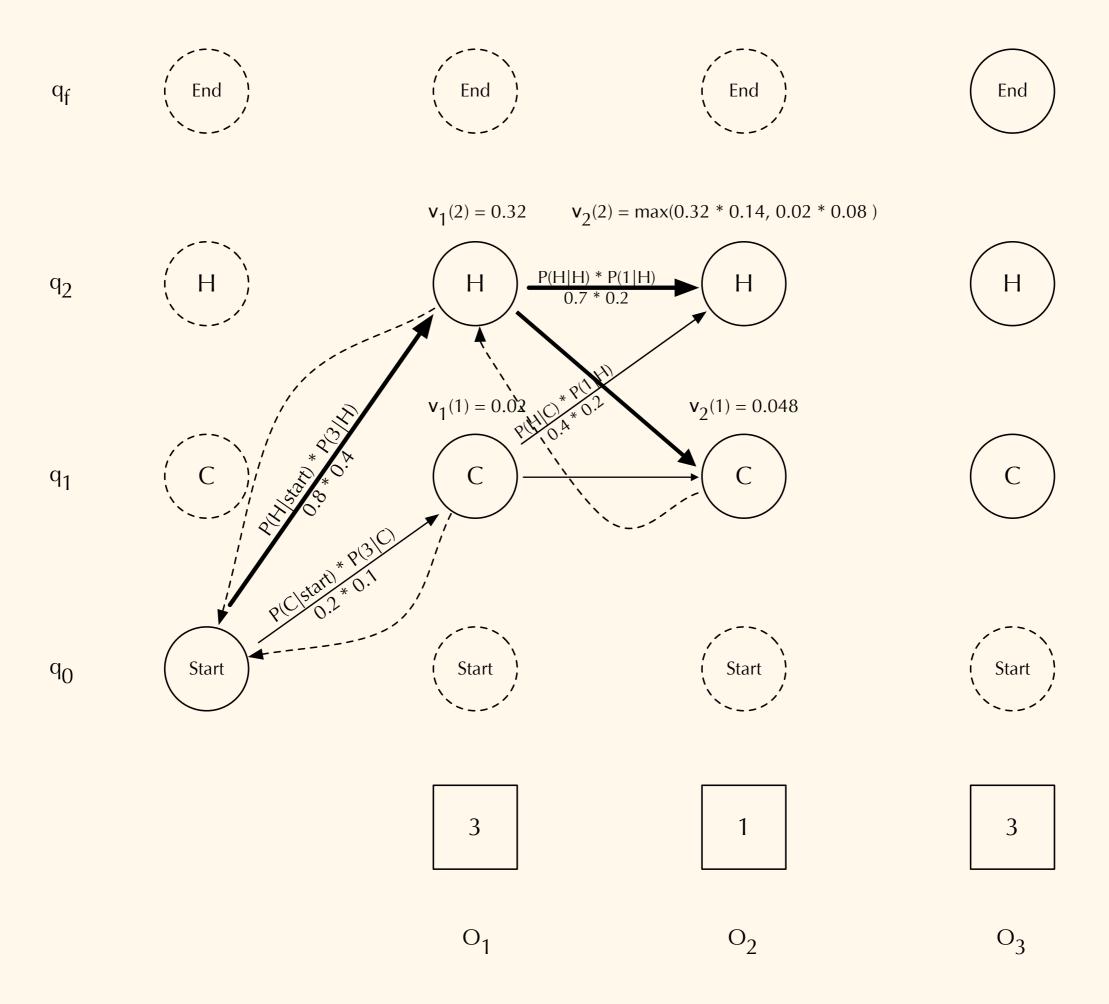


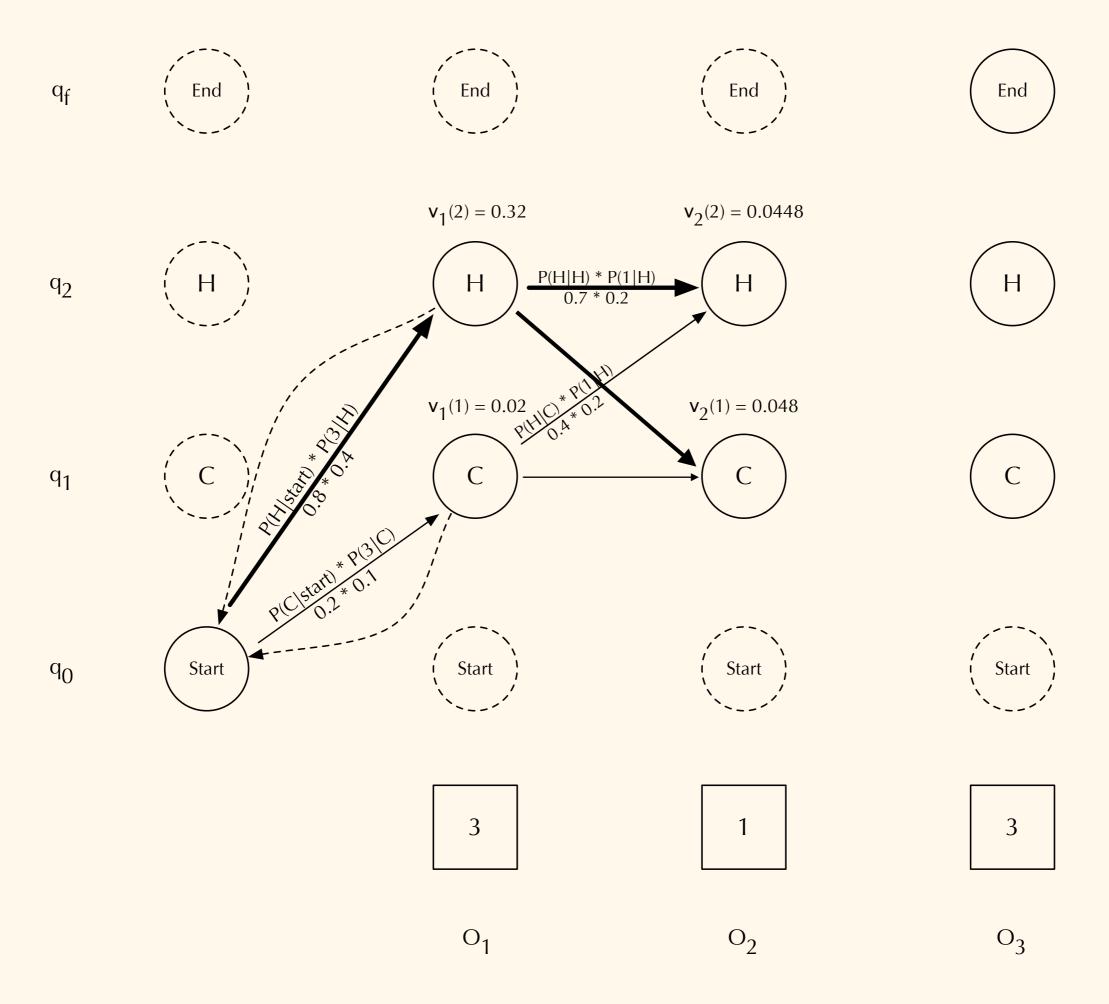


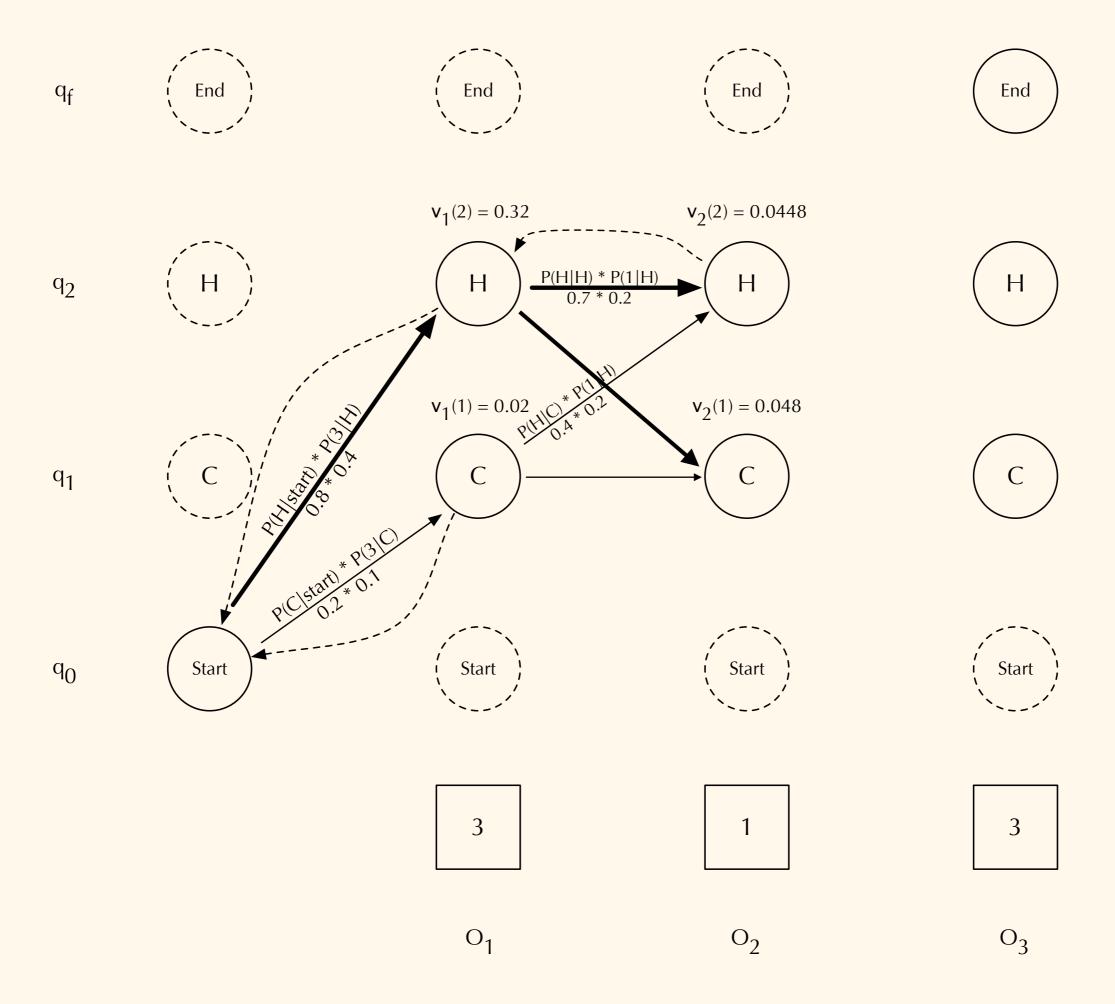


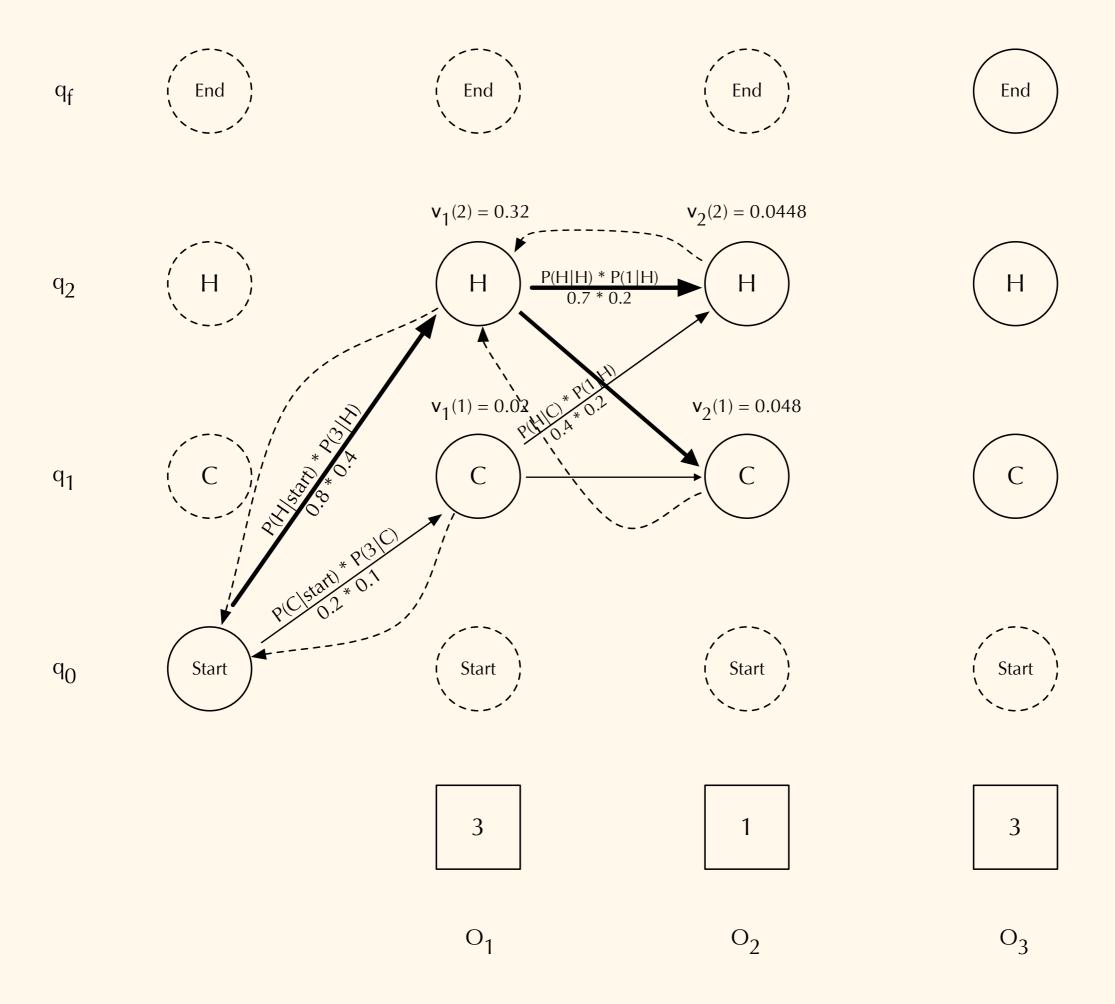


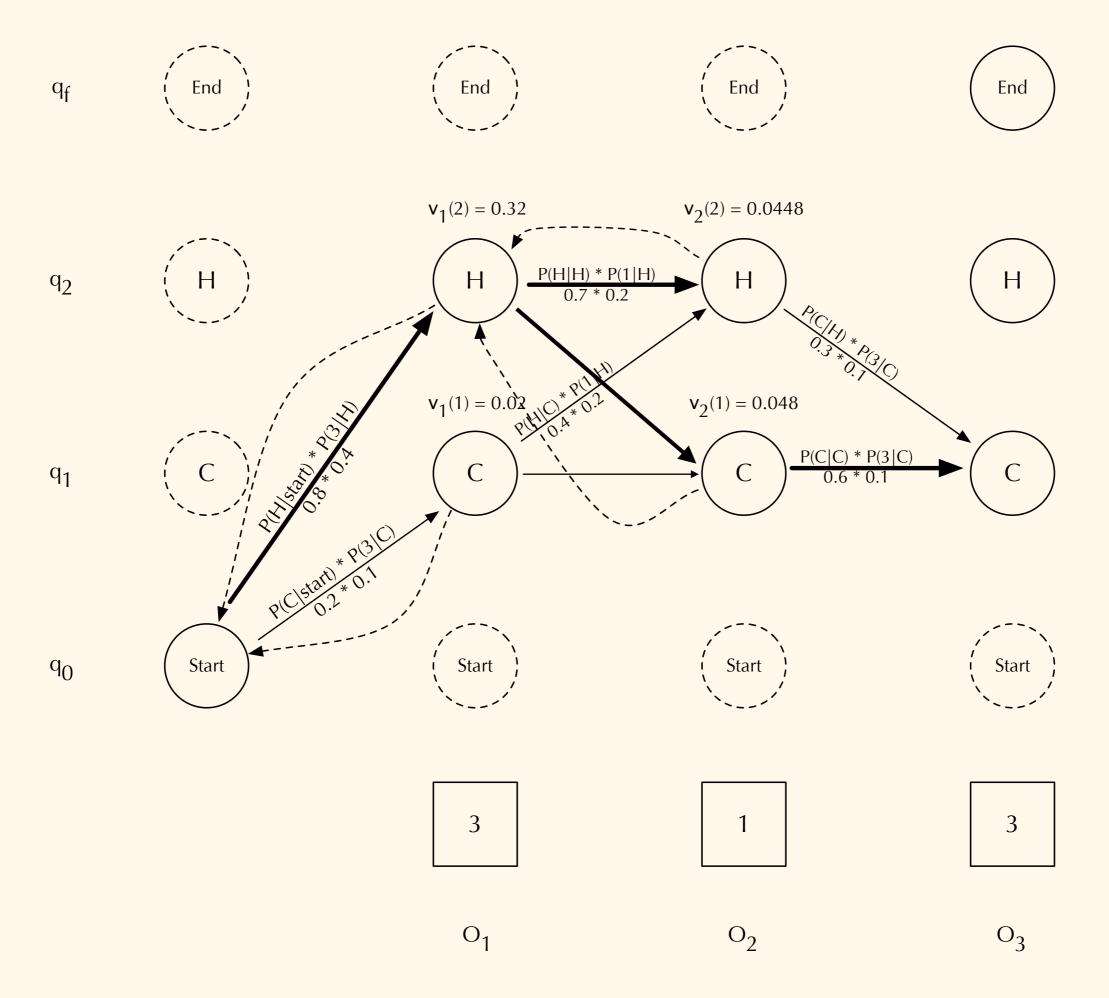


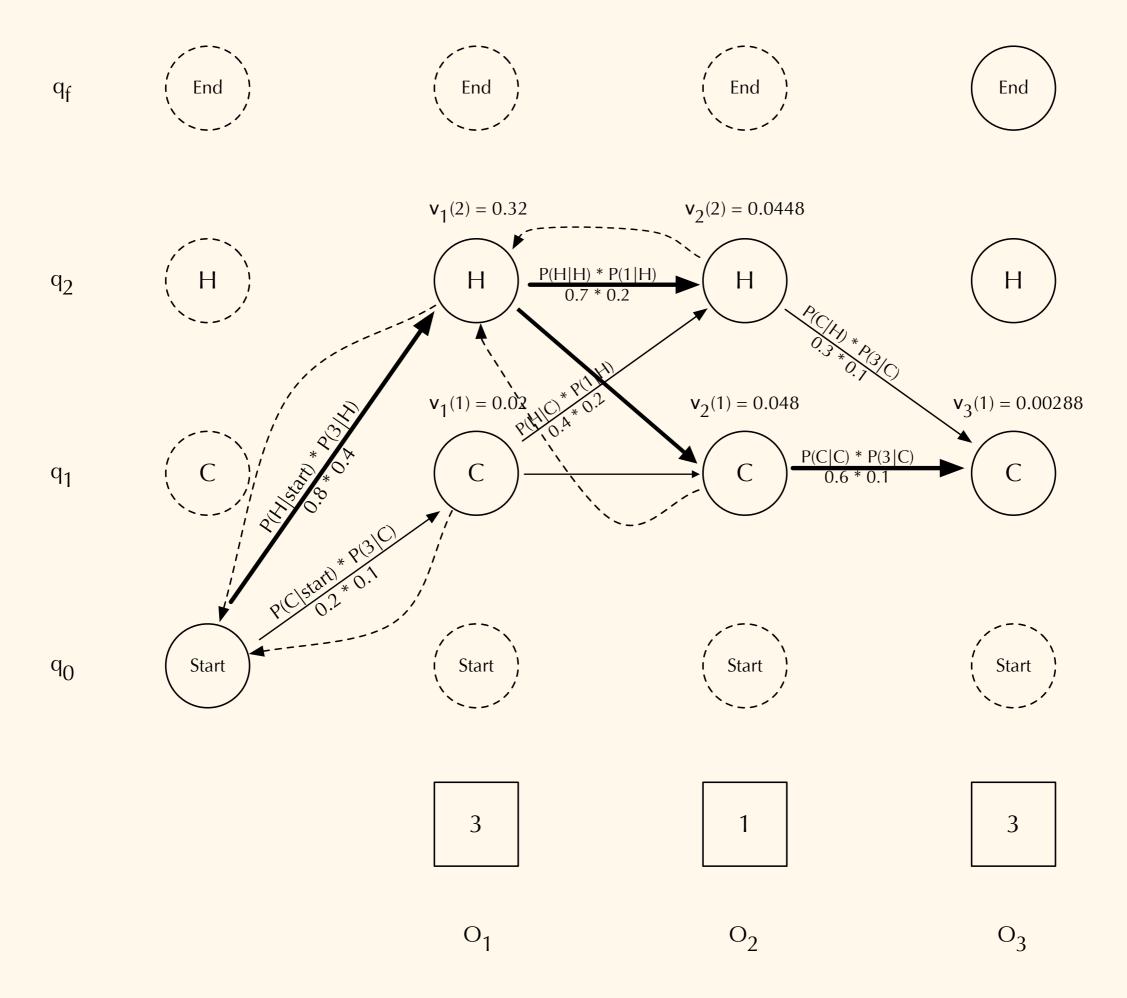


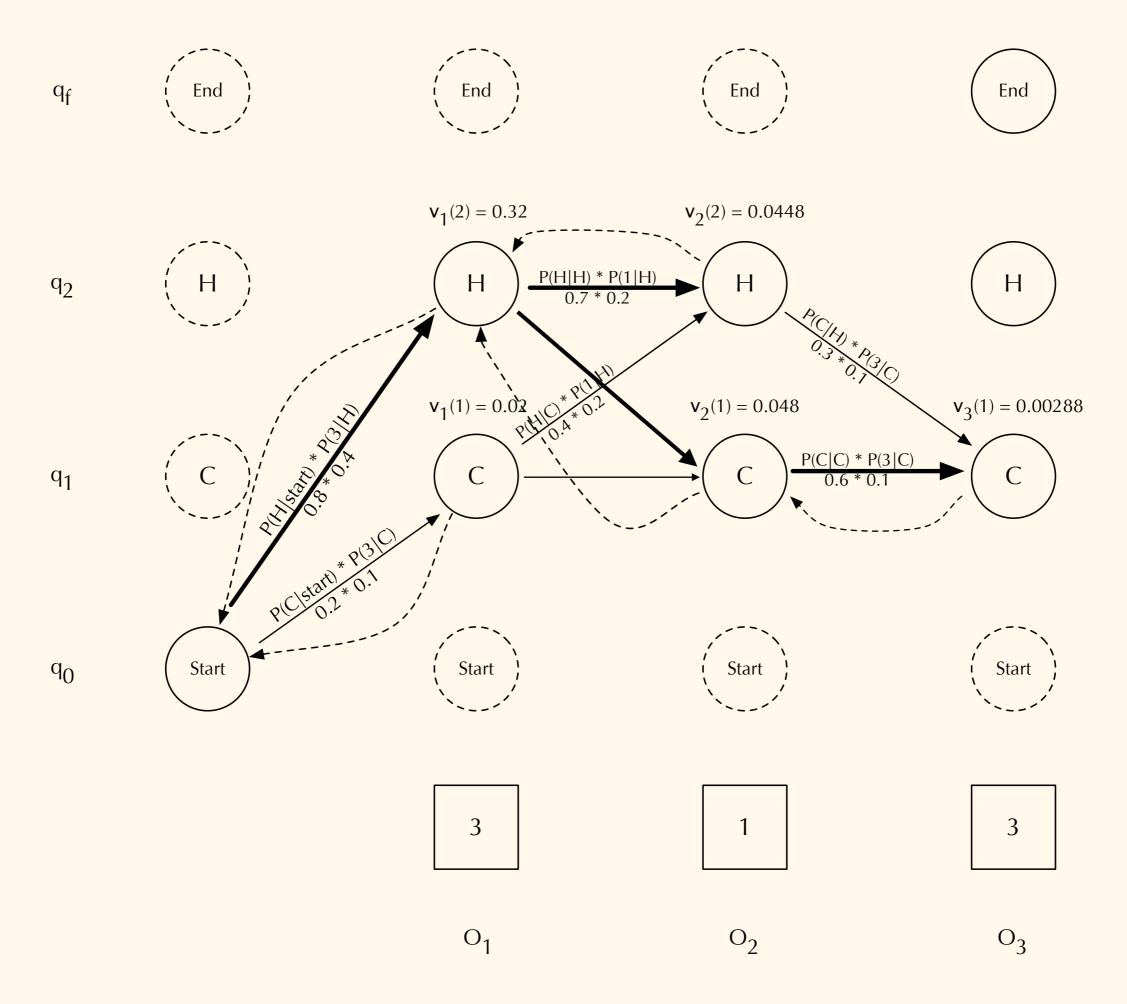


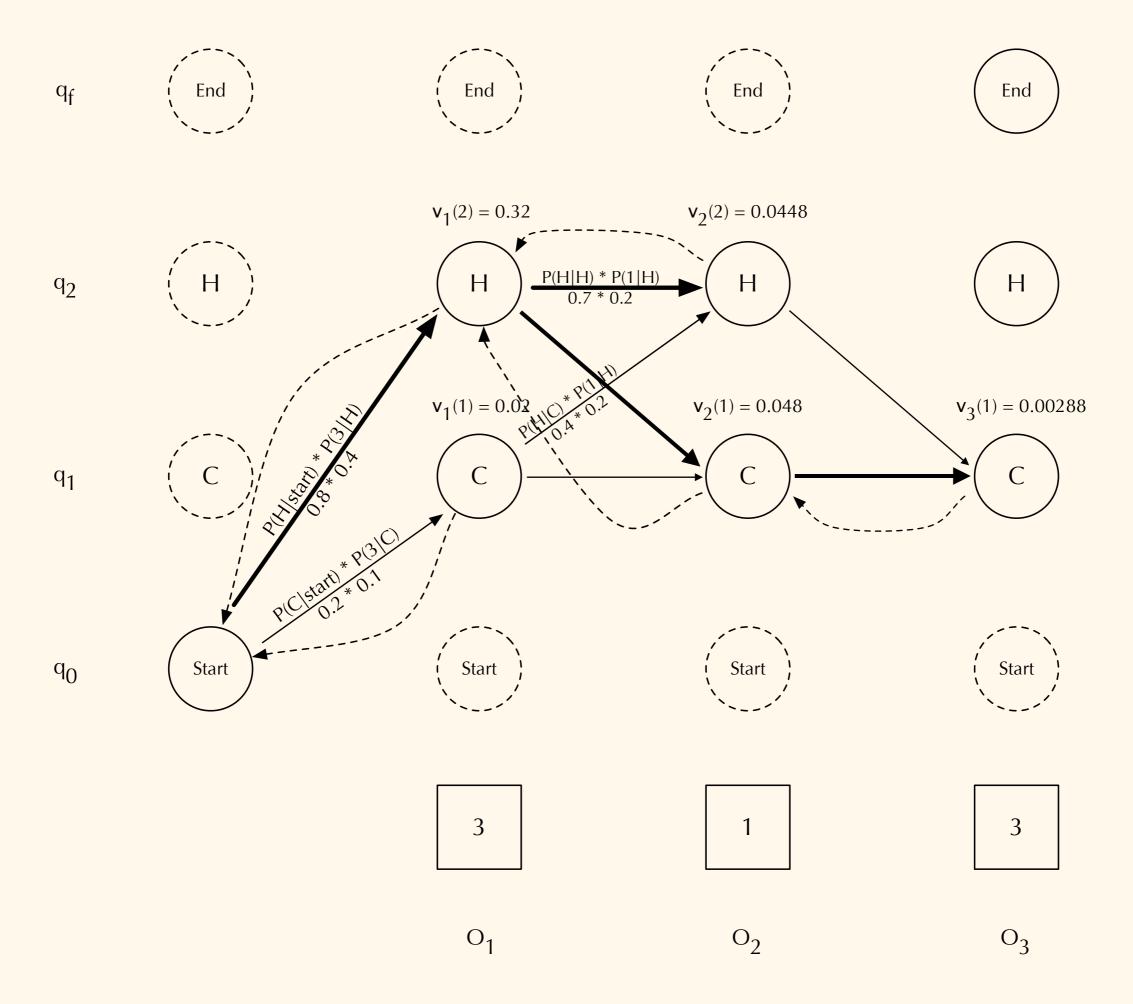


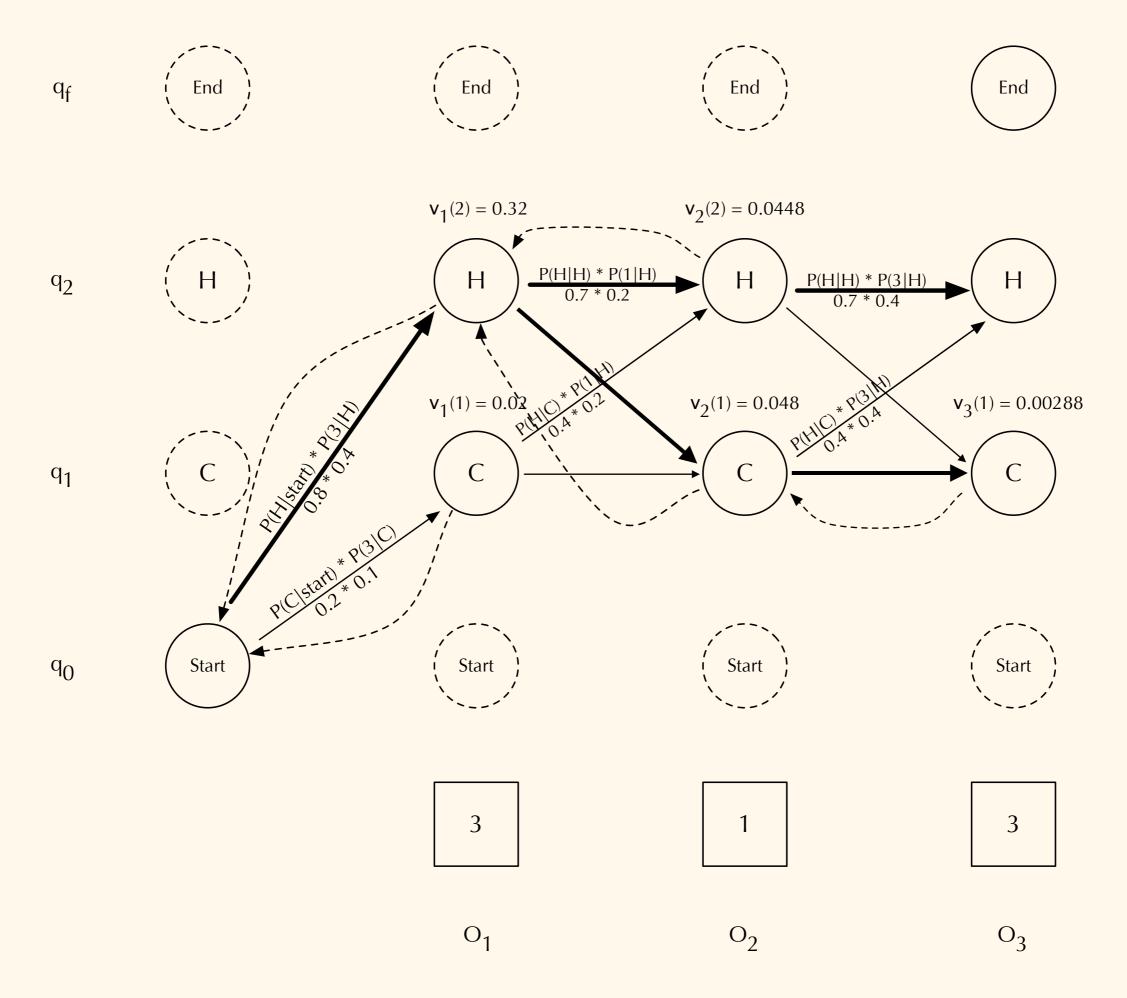


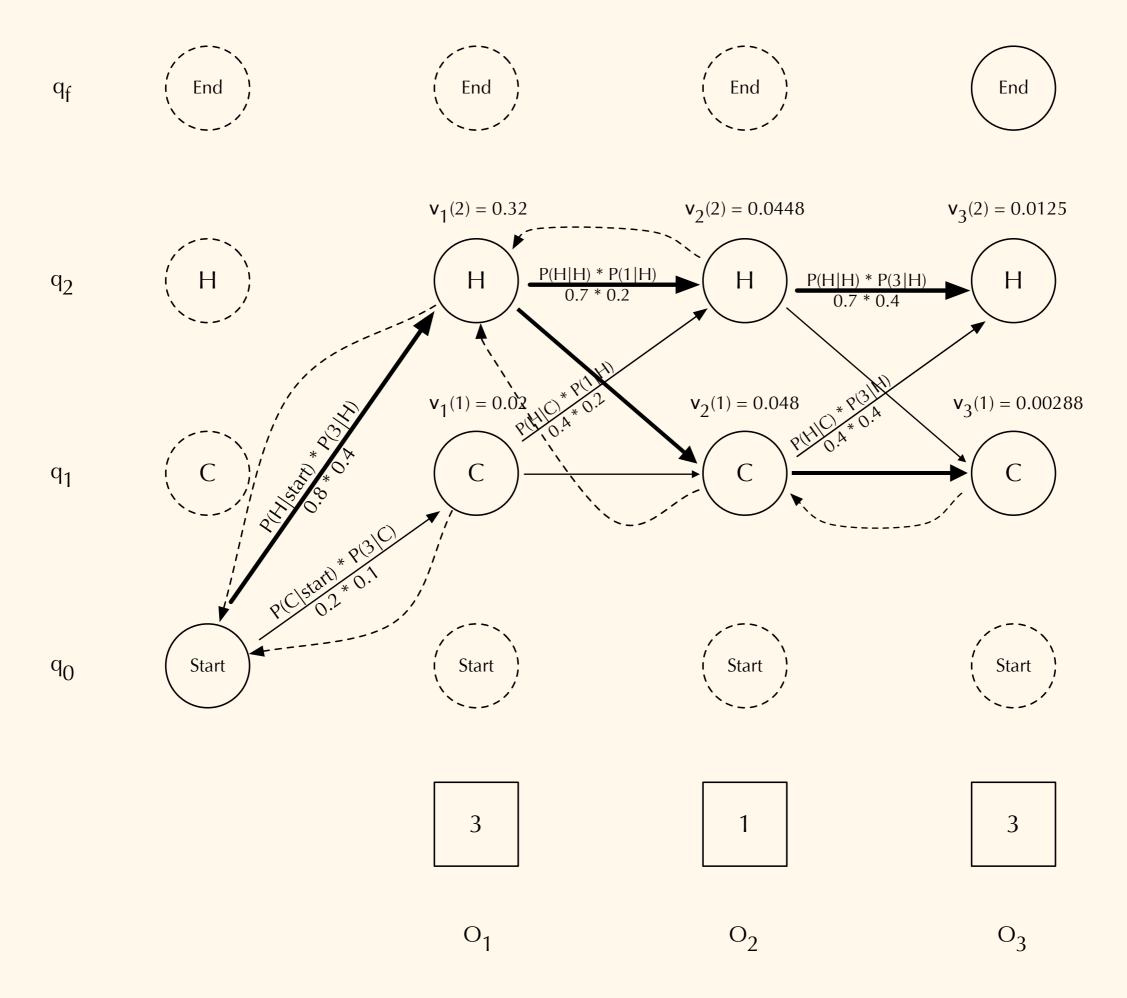


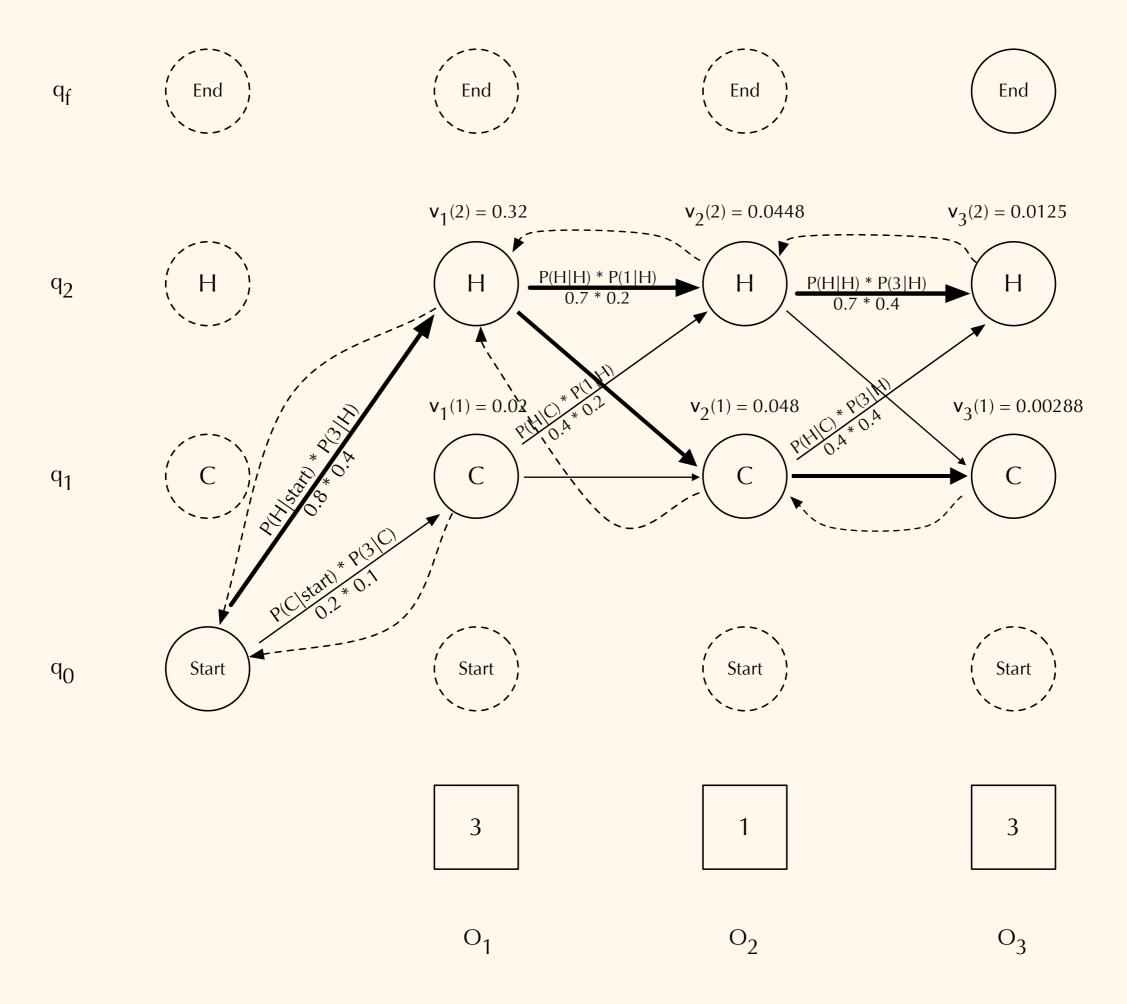


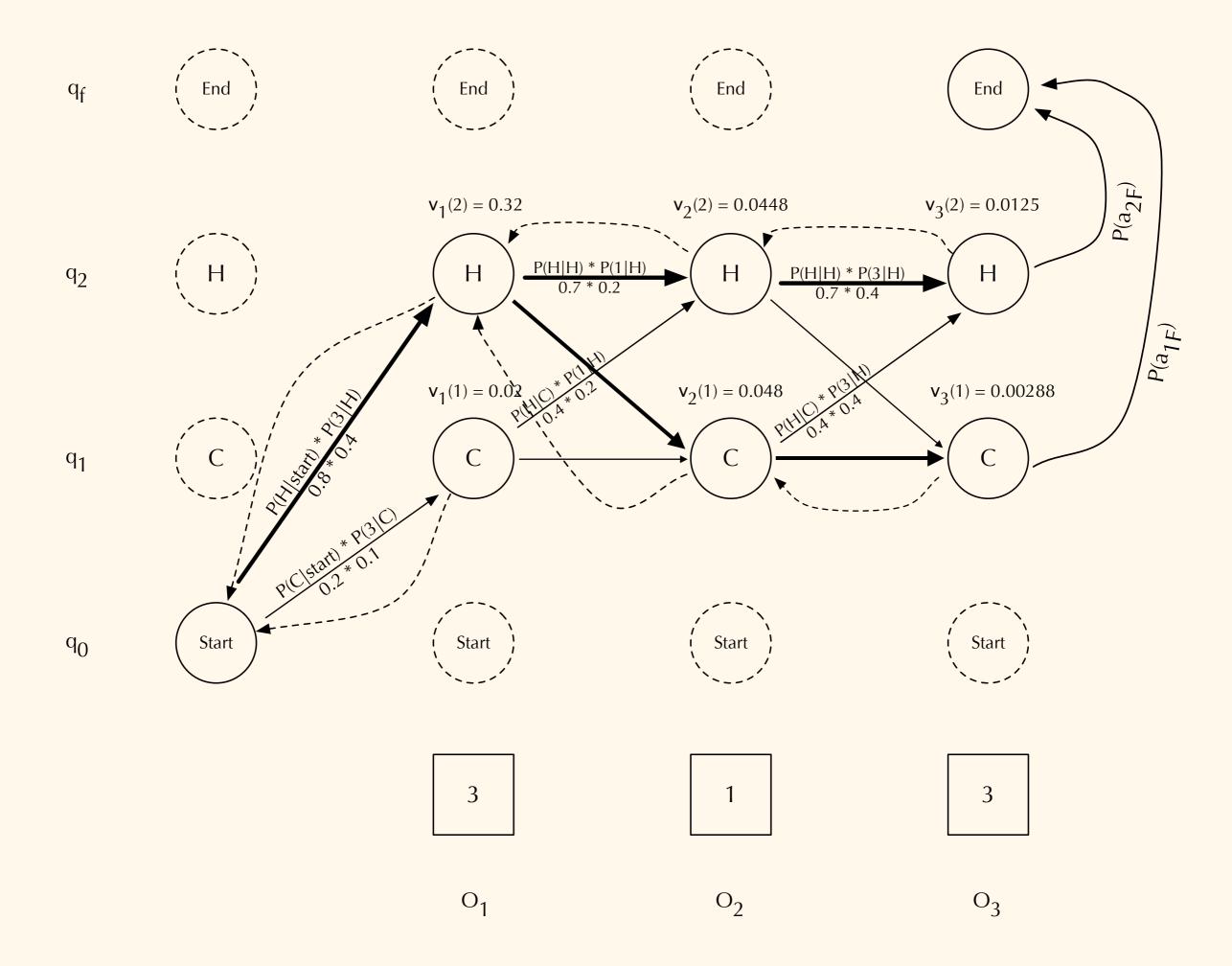


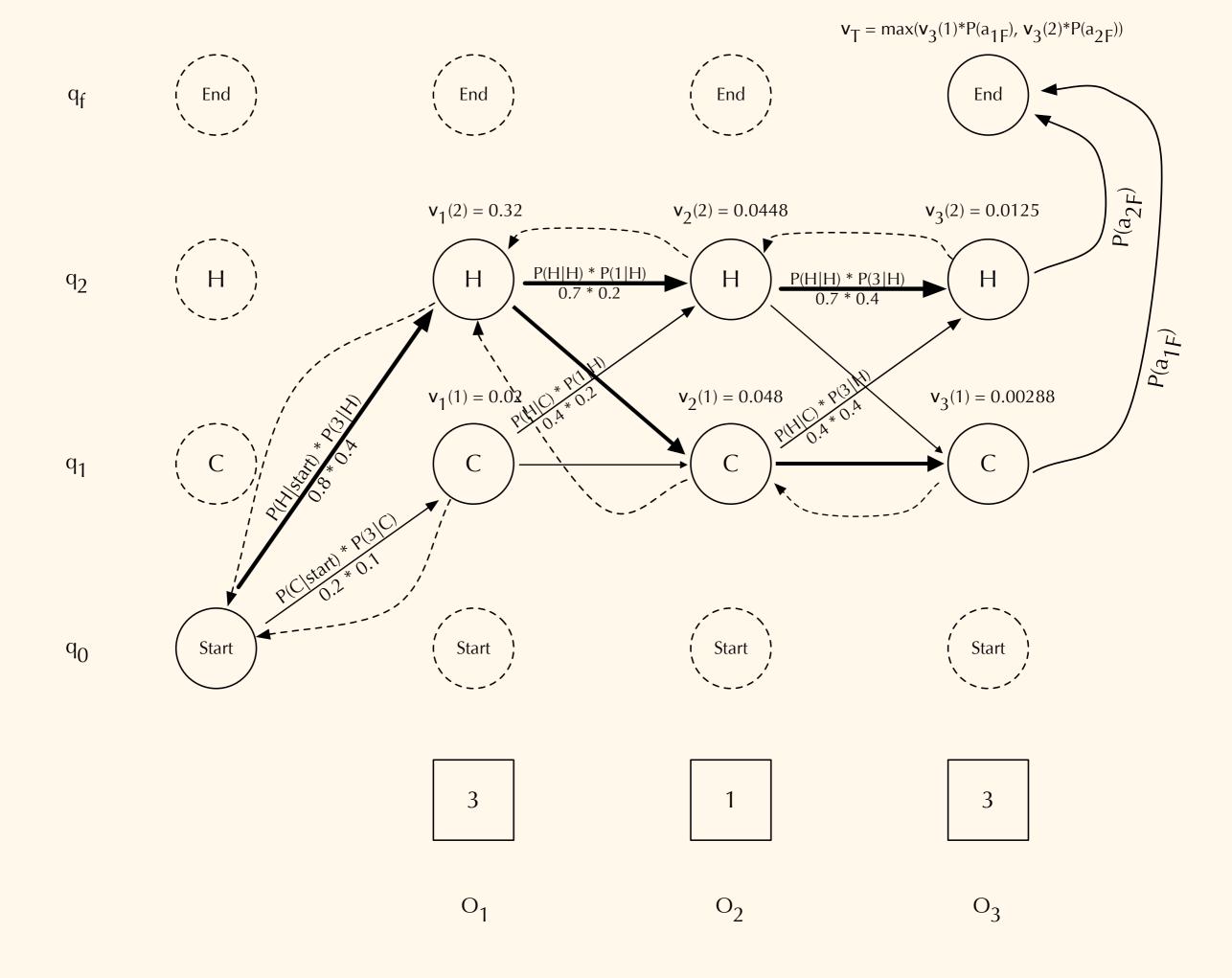


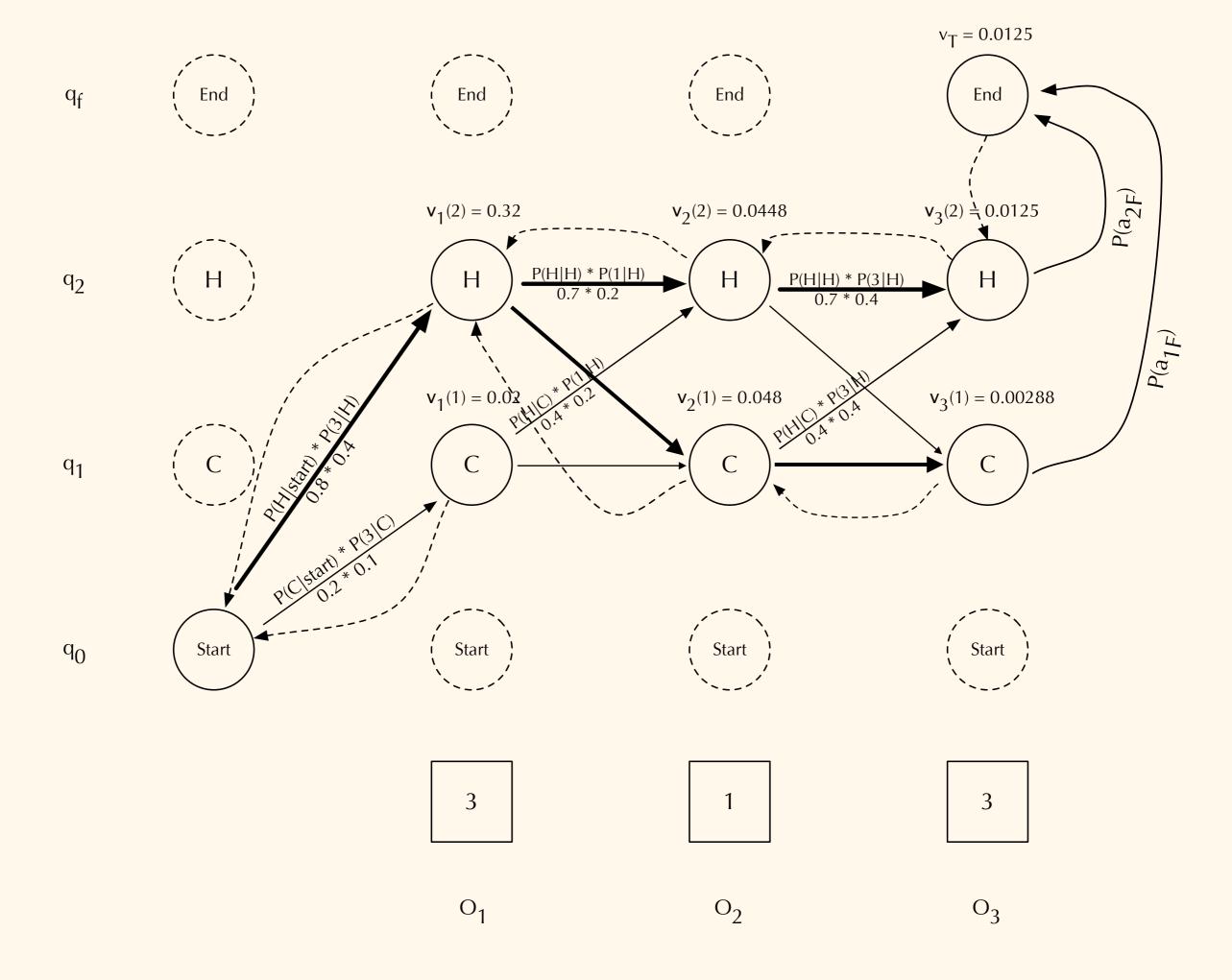












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Leonard Baum

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Lloyd Welch

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A related property is the *backward probability* $\beta_t(i)$, which represents the probability of seeing observations $O_{t:T}$, given that we are currently in state i at time t.

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A related property is the *backward probability* $\beta_t(i)$, which represents the probability of seeing observations $O_{t:T}$, given that we are currently in state i at time t.

This is calculated using the *backward algorithm*, which is very similar to the forward algorithm (but in reverse!).

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Probability of finishing (i.e., reaching end state) the observed sequence from state *i*.

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$$\beta_t(i) = \sum_{j=1}^{N} a_{ij} b_j(o_{t+1}) \beta_{t+1}(j)$$

$$\beta_T(i) = a_{i,F}$$

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Sum of the backwards probabilities of the different paths through the model that could happen from state *i* and time *t*.

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$$P(O|\lambda) = \alpha_T(q_F) = \beta_1(0) = \sum_{j=1}^{N} a_{0j}b_j(o_1)\beta_1(j)$$

$$\beta_T(i) = a_{i,F}$$

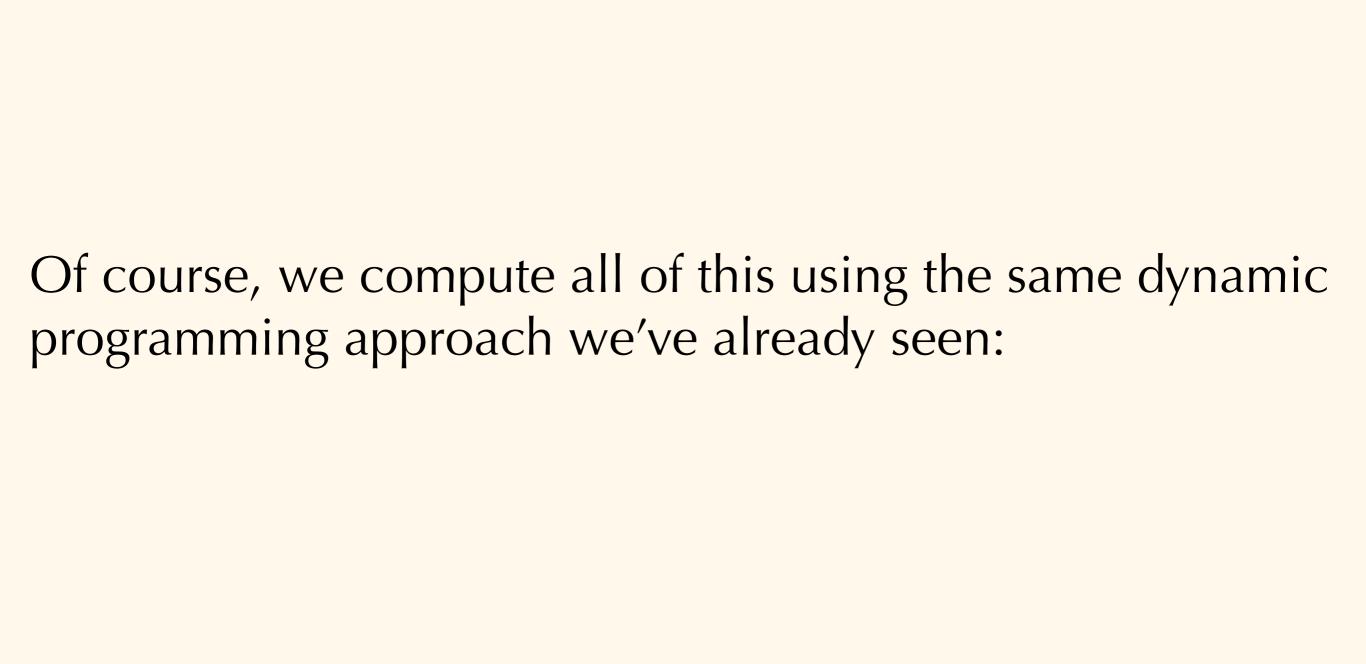
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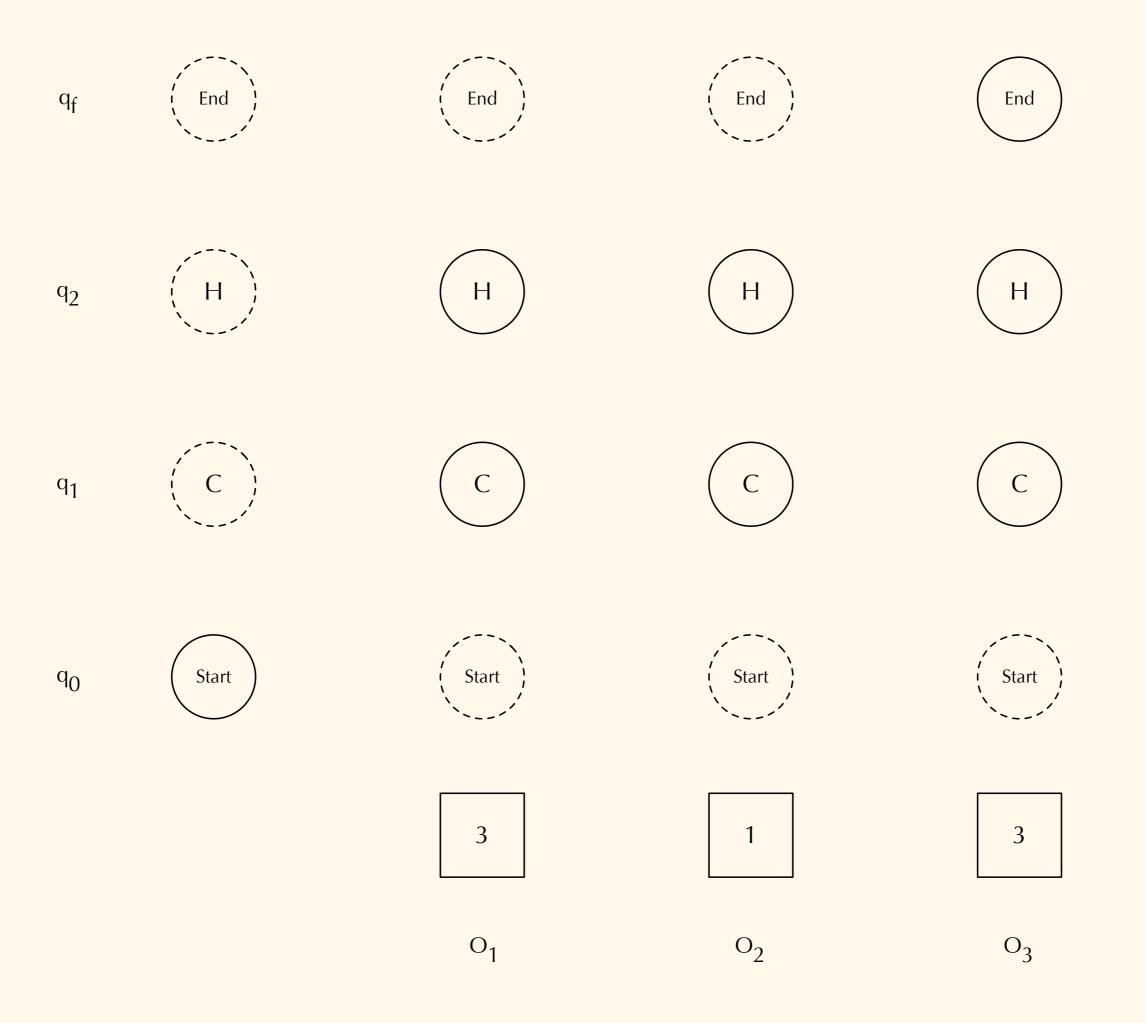
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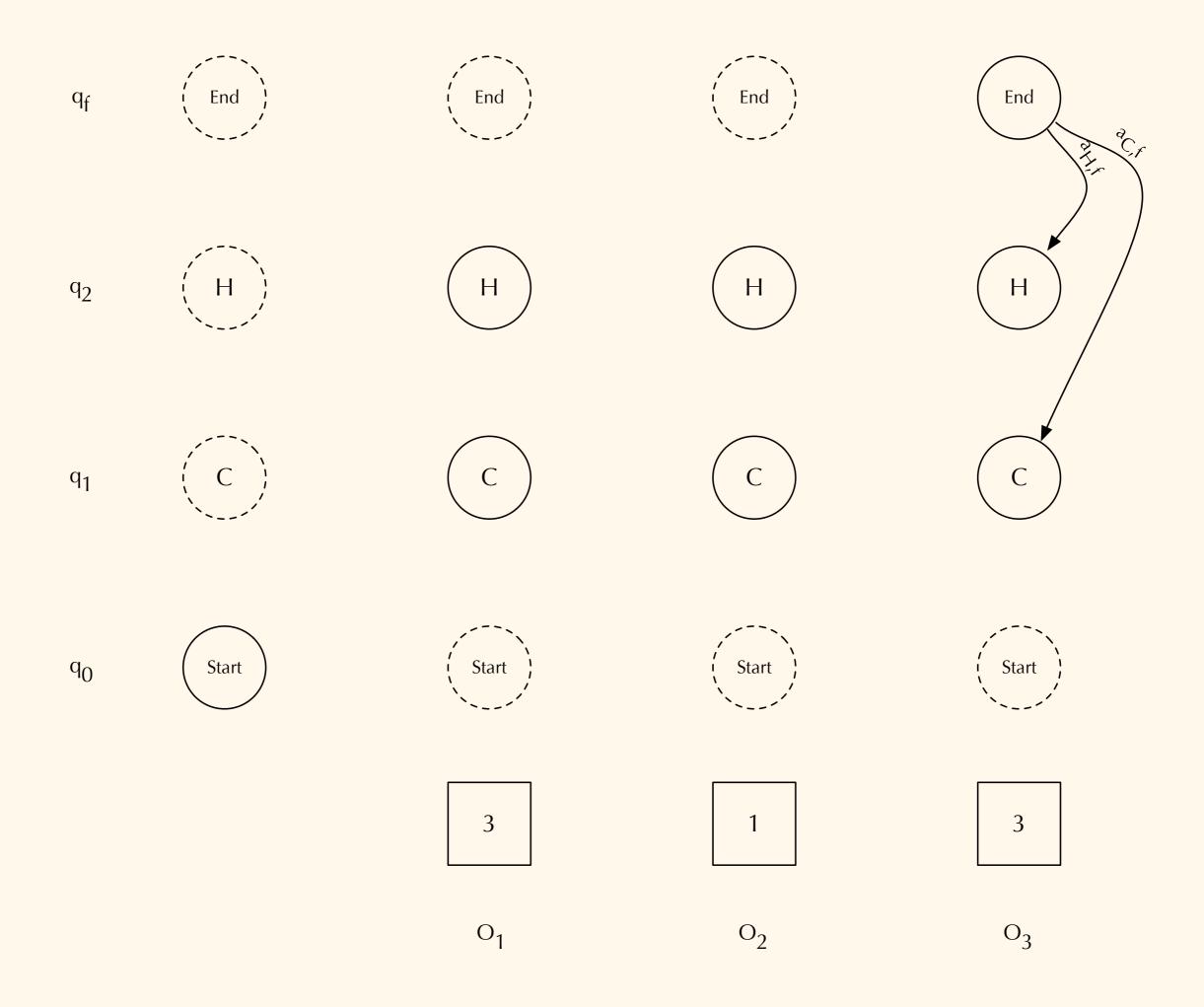
Sum of the backwards probabilities of the different paths through the model that could happen from state *i* and time *t*.

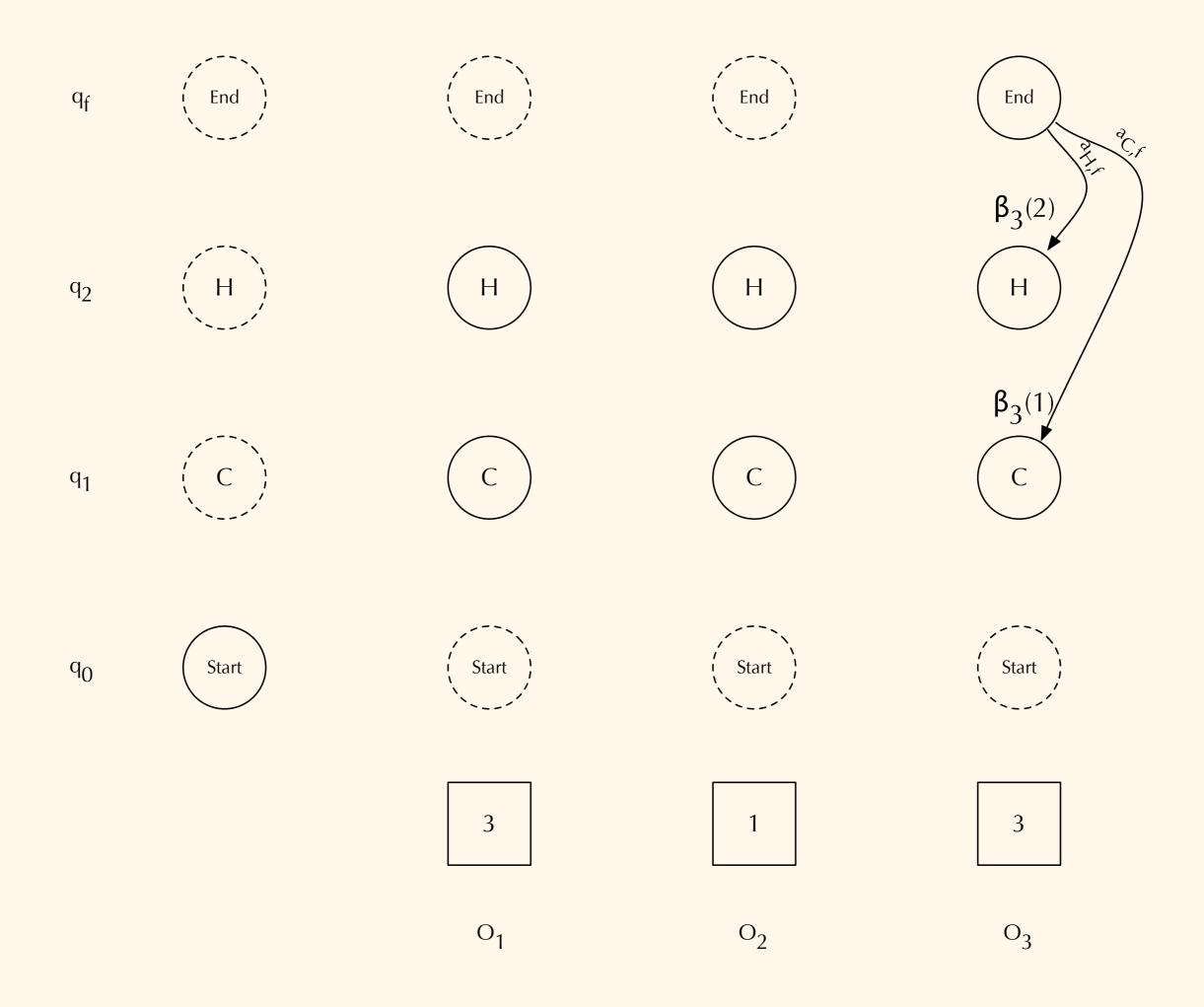
$$P(O|\lambda) = \alpha_T(q_F) = \beta_1(0) = \sum_{j=1}^{N} a_{0j}b_j(o_1)\beta_1(j)$$

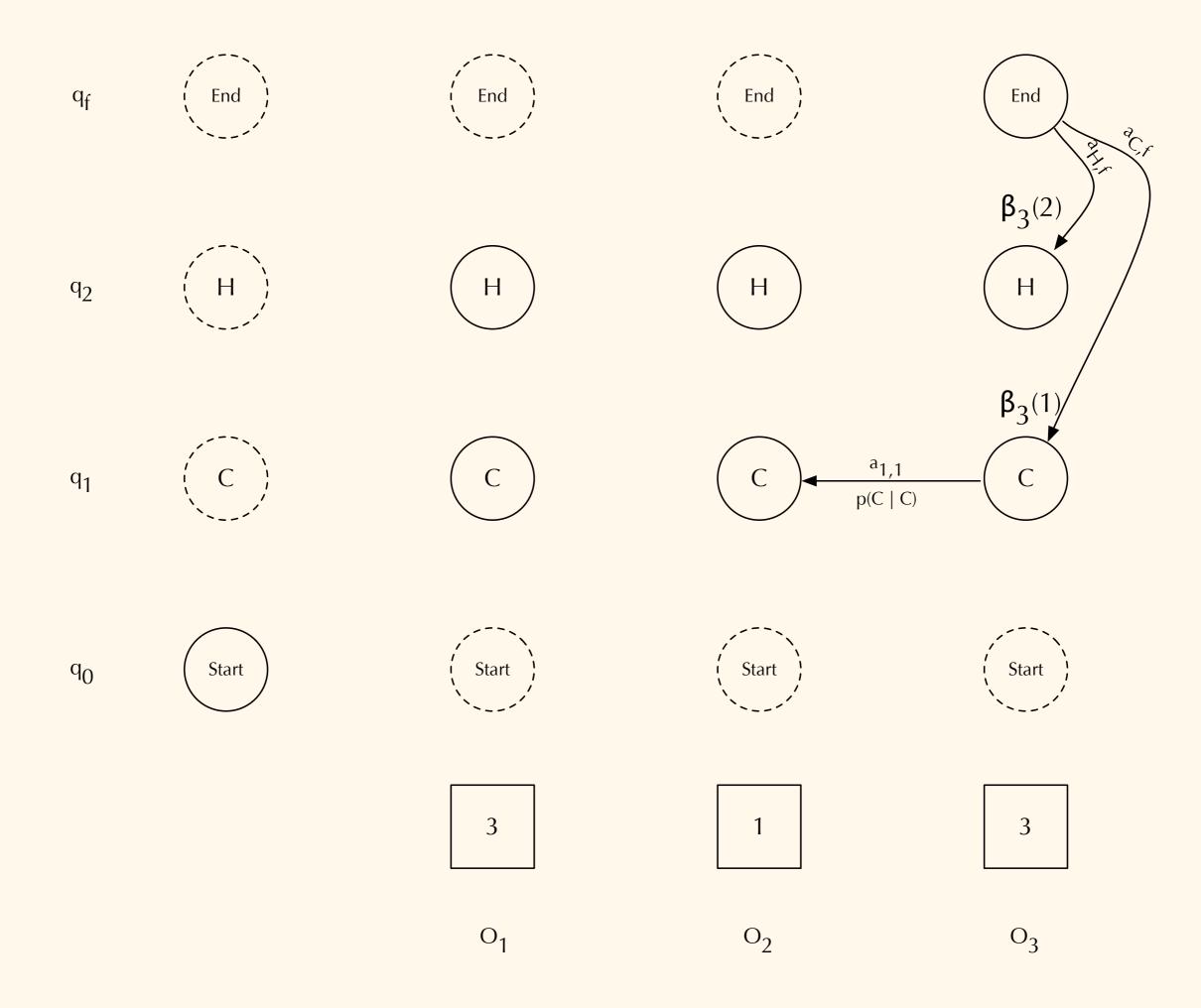
Final forward probability of observation given model.

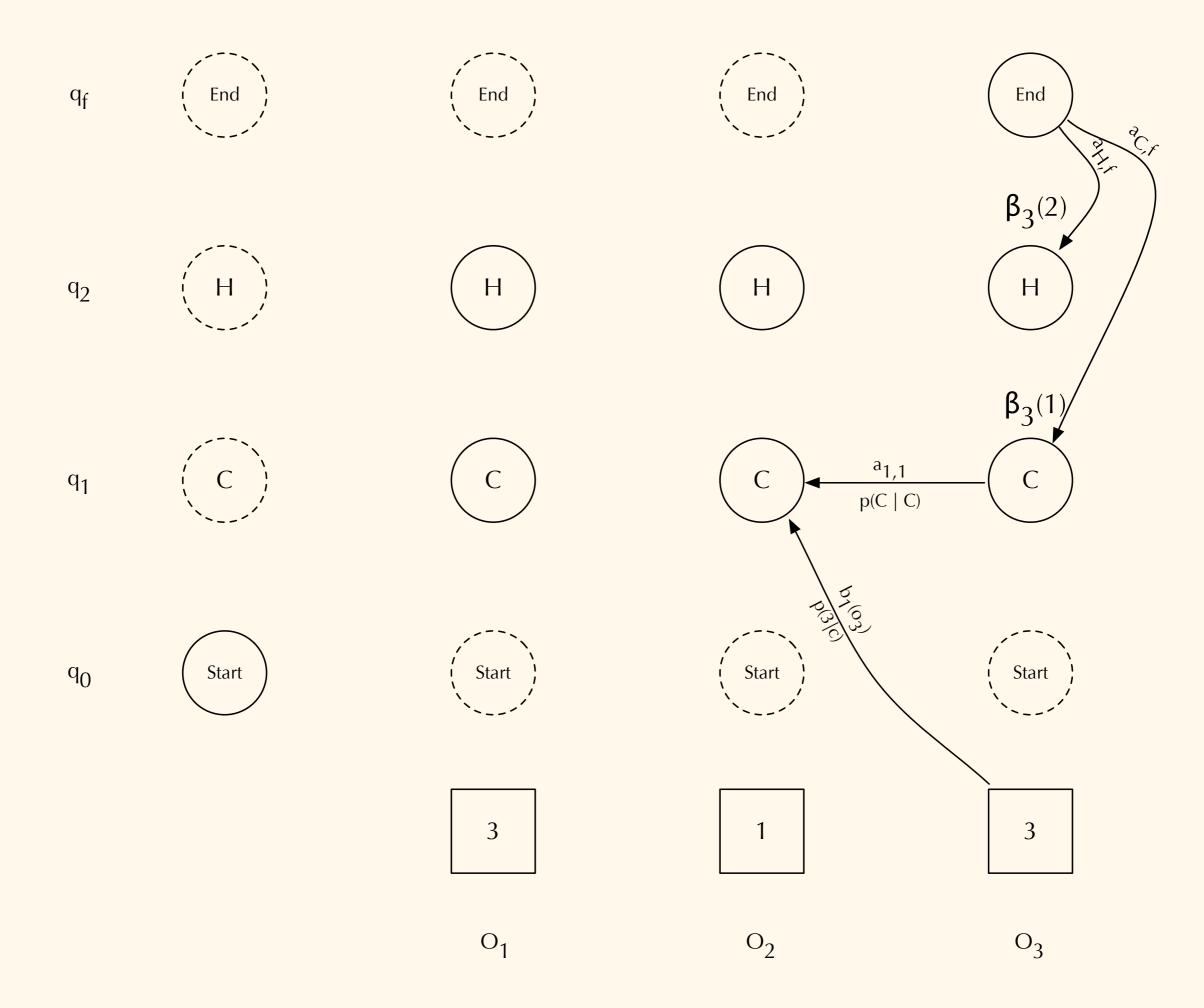


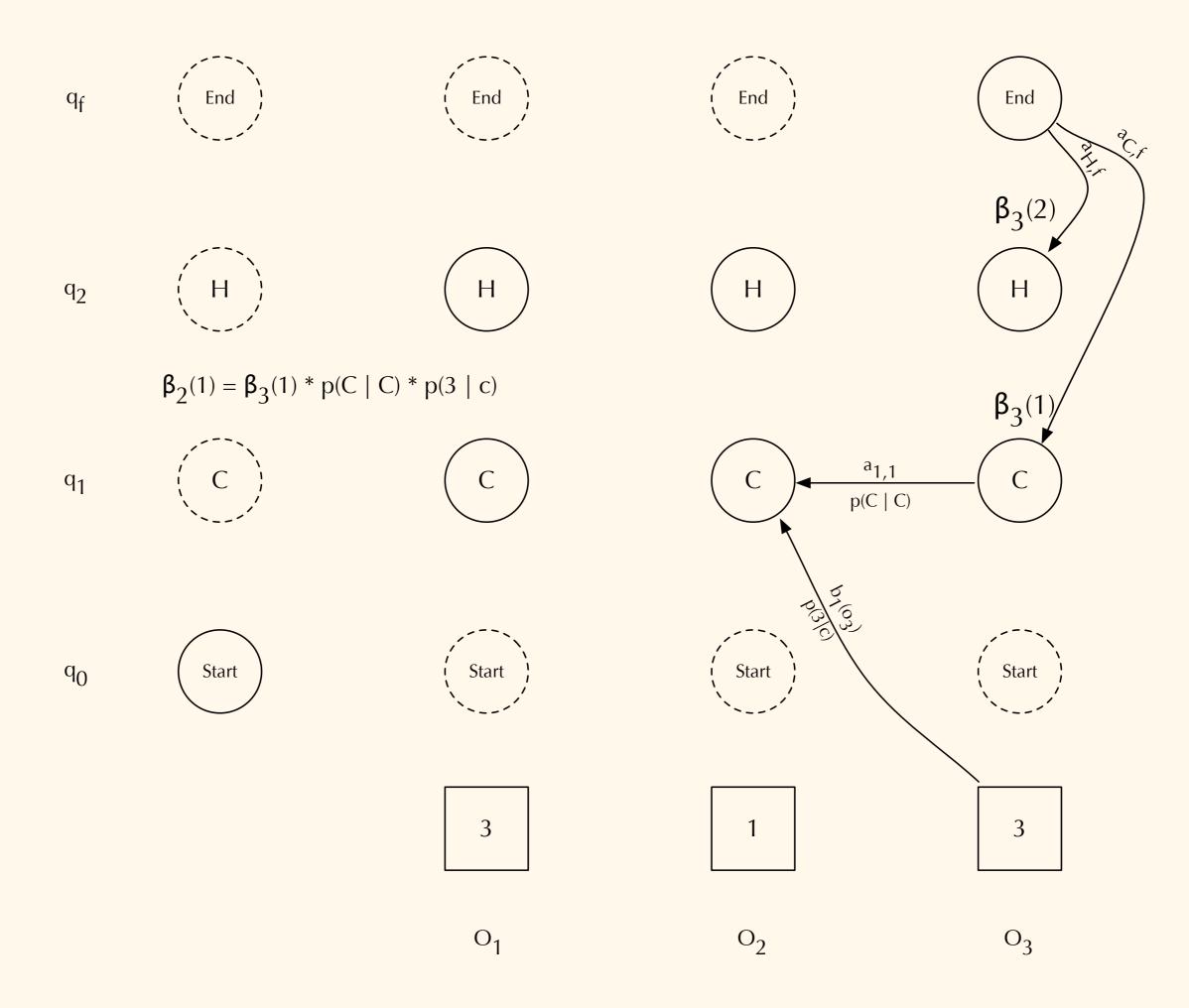


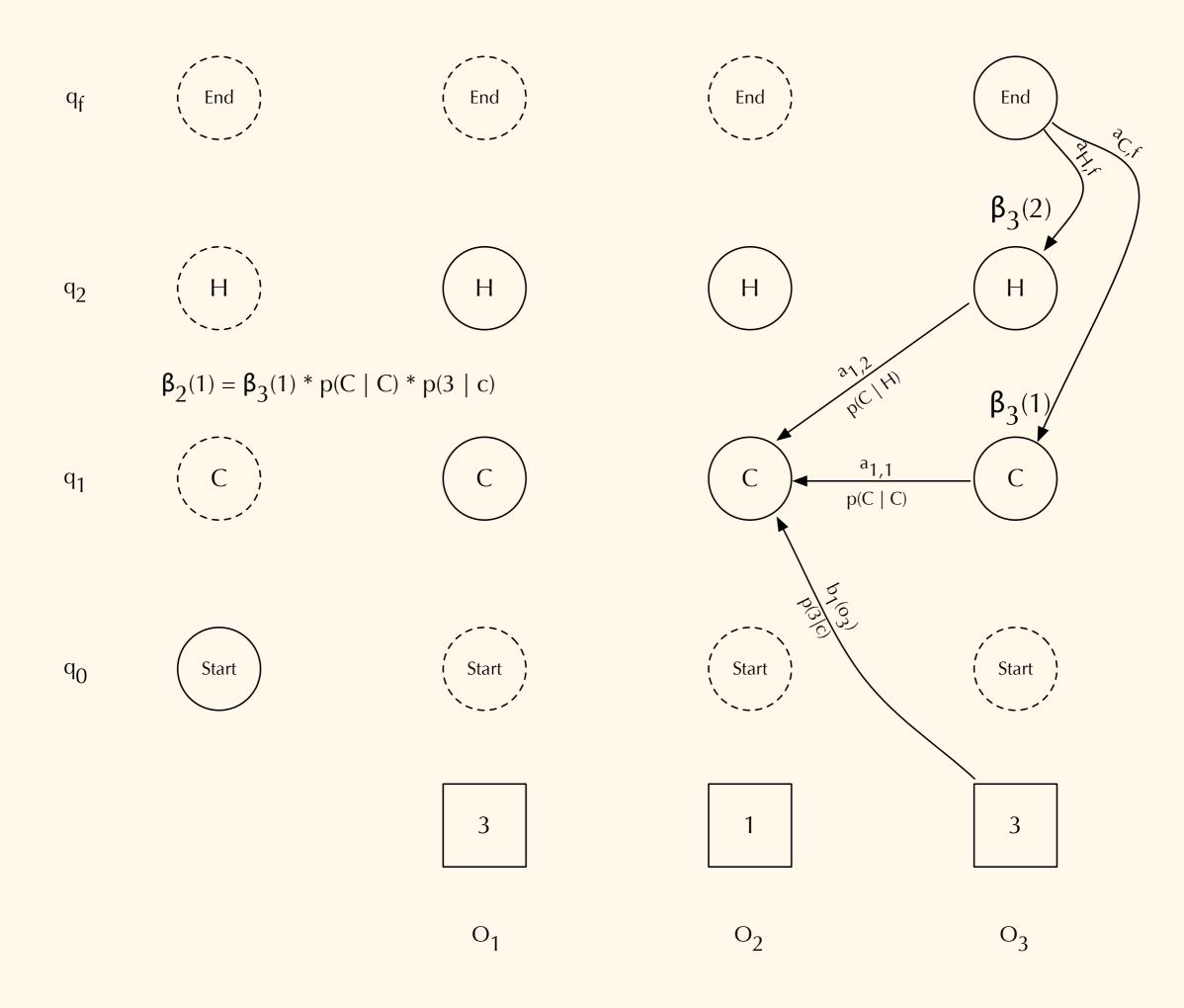


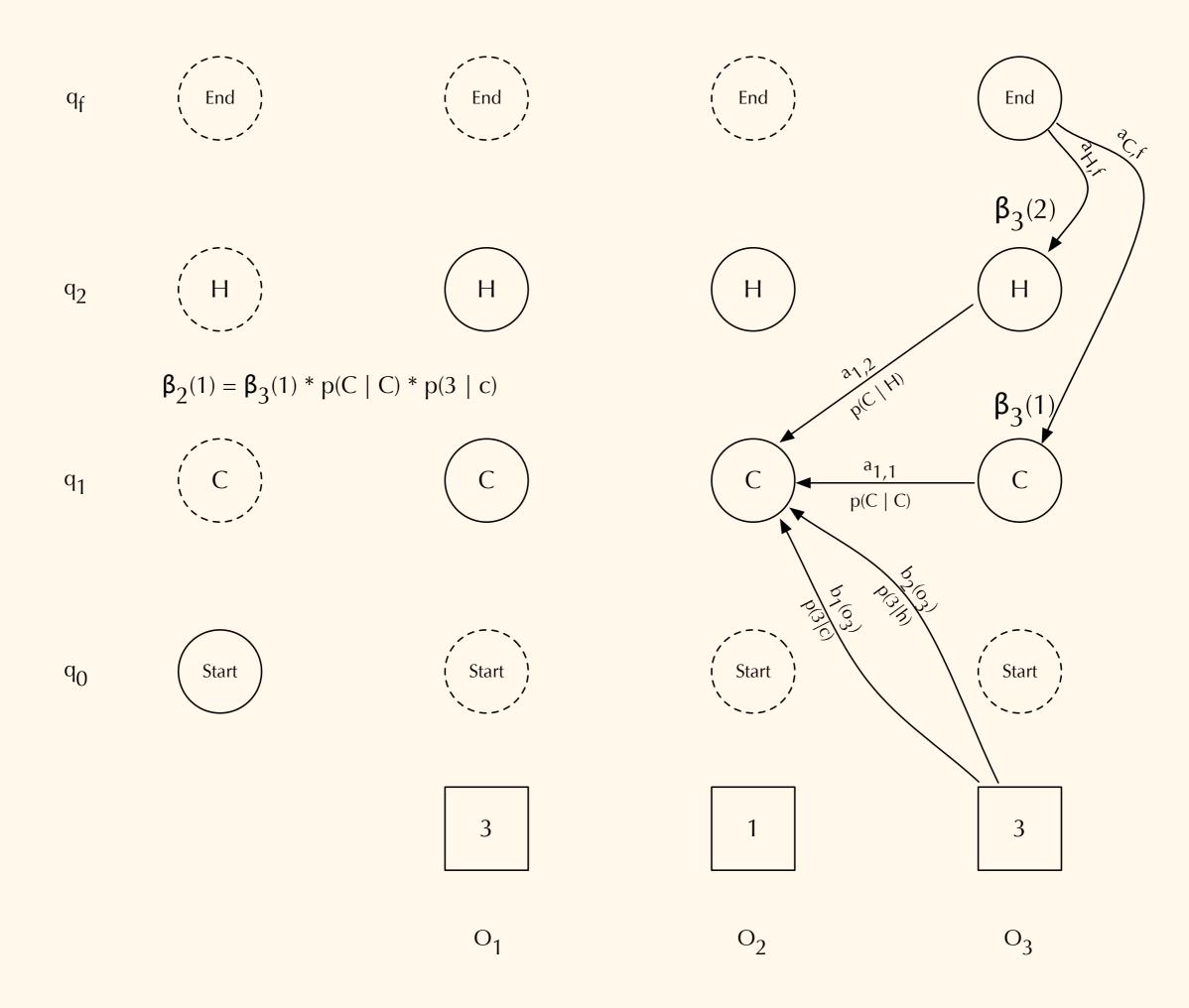


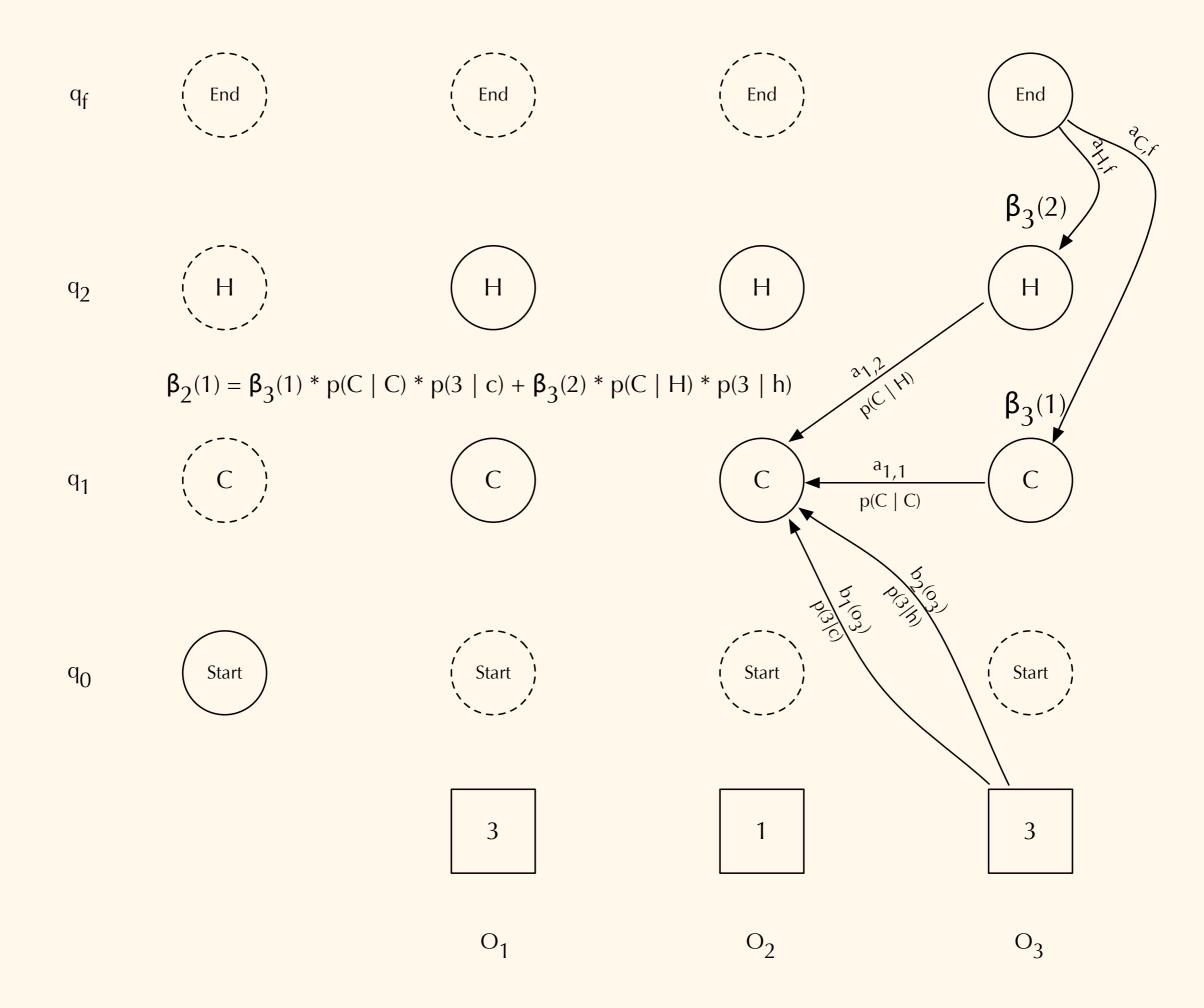












Baum-Welch is a variation on Expectation-Maximization...

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If we had an estimate of the probability of transition $i \rightarrow j$ occurring at each time t, we could sum them to get the total count for $i \rightarrow j$.

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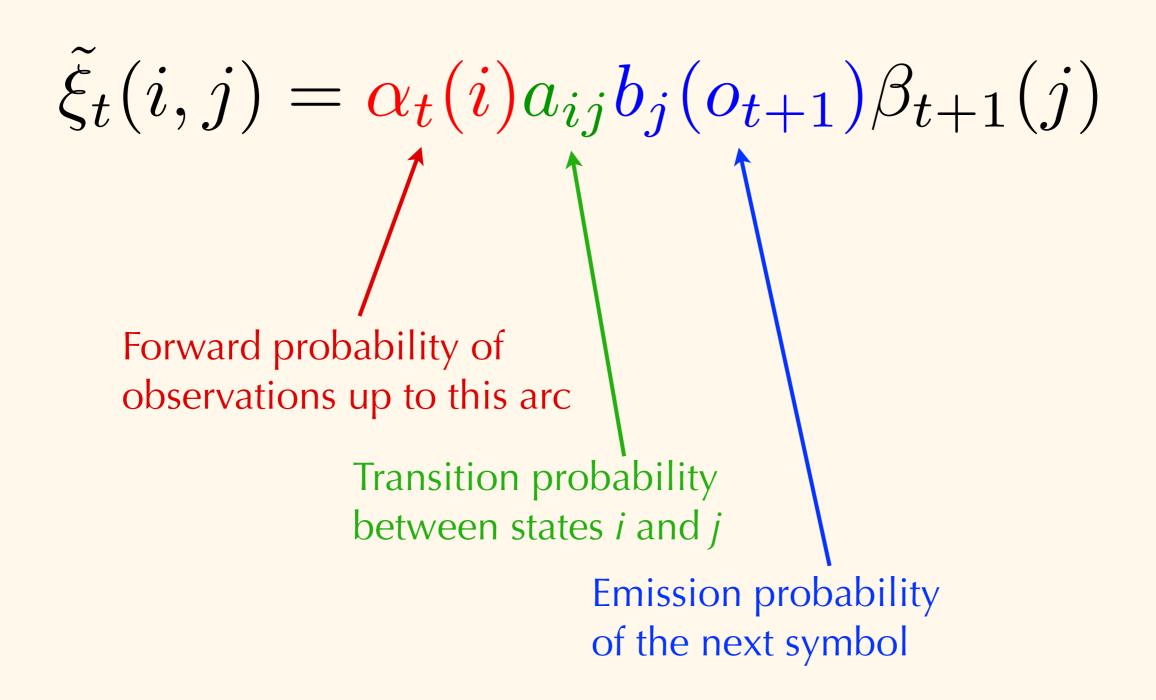
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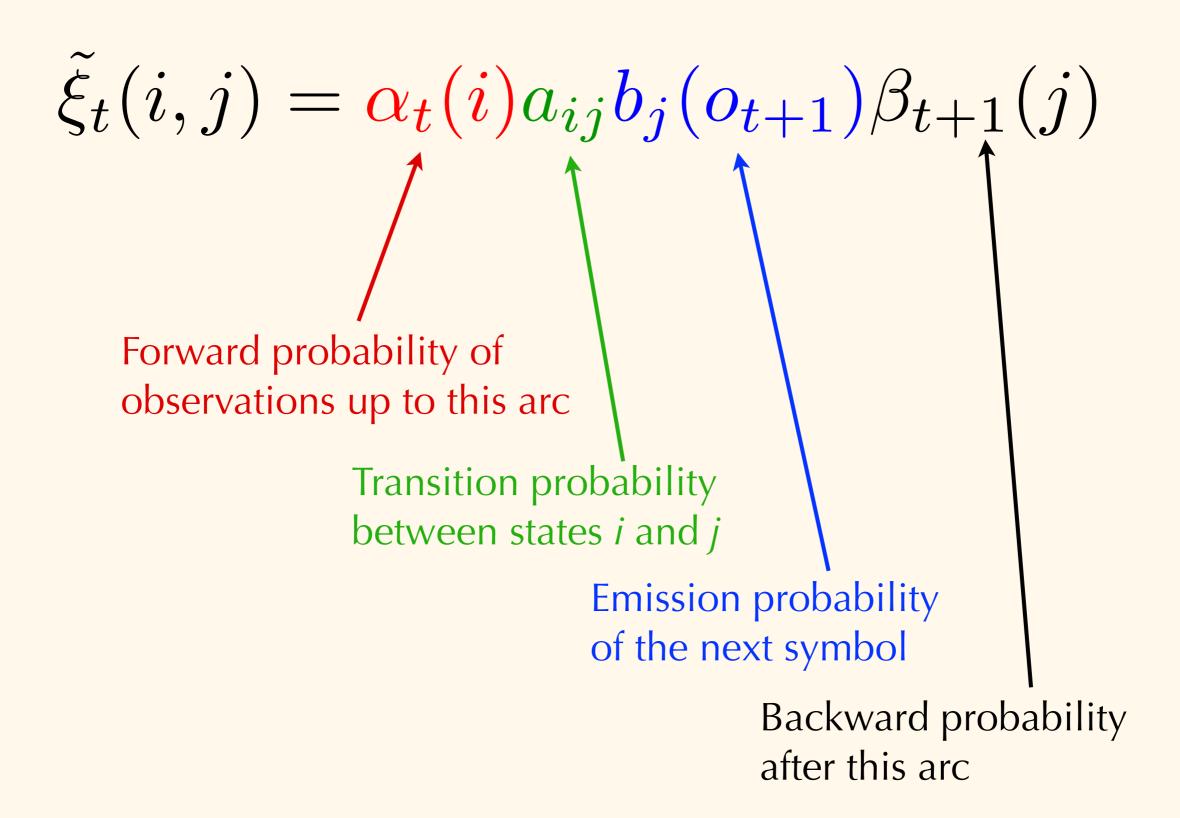
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 Forward probability of observations up to this arc

Transition probability between states i and j





We can transform $\tilde{\xi}_t(i,j)$, or $P(q_t=i,q_{t+1}=j,O|\lambda)$ into $\xi_t(i,j)$, or $P(q_t=i,q_{t+1}=j|O,\lambda)$, simply by dividing by $P(O|\lambda)$.

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$$P(O|\lambda) = \alpha_T(q_F) = \beta_1(0) = \sum_{j=1}^{N} a_{0j}b_j(o_1)\beta_1(j)$$

So, the final equation is:

$$\xi_t(i,j) = \frac{\alpha_t(i)a_{ij}b_j(o_{t+1})\beta_{t+1}(j)}{\alpha_T(N)}$$

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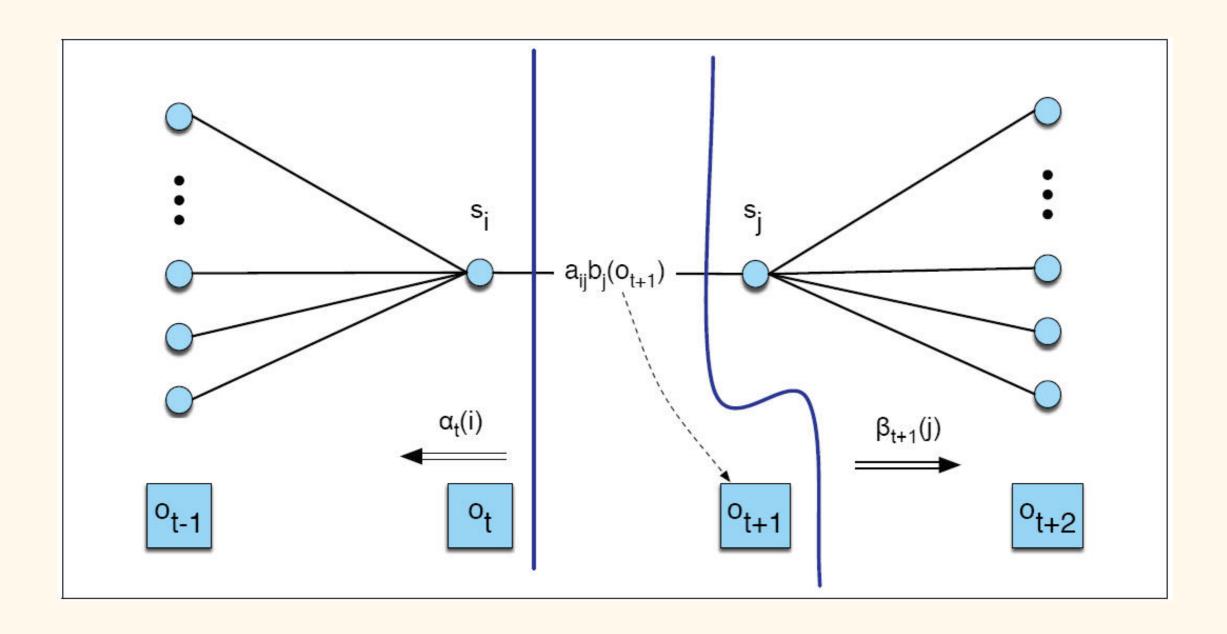
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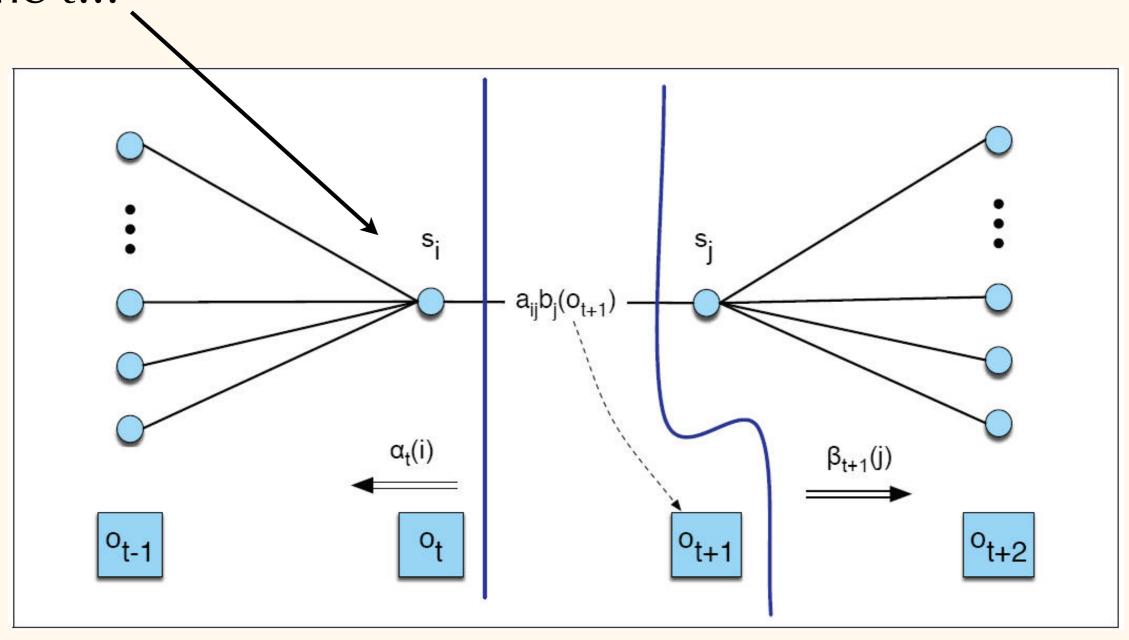
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Which looks a lot like:

$$\hat{a}_{ij} = \frac{\text{expected } \# \text{ transitions from state } i \text{ to state } j}{\text{expected number of transitions from state } i}$$



All of the ways the model could have gotten into state *i* at time *t*...

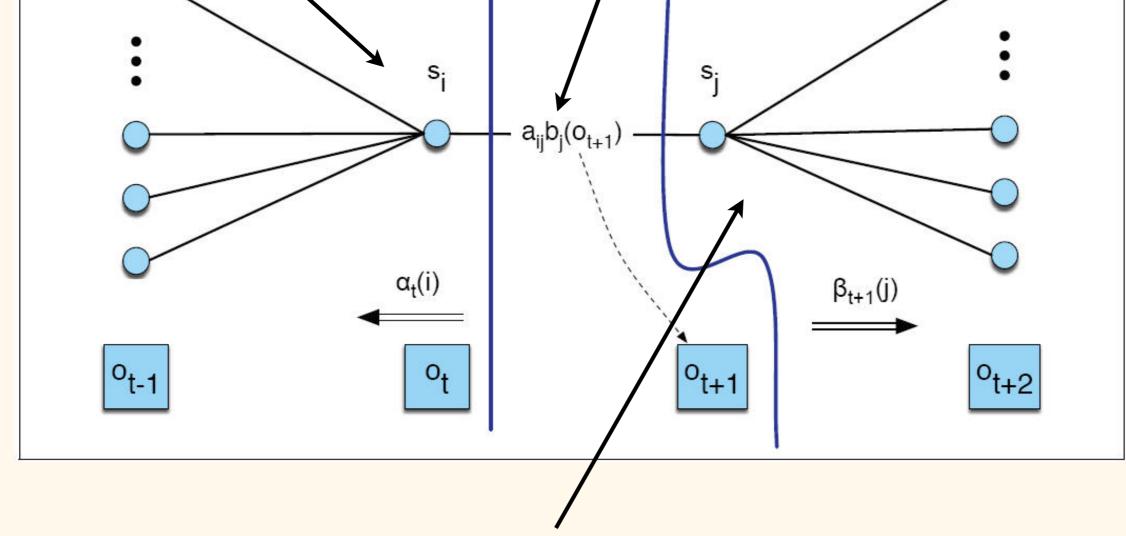


All of the ways the model could have gotten into state *i* at time *t*...

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All of the ways the model could have gotten into state *i* at time *t*...

... the likelihood of going from i to j while emitting o_{t+1} ...



... all of the ways the model could finish from state j at time t+1.

$$\hat{b}_j(v_k) = \frac{\text{expected } \# \text{ times in state } j \text{ and observing symbol } v_k}{\text{expected number of times in state } j}$$

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Using the same trick as before:

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"Probability of getting to this state at this time point, times the probability of the rest of the observations given this state and this time point"

$$\hat{b}_j(v_k) = \frac{\sum_{t=1}^{T} \gamma_t(j)}{\sum_{t=1}^{T} \gamma_t(j)}$$

"Only count observations where the observed emission was v_k ."

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We can go back, calculate *new* forward and backward trellises, and re-compute *A* and *B*.

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In practice, much depends on our initial estimates, and so we often use additional information when possible (e.g., encoding impossible transitions, etc.).